How Violence Against LGBTQ+ People Motivates Prosocial Attitudes Toward LGBTQ+ Group Members

Marcel F. Roman¹ and Jack Thompson²

¹Postdoctoral Research Fellow, Harvard University ²Postdoctoral Research Fellow, University of Exeter

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Abstract

We present a *Fickle Prosocial Violence Response Model* (FPVR) to explain how indirect exposure to civilian-perpetrated violence against marginalized minority groups motivates prosocial attitudes toward victimized groups. Although the mass public may not sympathize with marginalized groups, they may adopt prosocial attitudes toward marginalized groups subject to civilian-perpetrated violence if the violence is salient and perceptibly illegitimate. However, the adoption of prosocial attitudes may be fickle. We find evidence consistent with the model. Studies 1-3 show high-profile violence against LGBTQ+ people increases support for LGBTQ+ rights and reduces negative attitudes toward LGBTQ+ group members. But, the adoption of prosocial attitudes is short-term. Study 4 shows less salient violence against LGBTQ+ people may not engender prosocial attitudes at the outset. Our findings suggest violent events must be sufficiently salient to initially motivate prosocial beliefs. Nevertheless, salient civilian-perpetrated violence against marginalized groups may not sustainably motivate prosocial beliefs toward targeted groups.

Keywords: exposure to violence; political violence; prosocial attitudes; intergroup relations; LGBTQ+ politics

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1 Introduction

Since the 1969 Stonewall Uprising, there have been numerous instances of anti-LGBTQ+ violence in the US. Despite progress on LGBTQ+ rights (Flores, 2014), anti-LGBTQ+ violence and hate crimes have increased recently¹ while several states introduced a record number of anti-LGBT+ laws in the last year.² Perhaps the most prominent, recent, instance of perceptibly anti-LGBTQ+ violence was the 2022 Club Q massacre, where a gunman killed 5 clubgoers at a Colorado Springs LGBTQ+ nightclub. These violent acts, while sympathy-inducing within media and amongst some political elites, may reflect heteronormative societal norms resistant to change.³ Therefore, an open question is whether *indirect* (i.e. media observation of violence) exposure to high-profile civilian-perpetrated violence against LGBTQ+ group members motivates introspection among the mass public, shifting their attitudes prosocially toward LGBTQ+ people.

We synthesize several theoretical insights and present a *Fickle Prosocial Violence Re*sponse (FPVR) model to explain how violence against marginalized groups may elicit prosocial attitudes toward targeted groups. Although the mass public may not strongly empathize with marginalized minority groups (Cikara et al., 2014), violence against marginalized groups may elicit prosocial attitudes if the violence is salient, perceptibly illegitimate, and the media and/or elites respond sympathetically (Iyengar, 1994; Birkland, 1998; Branscombe and Miron, 2004; Harth et al., 2008; Vossen et al., 2017). However, prosocial attitude adoption may be short-term. Social group attitudes are typically entrenched, even in light of salient events (Sears, 1993; Tuch and Weitzer, 1997; Kite et al., 2019). Immediate adoption of prosocial beliefs after violence may be counterbalanced by countervailing information in a discriminatory society (Vuletich and Payne, 2019). Elite messaging and pressure to support targeted groups may dissipate after an event loses salience (Downs, 1972), undercutting

¹https://www.hrc.org/press-releases/new-fbi-hate-crimes-report-shows-increases-in-anti-lgbtq-attacks ²https://www.aclu.org/press-releases/over-120-bills-restricting-lgbtq-rights-

introduced-nationwide-2023-so-far

³*Heteronormativity* is "privileging gender conformity, heterosexuality, and nuclear families over "deviant" forms of gender expression, sexuality, and family (Pollitt et al., 2021)"

sustainable prosocial attitudinal shifts (Zaller, 1992; Birkland and Lawrence, 2009).

We find evidence consistent with the FPVR model by using several surveys and an unexpected-event-during-survey-design. Studies 1-3 demonstrate the public adopts prosocial attitudes toward LGBTQ+ community segments and their political rights shortly after civilian violence against LGBTQ+ group members (i.e. the Pulse massacre, Matthew Shepard's murder). However, these attitudinal shifts do not persist. Moreover, Study 4 demonstrates the Club Q massacre had no effect on anti-gay, anti-trans, attitudes. Consistent with the FPVR model, we provide evidence the null effects at the outset are due to the less salient nature of the Club Q massacre vis-a-vis the Pulse massacre and Shepard's murder. We provide corroborating evidence by demonstrating less salient violent incidents against LGBTQ+ people outside those in Studies 1-4 largely do not motivate prosocial mass attitudes.

Our theory and evidence makes several contributions. First, the FPVR model helps explain how violence against marginalized groups motivates prosocial beliefs toward targeted groups among the mass public. Our model is important in light of several salient instances of civilian violence against marginalized groups in the US: Vincent Chin's 1982 murder, a Chinese man murdered due to anti-Japanese resentment; James Byrd's 1996 murder, a Black man lynched by white supremacists in Texas; the 2015 Charleston Church massacre, where a white supremacist murdered 9 Black churchgoers; the 2015 Stanford sexual assault case (*People v. Turner*), where a Stanford undergraduate man sexually assaulted a woman; the 2019 El Paso massacre, where a white supremacist killed 23 people, mostly Latinos, to counteract a "Hispanic invasion"; the 2021 Atlanta spa shooting, where 8 people, mostly Asian women, were killed; and the 2022 Buffalo massacre, where a white supremacist killed 10 Black people because he felt non-whites were "replacing" whites. We show these highprofile violent events 1) may not serve as sustainable moments of reevaluation concerning the socio-political status of marginalized groups and 2) may not motivate prosocial attitudes at the outset if they are insufficiently salient. Thus, our model and evidence may explain why these events have not led to societal adjustment of beliefs perpetuating social inequalities.

Second, our analysis is the first to examine the attitudinal consequences of real-world *civil*ian violence against LGBTQ+ community segments on prosocial attitudes toward LGBTQ+people among the mass public. Prior research on violence and prosocial attitudes in the US focuses on state (i.e. police) violence against Black people (Tuch and Weitzer, 1997; Sigelman et al., 1997; Reny and Newman, 2021). Extending research on the attitudinal consequences of violence to the domain of civilian violence against LGBTQ+ community segments is theoretically important since prosocial attitudinal shifts toward victimized groups may be conditional on 1) whether the violence is state-imposed and 2) the targeted group.

One perspective is that the effects of *civilian* violence against LGBTQ+ group members on prosocial attitudes toward LGBTQ+ people may be *weaker* and *less temporally sustainable* than prior studies examining racialized *state* violence. The mass public may attribute *state* violence to systemic problems within reformable institutions, motivating policy preferences benefiting targeted groups (Oskooii, 2016). Civilian violence may be rationalized as a problem inherent to a troubled individual as opposed to the public's aggregate queerphobia (Ott and Aoki, 2002),⁴ undercutting introspection over one's own queerphobic beliefs. Moreover, unlike racialized state violence, the violence we examine are not associated with subsequent mass protest, which may sustain the violent event's salience, facilitating longlasting attitudinal shifts (Reny and Newman, 2021).

Conversely, the effects of violence against LGBTQ+ community segments on prosocial attitudes may be *stronger* and *sustainable* relative to the effects of racialized state violence. Racial attitudes are typically assumed to be stable (Tesler, 2015),⁵ and preexisting evidence shows prosocial responses to racialized state violence are fickle (Tuch and Weitzer, 1997; Chudy and Jefferson, 2021). However, as far as data could be recorded, the public has increasingly adopted prosocial attitudes toward LGBTQ+ people over time (Flores, 2014). Therefore, the public may be inclined to engage in sustained introspection over their queer-

⁴ "Queer" denotes a gender/sexual identity that does not correspond to heterosexual notions of sexuality and gender.

⁵But see Hopkins and Washington (2020)

phobia after violence against LGBTQ+ individuals.

Our evidence adjudicates between these perspectives and bolsters prior findings consistent with the consequences of racialized state violence. Even in a distinct domain like civilian violence against LGBTQ+ people, the public may adopt prosocial attitudes toward targeted groups, but only briefly.

Third, our analysis contributes to the *focusing event* literature (Birkland, 1998). Prior research shows salient events shift mass attitudes, but briefly since these events eventually lose salience (Sigelman et al., 1997; Birkland and Lawrence, 2009). Additionally, LGBTQ+ politics research demonstrates high-profile pro-LGBTQ+ court cases (Flores, 2015), Pride parades (Ayoub et al., 2021), and celebrities coming out (Miller et al., 2020), can motivate prosocial attitudes toward LGBTQ+ people. But, this research 1) places little emphasis on the sustainability of attitudinal shifts, 2) does not assess event salience variation at the outset, and 3) does not focus on the consequences of violence against LGBTQ+ people, which may *reflect*, instead of *undercut* queerphobic beliefs. We provide new evidence bolstering prior findings on brief attitudinal shifts after "focusing events" in an unexplored domain.

2 Violence and Prosociality

Preexisting theory and evidence demonstrate *direct* or *proximal* (i.e. via close social ties, like family, friends, acquaintances) violence exposure during inter-group conflict may motivate parochialism, encourage intra- but not inter-group altruism, and undercut emotional substrates facilitating inter-group prosocial behaviors and attitudes, including, positive evaluations of outgroups and support for their political rights (Rusch, 2014; Lupu and Peisakhin, 2017; Mironova and Whitt, 2018; Hadzic et al., 2020). Other evidence, building on *Post-Traumatic Growth* and *Altruism Born of Suffering Theory* (Staub and Vollhardt, 2010), shows inter-group violence can motivate prosocial, altruistic attitudes and behaviors toward outgroups (Bakke et al., 2009). Direct or proximal violence exposure may motivate inter-

group prosociality since victimization generates a basis for empathy (Sirin et al., 2021).

Although prior work suggests direct or proximal exposure to inter-group, mostly interethnic, violence motivates prosociality, it is less clear how one-sided⁶ *indirect* exposure to violence against LGBTQ+ people influences prosocial attitudes toward LGBTQ+ group members among dominant groups or the mass public. Hereafter, we define prosocial attitudes as positive feelings toward LGBTQ+ group members and policies facilitating their rights.

One expectation is that indirect exposure to one-sided violence may *not* motivate prosocial beliefs. Insufficient media coverage and attention to violent events may not produce agenda-setting effects that mobilize prosocial mass attitudes (Birkland, 1998). Additionally, *Social Identity Theory* (SIT) implies dominant group members garner self-esteem from minority group marginalization (Tajfel and Turner, 1982). Thus, the mass public may garner psychic benefits from indirectly observing violence against minority groups (Cikara et al., 2014). Consistent with *Inter-group Emotions Theory* (IET), these dynamics may be exacerbated by the absence of direct experiences with analogous violence facilitating empathy (Sirin et al., 2021). Moreover, the social distance between modal mass public members and, for example, LGBTQ+ people, may generate an empathy gap, undercutting the adoption of prosocial attitudes after indirect violence exposure (Cikara et al., 2014).⁷ Therefore, we may observe an empirical pattern consistent with Figure 1, Panel A, where prosocial attitudes among the mass public toward a marginalized group do not change after indirect exposure to civilian violence against said group.

Another expectation is that, under some conditions, indirect exposure to violence against marginalized groups may motivate prosocial attitudes to ameliorate conditions concomitant with the violence. *Focusing Event Theory* implies salient violent incidents can mobilize mass attitudes (Birkland, 1998). These attitudes may be more likely to be mobilized prosocially if the media and elites express the violence is illegitimate and are sympathetic toward the targeted group (Zaller, 1992; Iyengar, 1994). Indeed, sympathetic messages expressed by

⁶ "One-sided" refers to dominant group-perpetrated violence.

⁷For the Pulse massacre, the empathy gap may be amplified by the predominantly Latinx victims.

partisan elites after violence may help socially conservative co-partisans reconsider prejudicial attitudes (Harrison and Michelson, 2017). The media also has a powerful influence on LGBTQ+ mass attitudes. Positive LGBTQ+ media portrayals motivate support for LGBTQ+ rights cross-nationally (Ayoub and Garretson, 2017). Moreover, parasocial LGBTQ+ media contact reduces anti-LGBTQ+ prejudice and increases support for pro-LGBTQ+ policies (Miller et al., 2020).

Likewise, alternative insights from SIT and IET suggest if the mass public feels onesided civilian violence against marginalized groups is illegitimate, it may reflect poorly on their stigmatizing beliefs, even if minority group marginalization otherwise facilitates selfesteem (Harth et al., 2008). Dominant group or mass public members may emotionally regulate these psychic costs by reacting to violence against marginalized groups with sympathy and/or empathy (Branscombe and Miron, 2004), motivating the downstream adoption of prosocial attitudes toward marginalized groups (Harth et al., 2008). These propositions are consistent with evidence demonstrating empathetic feelings in response to observing LGBTQ+ discrimination elicit support for LGBTQ+ rights (Stotzer, 2009).

Some prior research implies the adoption of prosocial attitudes toward marginalized groups after violence exposure may be durable. The mass public has become increasingly inclusive toward LGBTQ+ community segments over several decades (Flores, 2014), suggesting the public may be durably receptive to sympathetic appeals after violence against LGBTQ+ group members. Indeed, Broockman and Kalla (2016) show a perspective-taking exercise can increase support for transgender anti-discrimination policies up to 3 months. Oskooii et al. (2021) show high-profile institutionalized discrimination against religious minorities can reduce mass support for policies negatively affecting targeted groups up to a year. Therefore, we might observe an empirical pattern consistent with Figure 1, Panel B, where the public adopts increasingly prosocial attitudes after indirect exposure to civilian violence against marginalized groups, and these attitudinal shifts are durable.



Figure 1: Stylized Expectations Concerning the Effect of Violence Against Marginalized Groups on Prosocial Attitudes. Horizontal lines denote prosocial attitudes toward marginalized groups (y-axis) over time (x-axis). The dotted line characterizes violence against a marginalized group.

3 The Fickle Prosocial Violence Response (FPVR) Model

However, we develop and present a *Fickle Prosocial Violence Response* (FPVR) model, which posits perceptibly illegitimate salient civilian violence against marginalized groups can motivate prosocial attitudes toward targeted groups. *But, these attitudinal shifts may be fickle* given reductions in event salience, the dispositional qualities of social group attitudes, and countervailing information in an otherwise discriminatory society.

Issue-Attention Cycle Theory posits the public may react to dramatic events highlighting ignored social issues, like violence against LGBTQ+ group members, in an initially proactive manner. However, attitudinal shifts seeking to resolve a social ill may not be sustainable when it becomes clear resolving the problem is difficult (e.g. reevaluating queerphobic beliefs offering a privileged status) and the problem becomes less salient over time (Downs, 1972). Prior research implies sympathetic media and elite messaging after violence must persist to generate sustainable prosocial responses (Zaller, 1992).

Moreover, prosocial attitudinal responses may be short-term impression management. Illegitimate violence rejected by society, media, and elites may motivate prosocial expressions toward the targeted group among the masses to save face (Harth et al., 2008), but may not result in long-term attitudinal shifts motivated by the difficult task of dismantling hierarchical social relations (Nguyen et al., 2021). Short-term impression management may not be capable of undermining predispositions toward marginalized groups rooted in pre-adult socialization (Sears, 1993; Kite et al., 2019). Long-term attitudinal shifts may also be undercut by countervailing pressure to adhere to queerphobic norms in an otherwise heteronormative society (Vuletich and Payne, 2019).

Framing theory may also help explain the potential absence of long-term prosocial attitudinal shifts. Story framing affects how the public assigns responsibility to an event and preferred policy and societal responses. Media outlets may adopt episodic or thematic frames in their news coverage. Episodic frames emphasize event-centered information with attention toward an individual's actions (e.g. the violent perpetrator) whereas thematic frames emphasize broader problems (e.g. queerphobia) (Iyengar, 1994). Ott and Aoki (2002) and Zahzah (2019) posit media frames of prominent instances of violence against LGBTQ+ people, such as Matthew Shepard's murder and the Pulse massacre, often emphasize the perpetrator's gratuitous violence instead of societal heteronormativity. These episodic frames may allow mass public members to simply express prosocial attitudes toward LGBTQ+ to absolve oneself of short-term guilt but lose sight of reflecting over their quotidian role facilitating a heteronormative society in the long-term (Ott and Aoki, 2002).⁸

In summary, an observable implication of the theoretical synthesis informing the FPVR model is that indirect exposure to *salient* and *sympathetic* messaging from media and elites after violence against LGBTQ+ group members may encourage the adoption of prosocial attitudes toward LGBTQ+ community segments. But, the adoption of prosocial attitudes toward LGBTQ+ group members may not be long-lasting. Therefore, we may observe an empirical pattern consistent with the solid line on Figure 1, Panel C. **H1:** Indirect exposure to civilian violence against LGBTQ+ group members will initially increase prosocial attitudes toward LGBTQ+ group members. **H2:** But, indirect exposure to civilian violence against LGBTQ+ community segments will not produce sustainable increases in prosocial attitudes.

⁸Moreover, if the violence is a mass shooting, conservative outlets, like Fox News, may emphasize gun rights, reducing sustained discussion of violence against LGBTQ+ group members that may motivate long-term prosocial belief adoption (Cassino, 2016).

Prior evidence corroborates the *FPVR* model. High-profile anti-Black police violence increased prosocial attitudes toward Black people, but these attitudinal shifts rebounded to the pre-violence equilibrium within weeks or months (Tuch and Weitzer, 1997; Chudy and Jefferson, 2021; Nguyen et al., 2021). Similarly, Birkland and Lawrence (2009) demonstrate Columbine immediately increased gun control support, but the increase was not sustainable.

3.1 Individual-Level Heterogeneity

Shared Marginalization. Group Empathy Theory posits marginalized group members who possess similar discriminatory experiences support each other (Sirin et al., 2021). Crossgroup support may be more likely if the discrimination a particular group experiences is perceptibly shared (Cortland et al., 2017). Members of other subjugated groups (e.g. nonwhites, women), may perceive similarities between their experiences and those of LGBTQ+ group members, especially with regard to targeted violence. Indeed, the Introduction shows women and non-whites have been historically subject to targeted violence in a conceivably similar manner as LGBTQ+ group members. Thus, group members discriminated against on other dimensions, like race and/or gender, may be more inclined to respond prosocially toward LGBTQ+ group members after indirect exposure to violence against LGBTQ+ community segments.

Political Liberalism. Relative to conservatives and moderates, liberals are less socially conservative concerning sexuality and gender and are more acceptant of marginalized social groups. Indeed, liberals are more favorable toward LGBTQ+ community segments and pro-LGBTQ+ policies (Flores, 2014). Conservatives are more likely to adopt anti-LGBTQ+ beliefs in response to threatening anti-LGBTQ+ elite rhetoric while liberals are resistant to such rhetoric (Górska and Tausch, 2022). Liberals are also more inclined to respond prosocially toward marginalized groups in response to high-profile state violence against said groups, while moderates and conservatives respond relatively apathetically (Reny and Newman, 2021). Therefore, liberals may be more likely than conservatives to adopt prosocial

attitudes toward LGBTQ+ group members in response to violence against LGBTQ+ people.

Geographic Context. Individuals living in areas with a higher composition of LGBTQ+ people may be more likely to come into contact with LGBTQ+ group members and develop relatively strong social ties with LGBTQ+ people (Tadlock et al., 2017). Harrison and Michelson (2019) identify consistent evidence contact with LGBTQ+ group members motivates prosociality toward different LGBTQ+ community segments. Given individuals living in areas with more LGBTQ+ people may be dispositionally favorable toward the LGBTQ+ community (Thompson, 2022), they may be more inclined to adopt prosocial attitudes toward LGBTQ+ community segments after high-profile civilian violence against LGBTQ+ group members. Indeed, prior research shows individuals living in LGBTQ+ geographic contexts resist exogenous anti-LGBTQ+ elite rhetoric (Górska and Tausch, 2022).

In summary, **H3a-c:** indirect exposure to civilian violence against LGBTQ+ group members will be more likely to motivate prosocial attitudes toward LGBTQ+ group members among: **a**) non-whites and women relative to whites and men; **b**) liberals relative to moderates and conservatives; **c**) individuals living in geographic contexts with more LGBTQ+ people relative to those living in contexts with less LGBTQ+ people.

3.2 Event-Level Salience Heterogeneity

The FPVR model implies violent events must be sufficiently *salient*, that is, covered by media and paid attention to by the public, to generate attitudinal shifts toward targeted groups (Downs, 1972; Zaller, 1992; Birkland, 1998). Indeed, prior studies demonstrating mass attitudinal shifts after US violent events are analyzing high-profile events (Tuch and Weitzer, 1997; Sigelman et al., 1997; Birkland and Lawrence, 2009; Reny and Newman, 2021). Moreover, prior research informing the FPVR model's assumptions suggests attitudinal shifts decay with reduced salience (Tuch and Weitzer, 1997; Birkland and Lawrence, 2009; Chudy and Jefferson, 2021; Nguyen et al., 2021). Importantly, salience *is not binary*. Violent Event A may be more salient than Violent Event B, but less salient than Violent



Fickle Prosocial Violence Response (FPVR) Model

Figure 2: Fickle Prosocial Violence Response Model

Event C, such that Event A does not sufficiently influence mass attitudes like Event C does. Thus, we may expect to observe an empirical pattern consistent with the dashed line on Figure 1, Panel C. **H4:** Initially more salient instances of civilian violence against LGBTQ+ group members will be more likely to motivate prosocial attitudes toward LGBTQ+ people than initially less salient instances of civilian violence against LGBTQ+ group members.

4 Event 1: The Pulse Massacre

Studies 1-2 evaluate the consequences of the Pulse massacre. The massacre occurred on June 12, 2016 at the Pulse LGBTQ+ nightclub in Orlando, Florida. The massacre was perpetrated by Omar Mateen, who pledged allegiance to ISIS. Mateen killed 49 and injured 53 clubgoers with a semi-automatic rifle.⁹ After taking hostages, Mateen was killed by the

⁹https://www.cnn.com/2016/06/12/us/orlando-shooter-omar-mateen/index.html

police. During the massacre, Pulse was hosting "Latin Night." 80% of victims were Latinx.¹⁰

The nation reacted sympathetically post-massacre. Republican Florida Governor Rick Scott expressed support for those affected while instituting a state of emergency. The Obama administration expressed condolences and ordered federal assistance to the police investigation and the community. In a press conference, Obama described the massacre as an "act of hate." Many on social media, including 2016 presidential election candidates, congresspeople, political figures, foreign leaders, and celebrities expressed condolences.

The massacre was salient. 90% of adults indicated they were closely following the incident immediately post-massacre (Figure A1). A survey during the massacre (June 10-26) suggests the public was aware of the shooting since it expressed more concerns about terrorism and gun violence post-massacre (Figure A4). Media coverage of topics related to Pulse, LGBTQ issues, and terrorism discontinuously increased post-massacre (Figure A2). Google searches related to Pulse, LGBTQ issues, and terrorism peak when the massacre occurs (Figure A3). Media coverage and Google searches related to these topics were either declining or limited pre-massacre, suggesting anticipatory effects are unlikely to drive attitudinal shifts toward LGBTQ+ issues or people post-massacre. However, coverage and searches decline to their pre-incident levels by July, implying salience is fleeting.

The massacre was not simply interpreted as a terror attack, but targeted, illegitimate, anti-LGBTQ+ violence.¹¹ 70-85% of adults believed the shooting was a hate crime (Figure A5).¹²

Therefore, consistent with the FPVR model, the mass public may respond prosocially to the perceptibly illegitimate Pulse massacre given the event's salience and concomitant

¹⁰The massacre's victims spanned the LGBTQ+ spectrum, but gay men may have been centered in the media post-massacre (Ramirez et al., 2018). Although this might mean the massacre was not interpreted as violence against a broader LGBTQ+ community, this is not a shortcoming with our analysis, but with how society interprets violence against LGBTQ+.

¹¹Omar Mateen was not explicitly motivated by anti-LGBTQ attitudes. Mateen randomly targeted nightclubs to inflict mass casualties (see: https://www.nbcnews.com/feature/nbc-out/what-really-happened-night-pulse-n882571). However, the mass public *perceived* the massacre as an anti-LGBTQ+ hate crime regardless of Mateen's motive (Figure A5).

¹²See Appendix A.6 and A.10 for details on Figure A5 data.

sympathetic response from both the media and political elites. But, given reduced media coverage and attention to the event over time, attitudinal responses may be short-lived.

4.1 Study 1: TAPS

4.1.1 Data and Design

Study 1 uses The American Panel Survey (TAPS, Wave 55), to assess if exposure to violence against LGBTQ+ people motivates support for policies benefiting LGBTQ+ community segments. TAPS is a monthly online survey administered by the Weidenbaum Center, with national probability sampling conducted by GfK/Knowledge Networks.

The outcome is same-sex marriage support (*SSM support*). SSM is an important LGBTQ+ rights dimension and it implicates multiple LGBTQ+ community segments. Gay, lesbian, and bisexual people who want to marry a same-sex partner benefit from legalized SSM. Transgender people who have not changed their "legal" gender but seek to marry their partner in heterosexual romantic relationships, in addition to transgender people in same-gender relationships, would benefit from legalized SSM.¹³ SSM approval is near-unanimous among LGBTQ+ people. 60% of LGBTQ+ people say SSM should be a priority even if it takes attention from other issues.¹⁴ TAPS asks respondents if they "generally support or oppose same-sex marriage," with an option to indicate "no opinion."¹⁵ We measure *SSM support* as a binary indicator equal to 1 if the respondent indicates they support SSM and 0 otherwise.

The independent variable is being interviewed after the Pulse massacre (*post-Pulse*). TAPS was fielded between 06/08/2016-07/08/2016. Pulse occurs on 06/12/2016, allowing us to implement an unexpected-event-during-survey-design (UESD) with TAPS comparing *SSM support* for respondents interviewed pre- and post-Pulse (Muñoz et al., 2020). *Post-Pulse* is a binary indicator equal to 1 if a respondent is interviewed after 06/12/2016. Since we cannot be certain respondents perceived the massacre, the *post-Pulse* coefficient is interpreted as an

¹³https://transequality.org/issues/resources/marriage-equality-and-transgender-people

¹⁴https://www.pewresearch.org/social-trends/2013/06/13/a-survey-of-lgbt-americans/

¹⁵See Section B.1 for outcome measurement details.

"intent-to-treat" (ITT) effect. However, Figures A1-A4 suggest the public was attentive to the massacre. Moreover, TAPS respondents are more likely to believe ISIS is an important issue *post-Pulse* (Figure B27), suggesting they "received the treatment" since the massacre's perpetrator pledged fealty to ISIS. If **H1** is supported, the *post-Pulse* coefficient would be *positive*.

In the absence of internal attention checks, we truncate our TAPS sample to those who completed the survey in a "reasonable duration" of time to account for online survey respondent inattentiveness, which may produce low quality responses attenuating associations of interest. See Section B.3.1 for more details and evidence truncation does not affect our empirical conclusions or TAPS' representativeness. After truncation, TAPS contains N = 1142respondents, with 682 (60%) interviewed before Pulse and 460 after (40%).

We demonstrate the *post-Pulse* coefficient is insulated from bias by validating UESD identification assumptions. The first assumption is ignorability. "Treatment" should be independent of potential outcomes conditional on random sampling. Thus, respondents interviewed pre and *post-Pulse* should be compositionally similar. Figure 3, Panel A supports the assumption. Respondents interviewed *post-Pulse* are compositionally similar to respondents interviewed pre-Pulse across 20 baseline covariates except age (see Section B.4 for baseline covariate measurement), a finding consistent with multiple testing.

Excludability is another UESD identification assumption: differences between respondents interviewed pre- and *post-Pulse* should be the sole consequence of the massacre. The "treatment" is not just the massacre, but collateral media attention. However, outside the massacre, there are no punctuated moments of media attention over LGBTQ+ issues or violence against LGBTQ+ people during the month TAPS was fielded (June, Figures A2 and A3), suggesting the absence of simultaneous events motivating pro-LGBTQ+ attitudes.

Additionally, it is unlikely preexisting *SSM support* time trends are driving the result (Muñoz et al., 2020). We subset TAPS to the pre-Pulse period and assess the placebo "effect" of being interviewed after the median pre-treatment date and find null results (Table B4).



Figure 3: SSM Support Increases post-Pulse. Panel A displays respondent covariate balance pre- and post-Pulse. Panel B characterizes the post-Pulse effect on SSM support with and without covariate adjustment. Panel C displays falsification tests characterizing the unadjusted post-Pulse effect on LGBTQ+-irrelevant outcomes. Black coefficients are statistically significant, grey otherwise. Estimates use population weights. All covariates scaled between 0-1. 95% CIs displayed from HC2 robust SEs. See Tables B5, B6, and B7 for regression tables characterizing the coefficients.

4.1.2 Results

Consistent with H1, respondents interviewed *post-Pulse* are 13 and 10 percentage points more likely to support SSM without and with covariate adjustment (p < 0.01, p < 0.05, Figure 3, Panel B). These coefficients are 20-26% of the outcome standard deviation.

Our results are robust. Our findings are likely not driven by secular dynamics outside the massacre. Falsification tests on treatment-irrelevant outcomes such as support for increasing taxes, common core, a citizenship pathway, abortion, the Keystone pipeline, ACA repeal and emission caps are null (Figure 3, Panel C). These tests suggest chance imbalance on age does not implicate balance on policy preferences.¹⁶ Given the close association between socially conservative religious beliefs like abortion restrictionism and SSM opposition (Uecker and Froese, 2019), the null effect of *post-Pulse* on abortion support on Figure 3 Panel C suggests our results are not driven by secular shifts in social conservatism or religiosity.¹⁷ The results

¹⁶Age imbalance may not induce bias since age is unrelated to *SSM support* in TAPS, so it does not explain *joint* treatment and outcome variation (Table B6).

¹⁷SSM and abortion support are only moderately correlated (Pearson's rho = 0.52), suggesting SSM support is explained by other factors, like the Pulse massacre, independent of dispositional religiosity or social conservatism. Religiosity is also constant pre- and *post-Pulse* (Figure 3, Panel A), further suggesting religiosity does not drive our results.



Figure 4: The Influence of Pulse on SSM Support Attenuates Over Time. X-axis is days cut from moment of Pulse massacre after the massacre (with days after intact). Y-axis is the *post-Pulse* coefficient. 95% CIs from robust SEs. See Table B8 for regression table characterizing reported coefficients in this figure. See Section B.8.5 for control covariate coefficients.

are not driven by outcome item non-response since non-response is balanced pre- and post-Pulse (Table B3). The results are not driven by seasonal trends, Pulse's influence is unique to 2016. 3 surveys fielded in June 2012, 2013, and 2017 show the influence of being interviewed after the massacre's calendar day on *SSM support* is null (Figure B31), suggesting no secular dynamics intrinsic to the month of June that could explain our findings (e.g. Pride Month). Our findings are robust to smaller bandwidths less susceptible to secular temporal trends (Figure B32). Finally, given we are deriving intent-to-treat coefficients, we test if *post-Pulse* is heterogeneous by political interest or news consumption. We do not find heterogeneity (Section B.15). This is not concerning since 90% of the public was following the shooting (Figure A1), suggesting high treatment reception regardless of dispositional political or media interest.

4.1.3 Temporal Persistence

We test **H2** by assessing if the influence of Pulse on *SSM support* is temporally durable. We remove observations in the days immediately *post-Pulse* but not after those days, and re-analyze the influence of being surveyed *post-Pulse*. The logic is that respondents interviewed immediately *post-Pulse* may be the most susceptible to shifting attitudes toward LGBTQ+

community segments. Removing them may help us evaluate attitudinal decay by comparing respondents interviewed just before and some days after Pulse. After removing respondents interviewed between 1-10 days *post-Pulse*, the influence of being interviewed *post-Pulse* on *SSM support* is null (Figure 4).¹⁸ Therefore, temporal attenuation is quick relative to prior studies demonstrating attitudinal shifts lasting several months to a year (Broockman and Kalla, 2016; Oskooii et al., 2021). Consistent with **H2**, the initial increase in *SSM support post-Pulse* was not durable.

4.1.4 Individual-Level Heterogeneity

We test **H3a-c** by assessing if the *post-Pulse* coefficient is larger among: a) non-whites relative to whites and women relative to men; b) liberals relative to moderates and conservatives; and c) individuals living in states with a higher proportion of LGBT-identifying people and living in counties with a higher density of same-sex couples relative to individuals who live in areas with less LGBT-identifying people and same-sex couples.¹⁹ Inconsistent with **H3a-c**, *post-Pulse* does not appear heterogeneous by marginalized group membership, liberalism, and LGBTQ+ geographic context (Table B15). These findings suggest the massacre had a *largely homogeneous initial influence on mass attitudes*.

4.2 Study 2: PI S-IAT Data

4.2.1 Data and Design

Study 2 examines if the public adopts positive attitudes toward LGBTQ+ community segments *post-Pulse*. We use Project Implicit (PI) data on US respondents self-selecting into

 $^{^{18}1/20}$ covariates are imbalanced after cutting 2, 4, 6, 14, 16, 21, 22 days *post-Pulse* (Table B18), suggesting Figure 4's results are not driven by imbalance.

¹⁹We use 2016 Gallup data on over 1 million U.S. respondents to identify the proportion of each state's population identifying as "lesbian, gay, bisexual or transgender." (see: https://news.gallup.com/poll/201731/lgbt-identification-rises.aspx) We use 2010 Census data to identify same-sex couple density (the number of same-sex couple households per 1000 households in a county, see: https://williamsinstitute.law.ucla.edu/visualization/lgbt-stats/?topic=SS& showCounties=true#density). We merge these state and county-level covariates to the TAPS data by using respondent zipcode information.

and completing an internet survey in 2016 asking questions on their explicit and implicit attitudes toward gay people via PI's Sexuality Implicit Association Test (S-IAT, N = 43,950).²⁰ On average, 175 U.S. respondents completed the PI S-IAT survey each day during 2016.²¹ For information on S-IAT sample composition and representativeness, see Section C.1.

The outcomes are the S-IAT *D*-score, straight bias, and heterocentrism. The S-IAT calculates normalized averages of how quickly respondents associate negative/positive attributes to gay/straight people relative to negative/positive attributes to straight/gay people in the form of a *D*-score. The *D*-score ranges from -2-2. Higher values suggest implicit bias against gay people (i.e. associating negative attributes to gay people) (Greenwald and Lai, 2020).²²

Given indirect measurement, the *D*-score may be less influenced by impression management to be perceived as pro-gay post-massacre (Greenwald and Lai, 2020). Therefore, we can assess relatively quick, negative, emotional responses (i.e. System 2 responses) to gay people in addition to more deliberate evaluations of gay people (i.e. System 1 responses) (Greenwald and Lai, 2020). Although the IAT is not insulated from introspection, the modest correlation between the *D*-score and explicit bias suggests the IAT measures attitudes that are difficult to manipulate. Therefore, the *D*-score is valuable since we can demonstrate even temporary prosocial attitudinal shifts may not be impression management. The *D*-score is well-established and associated with objective covariates characterizing subordination (Ratliff and Smith, 2021).

Heterocentrism and straight bias are explicit anti-gay bias measures. Heterocentrism is the difference in 10-point feeling thermometers for straight and gay men. Straight bias is a 7 point measure from "I strongly prefer gay to straight people" to "I strongly prefer straight to gay people." The D-score, straight bias, and heterocentrism are rescaled between 0-1.

Although *heterocentrism* is explicitly about gay men, and *straight bias* is implicitly about

²⁰Data available here: https://osf.io/yjqmw/. See https://implicit.harvard.edu/implicit/ education.html for Project Implicit information.

 $^{^{21}}$ We exclude respondents interviewed after 09/08/2016 due to order effects since the S-IAT measurement changes from 188 to 200 trials by cutting a task block at that moment.

 $^{^{22}}$ See Section C.3 for more *D*-score measurement details.

gay men, the *D*-score captures attitudes toward gay men and lesbians. In effect, the *D*-score implicates gay men, lesbians, and bisexuals (and transgender people in same-gender relationships). Moreover, even if our Study 2 outcomes are limited when it comes to measuring attitudes toward the broader LGBTQ+ community (e.g. transgender people), attitudes toward gay people are correlated with attitudes toward transgender people (Norton and Herek, 2013), which may be pronounced given the massacre affected transgender people.²³ Therefore, our Study 2 outcomes implicate large LGBTQ+ community segments. Given the outcomes characterize negative attitudes, if **H1** is supported, *post-Pulse* should be *negative*.

We use a UESD with the S-IAT to evaluate how anti-gay attitudes shifted *post-Pulse*. Given the large number of individuals taking the S-IAT daily, we estimate the influence of taking the S-IAT *post-Pulse* using respondents taking the S-IAT 5-50 days pre- and post-massacre in addition to the full 2016 sample between January-September.

We validate the UESD ignorability identification assumption. Unlike Study 1, respondents are not sampled, but self-select, into the S-IAT. Therefore, sample composition may shift due to external events or secular trends. We expect respondents surveyed shortly preand post-massacre will be compositionally similar. However, respondents may be increasingly dissimilar in samples including respondents taking the survey well before or after the massacre. Figure C34 verifies our expectation. For 5-20 day bandwidth samples (Panels A-D), there is statistical imbalance on respondent characteristics pre- and post-Pulse on 1-2/12 baseline covariates. For 25-50 day bandwidth samples, there is imbalance on 3-7 covariates (Panels E-J). Given the 15 and 20-day bandwidth samples are only imbalanced on race, we prioritize interpreting the influence of *post-Pulse* on anti-gay attitudes using these samples. These findings suggest our coefficient estimates, particularly for the 15 and 20-day bandwidth samples, are relatively insulated from omitted variable bias.²⁴

²³https://www.advocate.com/crime/2016/6/17/pulse-survivor-stop-being-shady-and-messy-just-love-one-another-video

²⁴Importantly, like Study 1, religiosity is constant pre- and *post-Pulse*, suggesting secular sample composition shifts in social conservatism are not driving our results.



Figure 5: Influence of *post-Pulse* on Anti-Gay Attitudes. The x-axis is the sample bandwidth. The y-axis is the *post-Pulse* coefficient. All covariates rescaled between 0-1. 95% CIs displayed from robust SEs. See Sections C.7.2 and C.7.3 for corresponding regression tables.

4.2.2 Results

Figure 5 displays *post-Pulse* ITT coefficients where the outcome is the *D-score*, straight bias, and heterocentrism. In the 15 and 20-day sample bandwidth estimates, respondents surveyed *post-Pulse* have a lower *D-score* (-0.01, p < 0.10) and heterocentrism (-0.01, p < 0.01), equivalent to 7% and 8% of the respective outcome standard deviations pre-Pulse. Although small, these coefficients are reasonable, likely underestimated, and substantively important vis-a-vis the target population (see Section C.6).

The massacre does not appear to statistically reduce *straight bias* except in sample bandwidths with higher covariate imbalance (e.g. 25-50 days). Given *straight bias* is highly explicit, the absence of a reliable shift in *straight bias post-Pulse* may be a function of impression management on part of respondents disposed against LGBTQ+ whose attitudes may otherwise shift in favor of LGBTQ+ through indirect bias measurement (Greenwald et al., 1998). In sum, we find additional support for **H1** in Study 2.

We conduct several robustness checks. Preexisting time trends are not driving our results (Section C.8). We rule out if systematic temporal trends near June motivate prosocial attitudes toward gay people other than the massacre (Section C.9). We rule out if our findings are due to a secular attitudinal trend in favor of marginalized groups (Section C.10). We also rule out if respondent self-selection generates sorting bias (Section C.13).



Figure 6: The Influence of *Post-Pulse* on Reducing Anti-Gay Attitudes Attenuates Over Time. X-axis is days cut from moment of Pulse massacre after the massacre (with 15 days after kept intact). Y-axis is the *post-Pulse* coefficient. All estimates from models adjusting for controls. 95% CIs from robust SEs. See Tables C33, C34, C35, and C36 for tables characterizing the displayed coefficients. See Sections C.7.8 and C.7.9 for control coefficients.

4.2.3 Temporal Persistence

We assess if the *D*-score and heterocentrism decrease is sustainable. Consistent with H2, descriptive statistics suggest anti-gay attitudes decreased post-Pulse, but rebounded to pre-Pulse levels around August (Figure C33). We conduct a formal test of the sustainability of attitudinal shifts post-Pulse and compare S-IAT respondents surveyed 15 days pre-Pulse to those surveyed 15 days after 1-72 days post-Pulse (leaving at least 15 days up to the end of the post-treatment sample in the 2016 S-IAT data). This exercise allows us to compare individuals surveyed prior to Pulse to those surveyed some time away from Pulse at multiple time intervals. Respondents in time intervals that cut more days post-Pulse are temporally further from the massacre and potentially more subject to attitudinal decay in pro-gay beliefs. Figure 6 demonstrates the *D*-score and heterocentrism decrease was sustained up to 40 days post-Pulse. However, after 40 days, post-Pulse attenuates toward 0.2^5 Although attitudinal shifts last 40 days, these shifts are still much shorter than prior studies demonstrating long-term attitudinal shifts toward marginalized groups after external stimuli (Broockman and

 $^{^{25}}$ After cutting 40 days *post-Pulse*, there is covariate imbalance, but this does not invalidate Figure 6. After covariate adjustment, the *post-Pulse* coefficients attenuate toward zero, suggesting temporal attenuation occurred *earlier* than our results suggest (Section C.12).

Kalla, 2016; Oskooii et al., 2021). Consistent with **H2**, Study 2 suggests the massacre motivated prosocial beliefs, but not durably.

4.2.4 Individual-Level Heterogeneity

We test **H3a-c** by assessing if the *post-Pulse* coefficient is larger among non-whites, women, liberals, and individuals living in geographic contexts with more LGBTQ+ people.²⁶ Inconsistent with **H3a-c**, we find the massacre's influence is relatively homogeneous. *Post-pulse* is not stronger for non-whites, women, liberals, or respondents in geographic contexts with more LGBTQ+ people (Tables C50-C51).

4.3 Mitigating Bundled Treatment Concerns

It is unclear if respondents adopted prosocial beliefs toward LGBTQ+ community segments because the Pulse massacre was a terror attack or attack against Latinxs instead of perceptibly anti-LGBTQ+ violence. We mitigate these concerns with several tests and evidence outlined in detail in Section A.12. We summarize these tests and evidence here. First, other terror attacks do not motivate pro-LGBTQ+ beliefs and Pulse did not motivate antipathy toward groups stereotypically associated with radical Islamic organizations, mitigating concerns our results are driven by the massacre being a terror attack. Second, other attacks against Latinxs do not motivate pro-LGBTQ+ beliefs and Pulse did not motivate positive attitudes toward Latinxs, mitigating concerns our results are driven by the massacre being violence against Latinxs. Third, we show the public was disproportionately attentive to LGBT topics *post-Pulse* relative to terrorism- and/or Latino-related topics, implying the public primarily perceived the event as anti-LGBTQ+ violence.

Our tests do not entirely mitigate the bundled treatment problem. Our results may be due to the *combination* of circumstances associated with Pulse. Therefore, we conceptually

 $^{^{26}{\}rm Geographic}$ context is measured like Study 1. We use respondent county data in the S-IAT to merge in information on LGBTQ+ geographic context.



Figure 7: Media Coverage of Gay-Related Content in 1998. Panels A/B display the number of NYT/Washington Post gay-related articles (y-axis) by month (x-axis). Panel C displays the number of gay-related articles related to Shepard or anti-gay violence by month. Dashed vertical line denotes the period Shepard is murdered.

replicate Studies 1-2 by assessing the influence of instances of violence against LGBTQ+ group member(s) in Studies 3-4 that were not terror attacks nor attacks against non-whites.

5 Event 2: Matthew Shepard's Murder

Studies 1-2 may not be externally valid. The Pulse massacre is a unique instance of violence against LGBTQ+ people. It is the deadliest instance of violence against LGBTQ+, is the second deadliest mass shooting, has predominantly Latinx victims, was ISIS-inspired terrorism, and occurred after seminal gay rights victories (e.g. same-sex marriage). Therefore, it may be prudent to assess if a distinct instance of violence against LGBTQ+ group member(s) also motivates prosocial beliefs. Consequently, we examine how the murder of Matthew Shepard, a white gay Wyoming college student, by two white men, influenced beliefs toward homosexuality during a more homophobic temporal context.

On October 6, 1998, Shepard was brutally beaten by Aaron McKinney and Russell Henderson. The incident was heavily covered by national media (Loffreda, 2001). Shepard died six days later on October 12. The murder was salient and the nation reacted sympathetically. A bipartisan group of Congresspeople condemned the murder and expressed condolences. A vigil was held outside the US capitol on October 15, where thousands of people, including current and former Congresspeople and celebrities, paid respects to Shepard. Advocates note Shepard's murder engendered a "seismic shift in attitudes towards the LGBTQ community."²⁷ Indeed, a decade later, Congress passed the Matthew Shepard and James Byrd Hate Crimes Prevention Act, which expanded the power to prosecute sexuality hate crimes.

On the month of Shepard's murder, the number of gay-related news articles was 150% (NYT) and 172% (WashPo) of the Jan-Sep 1998 average (Figure 7).²⁸ Consistent with the FPVR model, media attention to Shepard's murder was immediately intense but quickly declined, suggesting attitudinal responses may be short-lived.

5.1 Study 3

5.1.1 Data and Design

To evaluate if Shepard's murder decreased anti-gay attitudes, we identify surveys with similar items characterizing attitudes toward gay people shortly before and after Shepard's murder.²⁹ We identify two representative CNN telephone polls asking respondents if they believe homosexuality is "morally wrong" (moral wrong) 4 months before and 2 days after Shepard's death (CNN Jun. 1998, N = 1016; CNN Oct. 1998, N = 1036).³⁰ We stack these datasets and identify overlapping controls from each survey.³¹ We then compare respondents interviewed after Shepard's murder (*post-Shepard*) to those before to assess if anti-gay violence exposure decreased the belief homosexuality is morally wrong, consistent with **H1**. We focus on surveys with the moral wrong outcome for 3 reasons. First, it is asked on three surveys after Shepard's murder (in 1998, 2001, 2004), allowing an assessment of long-term attitudinal shifts. Second, there are multiple pre-Shepard surveys with the same item, allow-

²⁸See Appendix D.1 for details on media data.

 $^{^{29} \}rm We$ use the search terms "homosexuality" OR "homosexual" OR "gay" in Roper iPoll between 1996-2000 to identify gay-related items around Shepard's murder.

 $^{^{30}}$ We found two other items that could serve as potential candidates for assessing the influence of Shepard's murder on attitudes toward LGBTQ+ group members, however, we do not use them for various reasons that we outline in Section D.4.

 $^{^{31}\}mathrm{See}$ Section D.2 for more details on sampling methodology.

ing placebo tests to rule out if *post-Shepard* effects are due to secular progressive attitudinal trends concerning homosexuality's morality. Third, *moral wrong* implicates large LGBTQ+ community segments. Lesbian, gay, bisexual, and transgender people may all partake in "homosexual" behavior. Given the outcome characterizes a negative attitude toward LGBTQ+ community segments, the *post-Shepard* coefficient would be *negative* if **H1** is supported.

Our approach has shortcomings we assuage. First, given the absence of auxiliary data on attention to the murder, we cannot be certain respondents "received the treatment." Therefore, we interpret *post-Shepard* as an ITT effect. However, Figure 7 suggests the murder received significant media attention such that it might shift mass attitudes.

Second, given possible differences in sampling between the two surveys, our statistical conclusions may be due to sample composition. Balance tests between the two surveys demonstrate limited baseline covariate imbalance (Figure 8, Panel A), suggesting sample composition may not drive our results.

Third, unlike Studies 1-2, we cannot assess an immediate effect of anti-gay violence exposure even though the two surveys were fielded near Shepard's murder. There are four months between the surveys with the *moral wrong* outcome (Jun.-Oct. 1998). Therefore, our *post-Shepard* estimates may be due to intervening factors or secular progressive time trends. However, there is no anti-gay violence with the level of media coverage Shepard's murder garnered in between the field periods (Figure 7). Crowdsourced evidence suggests the last prominent instance of anti-LGBTQ+ violence prior to Shepard's murder was not between June-October 1998, but on February 1997 (the Otherside Lounge Bombing).³² Indeed, between June-September 1998, there were *zero* New York Times articles related to anti-gay hate crimes. Conversely, on the month of Shepard's murder (October 1998), there were 17 NYT articles related to anti-gay hate crimes (Figure D39). Two other intervening factors in 1998 may explain our results: 1) President Clinton signing an executive order against sexual orientation discrimination and 2) Tammy Baldwin's House election (the first lesbian

³²https://en.wikipedia.org/wiki/History_of_violence_against_LGBT_people_in_the_United_States

congressperson). We provide evidence these events are unlikely explaining our *post-Shepard* coefficient estimates (Section D.10).

Moreover, we rule out if our results are due to secular outcome time trends by conducting a temporal placebo test and demonstrating *moral wrong* levels do not change between Apr. 1997-Jun. 1998 (Figure 8, Panel B).³³ These results suggest prominent pre-study events, such as Ellen DeGeneres' televised coming out in April 1997, are not driving our results. Despite Study 3's shortcomings, we believe the design provides sufficient complementary evidence to Studies 1–2 along with suggestive evidence our theory generalizes beyond Pulse.

5.1.2 Results

Consistent with H1, Figure 8, Panel B shows respondents interviewed *post-Shepard* were 10 percentage points less likely to report homosexuality is morally wrong with or without covariate adjustment, 20% of the outcome standard deviation (p < 0.001).

We conduct falsification tests on outcomes related to non-LGBTQ+ marginalized groups to rule out secular supportive trends toward marginalized groups driving our results (Figure 8, Panel C).³⁴ Only 4/18 outcomes are significant and the *post-Shepard* coefficient is not consistently in support of non-LGBTQ+ groups, suggesting no systematic secular trend driving our results (see Section D.7 for more details). Like Study 1, the null effects of *post-Shepard* on abortion support suggest our results are not driven by secular shifts in social conservatism and/or religiosity.

5.1.3 Temporal Persistence

To assess the persistence of attitudinal shifts *post-Shepard*, we identify 6 surveys between 1978-2004 where the *moral wrong* item was asked,³⁵ allowing us to evaluate trends in the public's belief homosexuality is morally wrong pre- and *post-Shepard*. We do not use the

 $^{^{33}\}mathrm{See}$ Section D.2 for more details on the temporal place bo test.

 $^{^{34}\}mathrm{See}$ Section D.8 for more details on falsification test outcomes.

 $^{^{35}\}mathrm{See}$ Section D.5 for details on the 6 surveys.



Figure 8: Respondents Interviewed *Post-Shepard* Were Less Likely To Believe Homosexuality is Morally Wrong. Panel A characterizes balance between respondents interviewed pre- and post-Shepard's murder. Black coefficients are significant, grey otherwise. Panel B characterizes a) the influence of being interviewed on June 1998 relative to April 1997 on the belief homosexuality is morally wrong (temporal placebo) and b) the influence of being interviewed *post-Shepard* on *moral wrong*. Panel C characterizes falsification tests assessing the influence of *post-Shepard* on non-LGBTQ+ group attitudes. 95% CIs displayed from robust SEs. See Tables D57, D58, D59, and D60 for regression tables on balance tests, the temporal placebo, the *post-Shepard* coefficient estimates, and falsification tests.

CNN June 1998 poll on Figure 9 in our assessment of temporal persistence (see Section D.6 for details as to why).

Figure 9 displays event study estimates comparing *moral wrong* levels in 5 surveys between 1978-2004 to a survey fielded prior to Shepard's murder in 1994. From 1978-1994, belief in *moral wrong* is remarkably stable. Respondents surveyed in 1994 are not statistically distinct from respondents surveyed in 1992 or 1978. Consistent with our initial temporal placebo test, these findings suggest an absence of progressive attitudinal trends toward gay Event Study (Moral Wrong)



Figure 9: Belief in *Moral Wrong* is Stable Between 1978-2004 With the Exception of the Moment Shepard Was Murdered. Reference study is the 1994 CNN poll. Color denotes the inclusion/exclusion of controls (age, education, gender, partisanship, race). Shaded estimate denotes Shepard's murder (Oct. 1998). All estimates use survey weights. All covariates scaled between 0-1. See Table D61 for a regression table characterizing this figure. 95% CIs displayed derived from robust SEs.

people prior to Shepard's murder. However, in October 1998, immediately after Shepard's murder, there is a statistically distinguishable decrease in *moral wrong*. But, the mass public's beliefs in *moral wrong* reverse to pre-murder levels by 2001 and 2004. Consistent with **H2**, our results suggest Shepard's murder motivated a decrease in negative beliefs concerning "homosexuals," but this decrease was not sustainable.

5.1.4 Individual-Level Heterogeneity

We test **H3a-b** by assessing if the *post-Shepard* coefficient is stronger among a) non-whites and women and b) Democrats.³⁶ Given the absence of a) county-level geographic data in the two 1998 CNN polls and b) state-level LGBT population information in the 1990s, we cannot test **H3c**. We find some evidence consistent with **H3a** (Table H73). Although there is no

 $^{^{36}\}mathrm{Data}$ on liberalism is unavailable in the 1998 CNN polls, but Democratic partisanship is an appropriate proxy given its' strong association with liberalism.

post-Shepard heterogeneity by gender, non-whites are less likely to believe homosexuality is morally wrong relative to whites *post-Shepard*. Whites are 7 percentage points less likely to believe homosexuality is morally wrong *post-Shepard* whereas non-whites are 22 percentage points less likely, 44% of the pre-Shepard outcome standard deviation. Likewise, we find evidence supporting **H3b** (Table H73). The *post-Shepard* effect appears driven by Democrats. Democrats are 22 percentage points less likely to believe homosexuality is morally wrong *post-Shepard*, whereas non-Democrats are 2 percentage points less likely.

6 Event 3: The Club Q Massacre

Study 4 mitigates two shortcomings with Studies 1-3. First, Studies 1-3 all analyze initially highly salient events (i.e. high media coverage, attention). However, consistent with **H4** and the *FPVR* model, relatively initially less salient violent events may be less likely to motivate prosocial attitudes toward LGBTQ+ community segments. Study 4 allows us to evaluate the consequences of indirect exposure to a putatively high-profile, but relatively initially less salient, instance of violence against LGBTQ+ group members: the 2022 Club Q massacre. Consequently, Study 4 allows us to test **H4** and broader *FPVR* model implications related to initial event salience. Second, the outcomes in Studies 1-3 do not explicitly reference broader LGBTQ+ segments beyond gays and lesbians (e.g. transgender people). Conversely, Study 4 not only examines the same Study 2 outcomes using the 2022 PI S-IAT survey, but additional outcomes characterizing negative attitudes toward transgender people in the 2022 PI Transgender Implicit Association Test (PI T-IAT) survey.³⁷ Therefore, Study 4 allows us to examine the consequences of violence against LGBTQ+ group members on mass attitudes explicitly related to transgender people, a small, politicized, population (Lewis et al., 2022).

On November 19, 2022, in Colorado Springs, CO, Anderson Aldrich entered an LGBTQ+ nightclub, Club Q, and killed five clubgoers, including two trans people, while injuring 25

³⁷PI started collecting transgender attitude data in 2020 (https://osf.io/fb29q/).

others with an AR-15-style rifle.³⁸ Aldrich was eventually incapacitated by clubgoers and apprehended by police. Evidence suggests the violence was bias-motivated. Aldrich pleaded "no contest" in court to two hate crime charges.³⁹

The media and some elites reacted sympathetically to the violence. President Biden and Transportation Secretary Buttigieg immediately expressed condolences.⁴⁰ However, unlike the Pulse massacre and Shepard's murder, the elite response was relatively polarized. Buttigieg blamed the shooting on growing Republican anti-LGBTQ+ rhetoric.⁴¹ Tucker Carlson and several right-wing commentators blamed the violence on purported "grooming" activity from LGBTQ+ people.⁴² Republican politicians who expressed condolences were criticized for simultaneously engaging in anti-LGBTQ+ rhetoric.⁴³ LGBTQ+ advocates noted a rise in queerphobic posts across social media platforms post-shooting.⁴⁴

Moreover, relative to Shepard's murder and the Pulse massacre, the Club Q massacre was less salient. First, there were less NYT articles related to the Club Q massacre two months after the event relative to Shepard's murder and the Pulse massacre (Figure E42). Second, regression discontinuity-in-time estimates suggest that although online articles on topics related to mass shootings, the LGBT community, and hate crimes discontinuously increased after Club Q, there were more online articles on topics related to mass shootings and the LGBT community after Pulse (Figures E43-E44, Table E64). Third, Google search data demonstrates there was more attention to mass shootings, LGBT people, and LGBT hate crimes immediately during Pulse relative to immediately during the Club Q massacre

³⁸https://www.cnn.com/2022/11/20/us/colorado-springs-shooting-gay-nightclub

³⁹https://www.pbs.org/newshour/nation/club-q-shooter-who-killed-5-gets-life-in-prison ⁴⁰https://www.denver7.com/news/local-news/we-are-devastated-officials-react-to-deadlymass-shooting-at-club-q-in-colorado-springs

⁴¹https://www.yahoo.com/video/pete-buttigieg-says-political-attacks-145452238.html

⁴²https://www.nbcnews.com/tech/internet/right-wing-influencers-media-double-antilgbtq-rhetoric-wake-colorado-rcna58371

 $^{^{43} \}tt https://www.durangoherald.com/articles/lauren-boebert-defends-her-past-anti-lgbtq-and-anti-trans-tweets/$

⁴⁴See: https://www.isdglobal.org/digital_dispatches/groomer-discourseintensifies-and-neo-nazis-celebrate-in-wake-of-colorado-springs-attack/ and see: https://apnews.com/article/technology-shootings-business-social-media-colorado-75a3c597a60dca0f116d5deb6a6c1a6b

(Figure E45). Therefore, consistent with the FPVR model and H4, although Club Q was relatively high-profile, its' lower-profile status vis-a-vis Pulse and Shepard's murder suggests it may be less likely to initially shift mass attitudes.

6.1 Study 4

6.1.1 Data and Design

We use data on U.S. respondents self-selecting into the 2022 PI S-IAT (N = 184, 824, 506 daily average respondents) and T-IAT (N = 85, 303, 233 daily average respondents) surveys. See Section E.1 for information on S-IAT and T-IAT sample composition and representativeness.

The S-IAT outcomes are the same as Study 2's (anti-gay *D-score*, *heterocentrism*, *straight bias*). The three T-IAT outcomes are similar but slightly different. The anti-trans *D-score* is measured by assessing the speed by which respondents associate negative/positive attributes (words) to images of trans/cis celebrities. Higher values suggest respondents associated negative/positive attributes to trans people faster/slower than cis people. *Ciscentrism* measures relative warmth toward cisgender people vis-a-vis trans people. *Cis bias* is a 7-point scale measuring preferences for cisgender relative to trans people. See Section E.2 for more T-IAT outcome measurement details. Prior research finds the T-IAT outcomes are correlated with anti-trans policy preferences (Axt et al., 2021). All outcomes are rescaled between 0-1.

The main independent variable is *post-Club* Q, an indicator equal to 1 if a respondent self-selects into the S-IAT or T-IAT after November 19, 2022. The *post-Club* Q coefficients will be *negative* if prosocial attitudes increase *post-Club* Q.

We implement another UESD, estimating the influence of *post-Club* Q 5-40 days in 5-day intervals post-massacre.⁴⁵ We assess covariate balance for these bandwidth samples between respondents taking the S-IAT/T-IAT pre- and *post-Club* Q (Figures E46-E47).⁴⁶ Covariate

 $^{^{45}}$ There are no data after 40 days *post-Club* Q since the 2022 surveys end on December 2022.

⁴⁶Baseline control covariates are measured like Study 2.



Figure 10: Influence of *post-Club Q* on Anti-Gay, Anti-Trans Attitudes. The xaxis is the bandwidth sample (1-40 days). The y-axis is the *post-Club Q* coefficient. Panels characterize different outcomes. The top/bottom 3 panels characterize estimates from the 2022 PI S-IAT/T-IAT data. Black coefficients are from models adjusting for controls, grey otherwise. 95% CIs displayed from HC2 robust SEs. See Tables E65-E66 for regression tables characterizing these estimates.

imbalance increases as sample bandwidth increases, likely due to unobservable secular trends. Therefore, we primarily interpret the 20- and 15-day bandwidth samples in the S-IAT and T-IAT respectively, where there is the least imbalance (4/12 and 1/12 covariates imbalanced respectively).

6.1.2 Results

The post-Club Q coefficient is null across all outcomes in the S-IAT/T-IAT 20/15-day bandwidth samples (Figure 10). Although post-Club Q coefficients in larger bandwidth samples suggest a decrease in the anti-trans and anti-gay *D-score* (e.g. the 40-day bandwidth samples), these estimates should be viewed skeptically given they possess high covariate imbalance and are more likely to be perturbed by unobservable secular trends (Figures E46-E47). Consistent with the FPVR model and H4, less salient violent events like Club Q do not motivate attitudinal shifts like more salient events (e.g. Pulse or Shepard's murder).

6.1.3 Individual-Level Heterogeneity

We test H3a-c and assess if the *post-Club* Q coefficient is larger among a) non-whites and women, b) liberals, and c) individuals living in geographic contexts with more LGBTQ+ group members using the 20- and 15-day bandwidth samples for the S-IAT and T-IAT. We find limited heterogeneity across these characteristics (see Tables E67-E68). The only statistically significant heterogeneity we identify is that the *post-Club* Q coefficient is negative and stronger among women for the *Cis Bias* outcome (Table E68). However, we do not identify heterogeneity by gender in the S-IAT data or the other two T-IAT outcomes. Therefore, we interpret the influence of *post-Club* Q as largely homogeneous.

6.2 Evidence From Less Salient Violent Events

A limitation with Study 4 is that, although the Club Q massacre was less salient than Pulse and Shepard's murder, the null results may be due to the arguably more polarized temporal context given the recent rise of Republican anti-LGBTQ+ rhetoric and policies. Anti-LGBTQ+ laws implemented in Republican states (Figure E48) and right-wing anti-LGBTQ+ protests have increased in the past few years (Figure E49). Indeed, prior research shows LGBTQ+ mass attitudes may entrench in polarized contexts (Lewis et al., 2022). The FPVR model also corroborates this limitation, since sympathetic responses by bipartisan elites may be necessary to motivate prosocial mass attitudes (Figure 2).

To circumvent this limitation, we use crowdsourced data on less salient violent incidents against LGBTQ+ people between 2010-2022 and evaluate the influence of these events on prosocial attitudes.⁴⁷ We demonstrate the incidents outside of those in Studies 1-4 are significantly less salient (Figure F50). We identify 3570, 442, and 358 NYT article hits related

⁴⁷https://en.wikipedia.org/wiki/History_of_violence_against_LGBT_people_in_the_United_ States

to the Pulse massacre, Shepard's murder, and the Club Q massacre respectively (Figure F50, Panel C). Conversely, the next most salient violent incident against LGBTQ+ group members between 2010-2022 was Mark Carson's May 2013 murder with 30 hits (Figure F50, Panel B). Consistent with **H4**, other less salient violent incidents against LGBTQ+ group members outside those in Studies 1-4 have largely null effects on mass attitudes toward gay people (Figure F51). The few significant effects are not consistently in the same substantive direction, implying a random, unsystematic, causal process.

7 Limitations and Additional Robustness Checks

Our analyses have limitations. First, one issue with our analytic approach is that we use several distinct outcomes across different time periods while assuming they measure the same concept (i.e. prosocial LGBTQ+ attitudes). We believe this is an advantage since prosocial attitudes towards LGBTQ+ group members are *multidimensional* and not correlated with each other 1-to-1 (Flores, 2014). Moreover, queerness is *fluid* and is simultaneously defined and expressed differently over time (Lewis et al., 2022). Therefore, although the meaning of our outcomes may shift over time, our theory may continue to apply across temporal domains. In sum, our empirical approach helps demonstrate our theory is justifiable and broadly applicable by showing high-profile violence against LGBTQ+ people influences distinct prosocial attitudinal dimensions (e.g. policy preferences *and* affective attitudes toward *distinct* LGBTQ+ community segments) similarly across temporal domains.

Nevertheless, our outcomes capture the same concept. If our outcomes are measuring the same concept despite differences in measurement and temporal domain across studies, they should 1) be highly correlated consistently with each other *across several time periods*, and 2) have similar correlates over time. We show these criteria are met in Section G.

Second, although we provide evidence respondents likely perceived and responded to violence against LGBTQ+ people in a manner consistent with the FPVR model, we cannot be certain respondents "received the treatment." Future research should use designs encouraging stronger treatment reception (e.g. survey experiments) to assess if our analyses underestimate effects and/or temporal persistence. However, unlike designs offering stronger treatment reception, a (tragic) advantage of our design(s) is that they derive effects based on "real-world," externally valid events.

Third, our evidence has not tested all mechanisms consistent with the FPVR model's assumptions. Our design is advantageous in that we can assess the effects of violence on prosocial attitudes in an uncontrolled environment with plausible identification assumptions, undercutting demand effects or external invalidity. But, our data were not directly collected to test our hypotheses, making mechanism tests difficult. To the extent we can provide evidence for mechanisms outlined in the FPVR model (Figure 2), we show a) initial salience is necessary to motivate prosocial attitudes at the outset, b) declines in salience over time are concomitant with decay in prosocial attitudinal shifts, and c) there is limited support shared marginalization, ideology, and LGBTQ+ geographic context consistently moderates the initial adoption of prosocial attitudes.

Future research should test other *FPVR* model mechanisms (Figure 2). Psychological insights are promising. Violence exposure's influence on prosocial beliefs and their sustainability may be mediated through positive emotional responses toward marginalized groups (e.g. empathy, sympathy, anger, guilt) (Branscombe and Miron, 2004; Harth et al., 2008). Additionally, future research should assess how media frames condition the public's attitudinal responses. During Shepard's murder and Pulse, the media and elites framed the victims sympathetically (instead of unsympathetically). Concomitantly, prior research suggests the media used episodic frames focusing on perpetrator motivations instead of thematic frames emphasizing societal queerphobia (Ott and Aoki, 2002; Zahzah, 2019). It may be prudent to evaluate if framing differences condition prosocial responses and their temporal durability.

Fourth, another limitation is that we only focus on *indirect* exposure to *high-profile* violence. *Direct observation* of smaller-scale quotidian violence against LGBTQ+ group
members (e.g. observing hate crimes, assault, verbal abuse) may have a stronger, durable influence on prosocial beliefs. Future research should explore how different violence exposure types motivate prosocial beliefs.

Fifth, another limitation is that we only explore attitudinal shifts, not behavior. See Section A.11 for reasoning and evidence the lack of emphasis on behavior may not be a shortcoming.

8 Conclusion

We present a *Fickle Prosocial Violence Response* model to explain how indirect exposure to civilian violence against marginalized groups may influence prosocial attitudes toward targeted groups. Across four studies and three events, we provide evidence supporting the model and show indirect civilian violence against LGBTQ+ group members increases prosocial attitudes toward LGBTQ+ community segments. However, these prosocial responses are not temporally sustainable and less salient events do not motivate prosociality at the outset. Our core contribution is that we repeatedly demonstrate indirect exposure to salient civilian violence against marginalized groups may not sustainably undercut negative attitudes toward these groups. The *FPVR* model provides a general framework that can be tested and theoretically built upon in other domains outside anti-LGBTQ+ violence, such as violence against other marginalized groups (e.g. non-whites, immigrants, women).

Interestingly, we find limited individual-level heterogeneity in Studies 1-2 and 4, and some evidence non-whites and Democrats are more likely to adopt prosocial attitudes after Shepard's murder in Study 3.⁴⁸ The absence of heterogeneous effects in Studies 1-2 are not necessarily theoretically surprising. The *parallel publics* thesis posits salient events can generate common information exposure and therefore homogeneous attitudinal responses across population subgroups (Page and Shapiro, 2010). Relatedly, there was mainstream agreement among media and elites the Pulse massacre was tragic and reflected illegitimate

 $^{^{48}}$ We also find limited evidence of heterogenous decay in effects for Studies 1-3 (Section H).

behavior. Thus, messaging associated with the massacre was not a "group cue" that could motivate prosocial responses among some subgroups but not others (Zaller, 1992). Indeed, the effect homogeneity we identify is consistent with prior evidence showing *SSM support* moves in parallel over time across partisan and social subgroups (Coppock, 2023). Study 3's individual-level heterogeneity may be a function of temporal context. Relative to 2016, racial violence was salient in 1998. James Byrd was murdered 4 months before Shepard's murder. The Rampart LAPD scandal was also underway (involving the police beating of Ishmael Jimenez). Therefore, non-whites may have been primed to adopt prosocial attitudes toward groups facing conceivably analogous violence. Likewise, the mass public was less acceptant toward LGBTQ+ people in the 1990s. Therefore, socially conservative Republicans and independents may have been resistant to sympathetic messaging after Shepard's murder relative to liberal Democrats. Finally, Study 4's limited heterogeneity may be due to Club Q's limited salience vis-a-vis Pulse and Shepard's murder.

What would generate temporally sustainable effects? The *FPVR* model suggests sustained media attention may motivate sustained attitudinal shifts (Figure 2). Disturbingly, salient violent event *recurrence* may facilitate sustainable prosocial shifts. Additionally, the *FPVR* model posits elites play a role in making violent incidents salient. Therefore, elites who continue to strategically amplify issues related to a specific event long after occurrence may sustain attitudinal shifts (Zaller, 1992; Birkland, 1998). Moreover, perhaps *direct* or *proximal*, as opposed to *indirect*, violence exposure is necessary to durably shift mass attitudes. Prior work shows direct violence exposure produces temporally sustainable attitudinal/behavioral shifts (Lupu and Peisakhin, 2017; Mironova and Whitt, 2018; Hadzic et al., 2020). Prosocial consequences may also be temporally durable conditional on victim or perpetrator characteristics, and the scale of violence. We leave it to future research to continue to develop new theoretical insights and assess possibilities for durable effects.

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A Pulse Context

A.1 Demonstrating Pulse Was Salient



Figure A1: Survey Data Demonstrate the Pulse Massacre Was Salient. Panels A and B display how closely respondents were following the Pulse shooting in a June 2016 CBS and Kaiser poll respectively. Panel B compares attention to Pulse (x-axis) relative to other issues (y-axis). All estimates are population weighted. 95% CIs displayed from 1000 bootstrap simulations. See Section A.6 for more details on Figure A1 polls.

A.2 Media Coverage Over Time



Figure A2: Media Coverage of Topics Related to the Pulse Massacre Over Time. Panels A, C, and E display the count of Pulse-, LGBTQ-, and terrorism-related stories between January-October 2016. Panels B, D, and F display the ratio of Pulse-, LGBTQ-, and terrorism-related stories relative to the total number of stories in digital news. Loess models fit on each side of the moment the massacre occurs. Annotations denote RDiT estimates for the effect of Pulse on the article count and ratio using MSE optimal bandwidth selection (Calonico et al., 2015) (running variable degree = 1). See Appendix A.7 for more details on Figure A2 data.

A.3 Search Behavior Over Time



Figure A3: Search Behavior From Google Trends Demonstrates the Pulse Massacre Was Salient and Unexpected. Panels A, B, and C display the average search intensity for Pulse massacre-, LGBTQ-, and terrorism-related terms between January-October 2016. Vertical lines and annotations denote key events related to respective topics. See Appendix A.9 for more details on Figure A3 data.



A.4 Demonstrating Public Perceived Pulse

Figure A4: The Pulse Massacre Was Perceived by the Mass Public. Panels A-E characterize predicted values of belief country is less safe since 9/11, terror attacks are likely in the future, international terrorism is a critical threat, worry about terrorism, and worry about gun violence respectively. Panel F characterizes the the influence of Pulse (x-axis) on the aforementioned outcomes (y-axis) adjusting and not for imbalanced covariates (black = with controls, grey otherwise, see Figure A6 for balance plot). All covariates rescaled between 0-1. 95% CIs displayed derived from HC2 robust standard errors. Data are from the Chicago Council on Global Affairs Survey (June 10-26). See Section A.8 for more details on Chicago Council data. See also Table A1.

A.5 Demonstrating Public Perceived Massacre as Hate Crime



Figure A5: The Pulse Massacre Was Perceived as Targeted Anti-LGBTQ+ Violence. Panels A and B display beliefs the public felt the shooting was an anti-LGBTQ+ hate crime in a June 2016 CBS poll (Panel A) and July 2016 AP poll (Panel B). All estimates are population weighted.

A.6 Salience Data Details

CBS News June 13-14 Poll is a nationally representative adult survey (N = 1001). The poll used a random digit dial methodology. Interviews were conducted in English and Spanish using live interviewers. The data are weighted to reflect U.S. census figures on demographic variables. The margin of error for the weighted data is ± 4 percentage points. The item of interest on Figure A1, Panel A is: "How closely have you been following news about the recent shooting at a nightclub frequented by gays and lesbians in Orlando, Florida where at least 49 people were killed – very closely, somewhat closely, not too closely, or not at all closely?" The item of interest on Figure A5, Panel A is "Do you think the shooting at people who are gay and lesbian), or both?"

Kaiser Family Foundation June 15-21 Poll is a nationally representative adult survey (N = 1201). The poll used a random digit dial methodology. The item of interest on Figure A1, Panel A is: "How closely have you been following news about the recent shooting at a nightclub frequented by gays and lesbians in Orlando, Florida where at least 49 people were killed – very closely, somewhat closely, not too closely, or not at all closely?"

A.7 Media Attention Data Details

We acquired media data on the daily number of web articles related to the topics of interest from Mediacloud's Explorer Search Tool (https://explorer.mediacloud.org/) from January 1, 2016 to October 15, 2016 to generate Figure A2. The reason we do not include data after October 15, 2016 in our analyses is because we do not want our analyses to be perturbed by the 2016 election, which increased attention to LGBTQ-related topics due to Trump's anti-LGBTQ positions. The two measures of media attention we evaluate are the *article count* and *article ratio*. The article count is the raw number of web articles including a specific search term(s). The article ratio is the number of web articles including a specific search term(s) normalized over the total number of web articles.

We acquire article count and ratio data on three topics.

- 1. Pulse-related topics
- 2. LGBTQ-related topics
- 3. Terrorism-related topics.

Pulse-related topics are the article count sum and article ratio mean for queries on the terms "orlando massacre," "orlando shooting," "pulse nightclub," "pulse shooting." LGBTQ-related topics are the article count sum and article ratio mean for queries on the terms "anti-gay," "anti-lgbt," "gay marriage," "gay rights," "hate crime," and "same sex marriage." Terrorism-related topics are the article count sum and ratio mean for queries on the terms "isis," "lone wolf," "mass shooting," "terror attack," and "terrorism."

A.8 Chicago Council Study

A.8.1 Data Details

The Chicago Council on Global Affairs Poll is a nationally representative adult survey fielded between June 10-26, 2016 (N = 2061). The survey was conducted by GfK Knowledge Networks. The margin of sampling error for the weighted data is ± 2.4 percentage points. The data are subsetted to respondents who took between 10-60 minutes to complete the roughly 120 item survey (N = 1704).

A.8.2 Outcome Items

"Less Safe Since 9/11" Do you think that, as a country, we are more safe, about as safe, or less safe than we were before the terrorist attacks of September 11th, 2001? 1) More safe, 2) About as safe 3) Less safe. Measured binary = 1 if respondent indicates "less safe."

"Terror Attacks Likely" How likely is it that occasional acts of terrorism in the U.S. will be part of life in the future? 1) Very likely, 2) Somewhat likely, 3) Not very likely, 4) Not at all likely. Re-scaled from 0-1 with 1 = very likely.

"Terrorism = Critical Threat" Below is a list of possible threats to the vital interest of the United States in the next 10 years. For each one, please select whether you see this as a critical threat, an important but not critical threat, or not an important threat at all: International terrorism. 1) Critical threat, 2) Important but not critical threat, 3) Not an important threat. Measured binary = 1 if respondent indicates "critical threat."

"Worried (Terrorism)" Are you very worried, somewhat worried, not very worried or not worried at all that: You or someone you know will be the target of a terrorist attack. 1) Very worried, 2) Somewhat worried, 3) Not very worried, 4) Not at all worried. Re-scaled from 0-1 with 1 = very worried.

"Worried (Gun Violence)" Are you very worried, somewhat worried, not very worried or not worried at all that: You or someone you know will be the target of gun violence. 1) Very worried, 2) Somewhat worried, 3) Not very worried, 4) Not at all worried. Re-scaled from 0-1 with 1 = very worried.

A.8.3 Balance Plot



Figure A6: Covariate Balance for Survey Respondent Characteristics Before and After the Pulse Massacre in the Chicago Council on Global Affairs Survey (June 2016). Black coefficients are statistically significant, grey otherwise. All estimates use post-stratification survey weights to ensure representativeness. All covariates scaled between 0-1. 95% CIs displayed derived from HC2 robust standard errors.

A.8.4 Regression Table

| (1) | (2) | (0) | | | | Worry (Terror) | | | |
|---------------------|--------------|---------|-------------|--------|---------------|----------------|---------------|------------|--------------|
| (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) |
| Post-Pulse 0.10** | * 0.12*** | 0.07*** | 0.08*** | 0.03 | 0.05^{*} | 0.06*** | 0.05** | 0.03^{*} | 0.02 |
| (0.03 | (0.03) | (0.02) | (0.02) | (0.03) | (0.02) | (0.02) | (0.02) | (0.02) | (0.02) |
| Age | 0.19^{*} | | 0.09 | | 0.24** | | -0.06 | | -0.07 |
| | (0.08) | | (0.05) | | (0.08) | | (0.05) | | (0.05) |
| White | 0.10^{**} | | 0.05^{*} | | 0.01 | | -0.05^{**} | | -0.06^{**} |
| | (0.03) | | (0.02) | | (0.03) | | (0.02) | | (0.02) |
| Woman | 0.03 | | 0.02 | | 0.07^{**} | | 0.09^{***} | | 0.09*** |
| | (0.03) | | (0.02) | | (0.02) | | (0.02) | | (0.02) |
| Child | -0.01 | | -0.03 | | -0.09^{**} | | -0.01 | | 0.00 |
| | (0.03) | | (0.02) | | (0.03) | | (0.02) | | (0.02) |
| Non-Religious | -0.07 | | -0.02 | | -0.14^{***} | | -0.08^{***} | | -0.03 |
| | (0.04) | | (0.02) | | (0.04) | | (0.02) | | (0.02) |
| Income | -0.09 | | 0.13^{**} | | 0.23^{***} | | -0.02 | | -0.03 |
| | (0.06) | | (0.05) | | (0.06) | | (0.04) | | (0.04) |
| College | -0.05 | | -0.03 | | -0.03 | | -0.02 | | 0.00 |
| | (0.04) | | (0.02) | | (0.03) | | (0.02) | | (0.02) |
| Unemployed | -0.01 | | 0.01 | | -0.04 | | 0.01 | | -0.00 |
| | (0.03) | | (0.02) | | (0.03) | | (0.02) | | (0.02) |
| Rent | -0.02 | | 0.01 | | 0.09^{**} | | 0.01 | | 0.03 |
| | (0.03) | | (0.02) | | (0.03) | | (0.02) | | (0.02) |
| Liberal | -0.08^{*} | | -0.02 | | -0.02 | | -0.04 | | 0.03 |
| | (0.03) | | (0.02) | | (0.03) | | (0.02) | | (0.02) |
| Conservative | 0.16^{***} | | 0.04 | | 0.05 | | 0.02 | | -0.06^{**} |
| | (0.03) | | (0.02) | | (0.03) | | (0.02) | | (0.02) |
| Urban | -0.02 | | -0.02 | | -0.00 | | 0.01 | | 0.02 |
| | (0.03) | | (0.02) | | (0.03) | | (0.02) | | (0.02) |
| State FE N | Υ | Ν | Y | Ν | Υ | Ν | Y | Ν | Υ |
| R ² 0.01 | 0.09 | 0.02 | 0.10 | 0.00 | 0.08 | 0.01 | 0.07 | 0.00 | 0.08 |
| N 1704 | 1704 | 836 | 836 | 1415 | 1415 | 1693 | 1693 | 1696 | 1696 |

Table A1: The Pulse Massacre Was Perceived by the Mass Public

Note: ***p < 0.001, **p < 0.01, *p < 0.05. Models 1, 3, 5, 7, 9 do not adjust for control covariates while Models 2, 4, 6, 8, and 10 do. All models use weights for representativeness. HC2 robust standard errors in parentheses.

A.9 Google Trends Data Details

We acquired Google Trends search data at the weekly level from the gtrendsR R package. We generate three different search intensity measures capturing interest in the Pulse nightclub shooting, LGBTQ-related issues, and terrorism-related issues. The Pulse-related issue measure is the average of the Google Trends search intensity measures for separate queries on the "pulse nightclub," "pulse shooting," "orlando massacre," and "orlando shooting." The LGBTQ-related issue measure is the average of the Google Trends search intensity measures for separate queries on "gay rights," "gay marriage," "same-sex marriage," "hate crime," "anti-gay," and "anti-lgbt." The terrorism-related issue measure is the average of the Google Trends search intensity measures for separate queries on "terrorism," "terror attack," "lone wolf," "ISIS," and "mass shooting."

The search intensity measure is the number of total searches divided by the total searches of the geography (United States) and time range (January 1, 2016-October 1, 2016) it represents to compare relative popularity. The numbers are scaled on a range of 0-100 based on a topic's proportion to all searches on all topics. For more information see https://support.google.com/trends/answer/4365533?hl=en

A.10 Hate Crime Perceptions Data Details

The AP/Black Youth Project July Poll is a nationally representative adult survey (N = 1940) fielded between July 9, 2016 and July 12, 2016 The data are weighted to reflect U.S. census figures on demographic variables. The margin of error for the weighted data is \pm 4 percentage points. The item of interest on Figure A5, Panel B is: "You may recall that last month (June 2016), 49 people were shot and killed (and 53 people were injured) by 29-year-old Omar Mateen at Pulse nightclub in Orlando, Florida. From what you remember, do you think the shooting at the nightclub in Orlando, Florida was a terrorist act, a hate crime against people who are gay, lesbian, bisexual, and transgender, a hate crime against Latinos/Hispanics, or none of the above? Please select all that apply."

A.11 Behavioral Shifts Post-Pulse

One limitation of our main analyses is that we only explore attitudinal shifts, not behavior. Behavioral shifts may not be commensurate with attitudinal changes. However, we do not believe our lack of emphasis on behavior is a shortcoming. The fickle attitudinal shifts we identify are consistent with the FPVR model's proposition the mass public may engage in short-term impression management until perceptibly anti-LGBTQ+ violence is no longer salient, making them more comfortable to express their original beliefs, and implying behavioral shifts are unlikely. Nevertheless, we explore if the Pulse massacre motivated pro/anti-LGBTQ+ behaviors and find mixed evidence.

We assess if the Pulse massacre motivated three different pro/anti-LGBTQ+ behaviors: anti-LGBTQ+ hate crimes, donations to Florida-based pro-LGBTQ+ organizations, and blood donations. We find the massacre motivated an increase in anti-LGBTQ+ hate crimes, consistent with prior research suggesting violence has a contagious effect (Section A.11.1) (Towers et al., 2015); no increase in monetary donations to Orlando LGBTQ+-serving organizations (Section A.11.7); and an increase in blood donations for victims (Section A.11.8). These findings suggest the massacre motivated both pro- and anti-social *behavior*, but given we use aggregate data, we cannot determine if this is due to behavioral *changes* or *priming* of those who are predisposed to either be anti-social or pro-social toward LGBTQ+ people. Future research should continue to explore the behavioral consequences of exposure to salient civilian violence against marginalized groups in addition to attitudinal consequences.

A.11.1 Anti-LGBTQ+ Hate Crimes: Details

We evaluate if the Pulse massacre motivated anti-LGBTQ+ hate crimes, consistent with prior research suggesting mass violence may have a contagion or "copy-cat" effect (Towers et al., 2015). To assess trends in hate crimes, we use data from the FBI Uniform Crime Report on hate crimes across the United States at the daily level between January 1, 2016, and December 31, 2016. Importantly, because the Pulse massacre was understood as a terrorist attack not necessarily motivated by anti-LGBTQ+ bias, it was not classified as a hate crime, even though it was perceived by the mass public as an anti-LGBTQ+ hate crime (Figure A5). Therefore, our analyses assessing the effect of the Pulse massacre on hate crimes is not driven by the massacre itself.

Figure A7 displays anti-LGBTQ+ (Panel A), anti-Black (Panel B), anti-Jewish (Panel C), and anti-Latino (Panel D) hate crimes during 2016 at the daily-level over time. The descriptive statistics suggest anti-LGBTQ+ hate crimes increased for a brief period after the Pulse massacre, but not anti-Black, anti-Jewish, and anti-Latino hate crimes.

Regression discontinuity-in-time estimates using the Calonico et al. (2015) optimal bandwidth selection approach corroborates the descriptive statistics (Figure A8). Immediately after Pulse, there's an increase in roughly 2 daily anti-LGBTQ+ hate crimes. However, there is 0 increase in the number of daily anti-Black, anti-Jewish, or anti-Latino hate crimes. These findings are robust to a variety of kernel and polynomial specifications for the running variable (days to Pulse).

The regression discontinuity estimates characterizing are robust. They hold using a variety of bandwidths from 10-100 days (Figure A9), and many of the coefficients are larger than at least 90% of the effects from pre-treatment placebo discontinuities (Figure A10). These effects are also not driven by Pride month, since they do not manifest in years prior to Pulse (2010-2015) or years after Pulse (2017-2019) (Figure A11).





Figure A7: Descriptive Statistics Characterizing Different Hate Crimes Over Time in 2016. The x-axis is the date. The y-axis is the number of hate crimes in a given day. Dashed vertical line denotes the moment the Pulse massacre occurred (June 12). Loess lines fit on each side of the moment the Pulse massacre occurred. Panels A-D display anti-LGBTQ+, anti-Black, anti-Jewish, and anti-Latino hate crimes.



A.11.3 Anti-LGBTQ+ Hate Crimes: RDiT Estimates

Figure A8: Regression Discontinuity-in-Time *Post-Pulse* Coefficient Estimates and Hate Crimes The x-axis is the hate crime type. The y-axis is the *Post-Pulse* coefficient. Color denotes kernel and polynomial degree at use. 95% CIs displayed from robust standard errors.



A.11.4 Anti-LGBTQ+ Hate Crimes: Close to Bandwidth Estimates

Figure A9: Regression Discontinuity-in-Time *Post-Pulse* Coefficient Estimates Using Bandwidths Close to Discontinuity The x-axis is the bandwidth (in days). The y-axis is the *Post-Pulse* coefficient. Each panel denotes the kernel at use and running variable polynomial degree (0-2). 95% CIs displayed from robust standard errors.



A.11.5 Anti-LGBTQ+ Hate Crimes: Pre-Pulse Temporal Placebo

Figure A10: Comparing *Post-Pulse* Coefficient with Temporal Placebo Tests Prior to Pulse The x-axis is the temporal placebo coefficient size. Vertical line denotes true *post-Pulse* coefficient size. Annotation denotes the proportion of placebo coefficients the true coefficient is larger than. Panels denote kernel and polynomial degrees (0-2). 95% CIs displayed from robust standard errors.





Figure A11: Post-June 12 Placebo Tests on Years Prior to and After Pulse. The x-axis denotes the hate crime dataset year (2010-2015, 2017-2019). The y-axis characterizes the RDiT coefficient of a placebo indicator equal to 1 after June 12, the calendar date of the Pulse massacre. Color denotes kernel and polynomial degree at use (0-2). 95% CIs displayed from robust standard errors.



A.11.7 Donations to Pro-LGBTQ+ Organizations

Figure A12: The Pulse Massacre Did Not Motivate A Differential Increase in Donations to Pro-LGBTQ+ Organizations in Florida. The y-axis is the differential effect of Pulse on the logged donations to Florida pro-LGBTQ+ organizations that are a part of the One Orlando Alliance. The x-axis is the time to treatment (tax years 2011-2020). Dashed line denotes *post-Pulse* coefficients. 95% CIs displayed.

To assess if the Pulse massacre motivated support for pro-LGBTQ+ organizations serving the Orlando LGBTQ+ community, we assess if contributions (i.e. donations) to non-profit pro-LGBTQ+ organizations serving Orlando differentially increased relative to other nonprofit organizations after the Pulse massacre. We used two different datasets to conduct this assessment. First, we used tax return information on the universe of non-profits that submitted tax returns between 2011-2020 from the Internal Revenue Service (IRS).⁴⁹ This data includes our outcome of interest, the amount of monetary contributions declared in a given tax year (inflation adjusted to 2011 U.S. dollars). We log the contributions outcome (plus 1 to ensure identification, log(contributions + 1)). Second, we merged this information with data we collected identifying non-profits who were serving the Orlando LGBTQ+ community and were soliciting monetary support through the One Orlando Alliance, a conglomerate of LGBTQ+ serving organizations in Central Florida that engaged in resource sharing after the Pulse massacre.⁵⁰ Consistent with the sample we derived from the IRS data, we only included One Orlando organizations who filed tax returns for each year between 2011-2020 (suggesting they existed across the entire temporal domain of the panel) and were local, not national organizations (e.g. the Human Rights Campaign, ACLU). We exclude national or-

⁴⁹Source: https://www.irs.gov/charities-non-profits/form-990-series-downloads

⁵⁰Source: https://oneorlandoalliance.org/our-history/

ganizations identified on the One Orlando Alliance member list from the IRS data as well.⁵¹ One Orlando Alliance non-profit organizations are coded 1 if they are a part of the Alliance and 0 otherwise in the IRS data (*alliance*).

Figure A12 displays event study estimates from a synthetic controls approach developed by Xu (2017) characterizing the differential effect of Pulse on One Orlando Alliance organization donation receipts. We use the synthetic controls approach to reweight pre-treatment outcome data from the set of untreated non-profit organizations to generate a counterfactual that satisfies the parallel trends assumption to derive the plausibly causal effect of Pulse on contributions to One Orlando Alliance organizations. The event study demonstrates the effect of Pulse on differential donations to Orlando LGBTQ+-serving organizations is 0, suggesting Pulse did not motivate an increase in donations or contributions to key Orlando LGBTQ+-serving organizations.

⁵¹See https://oneorlandoalliance.org/our-members/ for the complete list of One Orlando Alliance affiliated organizations, the organizations included in the sample are: 1) Community Legal Services of Mid Florida, 2) Equality Florida, 3) Family Equality, 4) Hope & Help, 5) Hope Community Center, 6) Legal Aid Society of the Orange County Bar Association, 7) Mental Health Organization of Central Florida, 8) Miracle of Love, 9) Orlando Gay Chorus, 10) Planned Parenthood of Southwest and Central Florida 11) Victim Service Center of Central Florida.



Figure A13: The Pulse Massacre Increased Search Interest in Donating Blood.

Given we do not possess direct data on blood donations for victims of the Pulse massacre, we use Google Trends data to identify the intensity of search interest in "blood donation" over time for the year 2016. Figure A13 clearly demonstrates search interest in "blood donation" substantially increases during the week of the Pulse massacre, but quickly drops off in the following weeks. Although search interest in "blood donation" may not necessarily translate into real-world behavioral action to donate blood, we are confident that our Google Trends analysis provides a rough proxy of real-world blood donation behavior due to qualitative accounts of blood donation after the Pulse massacre. According to the Orlando Sentinel,⁵² Orlando hospitals who took in Pulse massacre victims "never had a shortage of blood and no victim experienced a delay in getting the right type of blood." This is because "Thousands of people began donating blood, throughout Florida and even in other states, starting hours after the June 12 shooting. The donations far exceeded the blood needed for the shooting." Moreover, "In the week after the attack, OneBlood took in 28,000 pints of blood; the agency's average weekly volume is about 18,000 pints…It was the biggest response since the Sept. 11 terror attacks in 2001."

The search intensity measure is the number of total searches concerning blood donations divided by the total searches of the geography (United States) and time range (January 1, 2016-December 31, 2016) it represents to compare relative popularity. The numbers are scaled on a range of 0-100 based on a topic's proportion to all searches on all topics. For more information see https://support.google.com/trends/answer/4365533?hl=en

⁵²Source: https://www.orlandosentinel.com/news/pulse-orlando-nightclub-shooting/os-oneblood-ceo-pulse-20160629-story.html

A.12 Ruling Out Bundled Treatment Concerns With Placebo and Falsification Tests

One potential shortcoming of the two studies (Studies 1, 2) assessing the effects of *post-Pulse* on attitudes toward LGBTQ+ community segments is that the treatment is "bundled" in the sense that the Pulse massacre was perceived by the mass public as not only an anti-LGBTQ+ hate crime, but a terrorist attack in addition to an anti-Latino hate crime. Indeed, a plurality of the mass public perceived the shooting as *both* a terrorist attack and LGBTQ+ hate crime (Figure A5, Panel A). In addition, nearly 20% of the mass public perceived the shooting as a Latino hate crime (Figure A5, Panel B). Therefore, inconsistent with our theory, our findings may be driven by the fact the Pulse massacre was either a terrorist attack or anti-Latino hate crime.

We rule these possibilities out in several ways. First, we assess the effects of two prominent "Islamic" terrorist attacks on attitudes toward LGBTQ+ community segments: the April 2013 Boston bombing and the December 2015 San Bernardino shooting (Gunaratna and Haynal, 2013; Fitzpatrick, 2018). These incidents were highly salient to the mass public. Roughly 80% of the public reported they were following the Boston Bombing closely immediately after the bombing (higher than other salient issue at the moment of the bombing: the Gun Control Debate, Immigration Policy Debate, Texas Fertilizer Explosion, Poison Letters to Obama, Syrian Chemical Weapons, and Flight Delays, see Figure A15). The public also rated the San Bernardino attack the second most important issue or two of 2015, beating the Gay Marriage Decision, the Republican primary, and the Iran Deal (Figure A17). If pro-LGBTQ+ beliefs manifest after these terrorist attacks, then our findings may not be driven primarily by the perception of anti-LGBTQ+ violence, but rather the perception of a terrorist attack motivated by radical beliefs associated with Islam.

Figures A16 and A18 display coefficients characterizing the influence of the Boston bombing and the San Bernardino shooting on the *D*-score, straight bias, and heterocentrism outcomes using Project Implicit Sexuality IAT data from 2013 and 2015 respectively (5-50 day bandwidths from the moment of the event of interest).⁵³ With the exception of late-term effects for the *D*-score outcome in the 2015 data assessing the influence of the San Bernardino shooting, these events have had a null influence on the various outcomes of interest. Although the *D*-score appears to decrease after the San Bernardino shooting, heterocentrism does not decrease as well in a manner similar to the post-Pulse effects. Moreover, the Dscore decreases 30 days from the San Bernardino shooting, as opposed to just 15 days from the Pulse shooting. Therefore, the results characterizing the effect of the San Bernardino shooting are more likely to be a function of unobserved secular temporal trends unrelated to the shooting relative to the results characterizing the effect of Pulse. Indeed, the samples at use 30 days from the San Bernardino shooting are imbalanced on several covariates, including ideology (Figure A19). Moreover, we provide additional evidence that terror attacks associated with Islam do not systematically motivate pro-gay attitudes. We assess the effect of several Islamic extremist terror attacks between 2009-2020 on the D-score, heterocentrism, and straight bias outcomes. We do not find consistent, systematic evidence that these attacks

⁵³All outcomes for the Boston and San Bernardino attack analyses are measured similarly as those in the main text for Study 2. The Boston and San Bernardino attack analyses also adjust for the same covariates outlined in Study 2.

motivated pro-gay attitudes (Figure A20). In the aggregate, these findings imply the fact the Pulse massacre was a terrorist attack inspired by ISIS is not the main channel driving our results.

Second, we demonstrate that the effects are not motivated by negative attitudes toward Muslims/Islam in response to terror attacks associated with extremist Islamic organizations. For instance, the American mass public may seek to distinguish themselves from an Islamic/Muslim ideology that is perceptibly socially conservative on the dimension of sexuality and/or queerness in response to terror attacks (i.e. "pinkwashing," see Meyer (2020)). Consequently, the mass public may adopt prosocial beliefs toward segments of the LGBTQ+ community after the Pulse massacre as a function of concomitant animosity toward Muslims. If this mechanism explains our findings, then we may expect the mass public to adopt negative attitudes and/or behaviors toward Muslims and/or their political rights after the Pulse massacre.

We provide three pieces of evidence this mechanism may not be operative. Relative to respondents interviewed before Pulse, respondents interviewed after Pulse in the Project Implicit Arab IAT data do not adopt negative attitudes toward Arabs (Section C.10), an ethnic group strongly associated with Islam (d'Urso, 2022). In addition, we use two nationally representative ABC News telephone surveys fielded shortly before and after the Pulse massacre to demonstrate members of the mass public interviewed after the Pulse massacre are not more likely to support banning Muslims from entering the United States (i.e. the "Muslim Ban", see Figure A23). Finally, we use day-level hate crime data from the FBI Uniform Crime Report between January 1, 2016-December 31, 2016 to assess if the Pulse massacre motivated anti-Muslim or anti-Arab hate crimes. Behaviorally, the mass public may engage in anti-Muslim organizations (Welch, 2006). We assess the discontinuous effect of the Pulse massacre on the daily number of anti-Muslim/anti-Arab hate crimes, and find the Pulse massacre did not result in an increase in anti-Muslim/anti-Arab hate crimes (Figure A24).

These empirical findings suggest our results are not driven by a heightened animosity toward Muslims in response to terror attacks associated with Islam among the mass public. Instead, these findings provide further support for our claim that the mass public perceived the Pulse massacre as an instance of anti-LGBTQ+ violence, and adopted prosocial attitudes toward segments of the LGBTQ+ community accordingly, at least briefly.

Third, we assess the effect of a prominent anti-Latino hate crime on attitudes toward LGBTQ+ community segments: the August 2019 El Paso Shooting (Leander et al., 2020). According to Google Trends, this incident was the most prominent hate crime of 2019, with the exception of the Jussie Smollett debacle in January/February 2019 (Figure A21). Again, if pro-LGBTQ+ beliefs manifest after the 2019 El Paso shooting, then our findings may not be driven via the channel of anti-LGBTQ+ violence, but anti-Latino violence. To assess the influence of the El Paso Shooting on pro-LGBTQ+ beliefs, we use UCLA Nationscape data from the Democracy Fund Voter Study Group (5-50 day bandwidths from the El Paso shooting),⁵⁴ a large non-probability survey fielded each week by Lucid between July 2019-February 2021 weighted to national government population estimates.

 $^{^{54}}$ See https://www.voterstudygroup.org/data/nationscape for details



Figure A14: Google Search Intensity Across Different Search Terms. The x-axis is week. The y-axis is relative Google Search intensity between search terms for "LGBT," "terrorism" and "latino". Dashed vertical line denotes the moment of the Pulse massacre.

The outcomes of interest in the Nationscape data are LGBT unfavorability (1 = "very unfavorable" or "somewhat unfavorable," <math>0 = "somewhat favorable," "very favorable," or "haven't heard enough") and no trans military <math>(1 = "disagree" to allowing transgender people to serve in the military," <math>0 = "agree" or "not sure"). Therefore, negative coefficients characterizing being interviewed after the 2019 El Paso shooting suggest the mass public is adopting relatively positive attitudes toward LGBT people and the notion trans people may serve in the military.

For bandwidths between 5-50 days before and after the El Paso shooting, we do not find respondents interviewed after the shooting hold more favorable attitudes toward LGBT and the notion trans people may serve in the military (Figure A22). These findings imply the fact the Pulse massacre was an instance of violence against predominantly Latinx people is not the main channel driving our results. Instead, these findings suggest our main results assessing the effect of the Pulse massacre are driven by the perception the event was anti-LGBTQ+ violence.

Moreover, if the massacre motivated prosocial beliefs toward LGBTQ+ community segments because it was also an instance of violence against Latinxs, we may expect the massacre to motivate prosocial beliefs toward Latinxs. However, using an additional unexpected-eventduring-survey design with the 2016 General Social Survey, we do not find that the massacre motivated reductions in old-fashioned ethno-racism toward Hispanics (Figure A26), a wellestablished measure of ethno-racism (Tesler, 2013). We also do not find the massacre increased support for a pathway to citizenship for undocumented immigrants in the TAPS survey (Figure 3, Panel C). A pathway to citizenship disparately benefits Latinxs given twothirds of Latinxs are either immigrants or children of immigrants. These findings further suggest our results are driven by the fact the massacre was perceived as anti-LGBTQ+ violence, not anti-Latinx violence.

Additionally, we provide evidence that the mass public was particularly attentive to LGBT topics relative to issues related to terrorism or Latinos. Google Trends data shows
that the relative search intensity for "lgbt" was much higher and pronounced the moment of the Pulse massacre than "terrorism" or "latino (Figure A14)." These findings further imply the Pulse massacre primarily motivated prosocial attitudes toward LGBTQ+ community segments through the perception of violence against LGBTQ+ people.

Lastly, our second Event in the main text suggests bundled treatment considerations are moot (i.e. Matthew Shepard's murder). Shepard's murder was unequivocally understood as an anti-gay hate crime by the mass public and political elites. Unlike Pulse, it was not simultaneously a terrorist attack or an instance of violence against Latinx people. Contemporary hate crime laws in the United States are even named after Matthew Shepard. Shepard's murder was not a terror attack nor an attack against a member of a politically non-dominant ethno-racial group. However, we find a similar pattern of results to the Pulse massacre, where positive attitudes toward gay people increase immediately after his murder, but dissipate in the long-run.





Figure A15: Salience of Boston Bombing. The x-axis the the proportion of respondents following each issue closely, the y-axis is the respective issue. Panel A is data from the April 18 Pew Survey. Panel B is data from the April 25 Pew Survey. 95% CIs displayed from 1000 bootstrap simulations. All estimates use survey population weights.



Figure A16: Influence of Boston Bombing on Anti-Gay Attitudes. The x-axis is the bandwidth (in days) used from the PI S-IAT data. The y-axis is the post-event coefficient. Annotations denote sample size (in thousands) corresponding to each respective coefficient estimate along the bandwidth size. All covariates rescaled between 0-1. 95% CIs displayed from HC2 robust standard errors.

A.12.2 Terror Attack: San Bernardino (2015)



NBC/WSJ December 2015 Poll

Figure A17: Salience of San Bernardino Shooting. The x-axis the proportion of respondents indicating each issue was the most important (or two) of 2015, the y-axis is the respective issue. Data are from the NBC/Wall Street Journal December 2015 poll.



Figure A18: Influence of San Bernardino Shooting on Anti-Gay Attitudes. The x-axis is the bandwidth (in days) used from the PI S-IAT data. The y-axis is the post-event coefficient. Annotations denote sample size (in thousands) corresponding to each respective coefficient estimate along the bandwidth size. All covariates rescaled between 0-1. 95% CIs displayed from HC2 robust standard errors.



Figure A19: Balance on IAT Taker Composition Before and After the San Bernardino Shooting. Each panel characterizes covariate balance for different bandwidths (see plot title). The x-axis is the *post-shooting* coefficient derived from separate regression models regressing a baseline covariate (y-axis) on *post-shooting*.

A.12.3 Terror Attack: Other Events



Figure A20: Influence of Other Terror Attacks on Anti-Gay Attitudes. Each panel characterizes the effects of terror attacks on anti-gay attitudes for 15 and 20-day bandwidth samples. The x-axis is the post-attack coefficient, the y-axis is the event. Color denotes outcome at use. Terror attack data are sourced from the following crowdsourced list: https://en.wikipedia.org/wiki/Terrorism_in_the_United_States#Islamist_extremism. PI S-IAT datasets on self-selected U.S. adults are used from each year that each attack occurs within. 95% CIs displayed from robust SEs.

A.12.4 Anti-Latino Attack: El Paso (2019)



Figure A21: Salience of El Paso Shooting As A Hate Crime. The x-axis is the date (in weeks), the y-axis is the Google search hit intensity for "hate crime." From left to right, dashed vertical lines denote Jussie Smollett reporting a hate crime, his indictment for faking the hate crime, and the El Paso shooting.



Figure A22: Influence of El Paso shooting on Anti-LGBTQ+ Attitudes. The xaxis is the bandwidth (in days) used from the Nationscape data. The y-axis is the post-event coefficient. Annotations denote sample size (in thousands) corresponding to each respective coefficient estimate along the bandwidth size. Differences in sample sizes across outcomes are not due to non-random missingness, but rather the random omission of the *no trans military* outcome item in the Nationscape data for respondents in the weekly subsamples. All covariates rescaled between 0-1. 95% CIs displayed from HC2 robust standard errors. All estimates use survey poulation weights.



A.12.5 Muslim Ban Falsification

Figure A23: Influence of Pulse Massacre on Support for Muslim Ban. Panel A characterizes covariate balance between the ABC News May 2016 and ABC News June 2016 surveys. Panel B characterizes the influence of being interviewed in the June 2016 survey on support for the Muslim Ban with and without control covariates (i.e. the balance covariates). All covariates rescaled between 0-1. 95% CIs displayed from HC2 robust standard errors. All estimates use survey poulation weights.

Details: To assess the influence of Pulse on attitudes toward the Muslim Ban, we stacked two different ABC News Telephone Polls fielded shortly before and after the Pulse massacre. The first ABC survey was fielded between May 16-19, 2016 (N = 1005), less than one month before the massacre. The second ABC survey was fielded between June 20-23, 2016 (N =1001), just a week after the massacre. The outcome of interest is support for the Muslim Ban. The two surveys ask respondents if they "would support or oppose a temporary ban on Muslims who are not U.S. citizens from entering the United States?" The outcome is coded 1 if the respondent indicates "support, strongly" or "support, somewhat," 0 if the respondent indicates "oppose, somewhat" or "oppose, strongly." We assess the effect of being interviewed *post-Pulse* relative to pre-Pulse. If respondents are inclined to restrict the rights of Muslims *post-Pulse*, then the *post-Pulse* coefficient with respect to the Muslim Ban outcome would be positive. We also adjust for a number of covariates (age, woman, white, college education, income, liberal ideology, Democrat, and state-level indicators for Texas, California, New York, Florida, and Pennsylvania), that we also assess balance for, suggesting that respondents interviewed before and after Pulse in the ABC polls are compositionally similar (Figure A23, Panel A). We do not find evidence respondents interviewed after Pulse are more likely to support the Muslim Ban (Figure A23, Panel B). With or without covariate adjustment, the post-Pulse coefficients being either 0 or near-zero, and statistically insignificant.

A.12.6 Muslim Hate Crime Falsification



Figure A24: Discontinuous Effect of Pulse Massacre on Anti-Muslim/Anti-Arab Hate Crimes. Panel A characterizes daily anti-Muslim/Arab hate crimes over time in 2016. Solid lines are loss models fit to each side of the moment of the Pulse massacre occurs. The dashed vertical line characterizes the moment the Pulse massacre occurs. Panel B characterizes regression discontinuity-in-time (RDiT) coefficient estimates (y-axis) of the effect of the Pulse massacre on anti-Muslim/Arab hate crimes across kernel/polynomial specifications (x-axis). Bandwidth selection is data-driven, mean-squared optimal (see Calonico et al. (2015)). 95% CIs displayed derived from robust SEs.



A.12.7 Latino Old-Fashioned Ethno-Racism: Balance Tests

Figure A25: Balance Tests Between Respondents Interviewed Before and After the Pulse Massacre (GSS '16) The x-axis is the *post-Pulse* coefficient, the y-axis is the balance covariate. Each panel characterizes the bandwidth (5-50 days, then full sample) and sample size for each bandwidth sample. Black coefficients are statistically significant, grey otherwise. All estimates are population-weighted. Data is from the 2016 General Social Survey. 95% CIs displayed derived from robust SEs.





Figure A26: Effect of Pulse on Old Fashioned Ethno-Racism (GSS '16). The x-axis is the bandwidth sample at use (in days), the y-axis is the *post-Pulse* coefficient. The outcome for Panel A is a binary indicator if the respondent indicates they favor intermarriage with a Latino for a familial relative, the outcome for Panel B is a binary indicator if the respondent indicates they oppose intermarriage with a Latino for a familial relative. Black coefficients are from models adjusting for age, gender, race, college-education, income, partisanship, and ideology, grey otherwise. Data is from the 2016 General Social Survey. 95% CIs displayed derived from robust SEs.

B Study 1: TAPS

B.1 Outcome Measurement

To measure support for same-sex marriage, we use an item in the June 2016 TAPS survey asking respondents if "you generally support or oppose same-sex marriage." with options to choose: 1) Support; 2) Oppose; and 3) No opinion.

B.2 Manipulation Check



Figure B27: Belief ISIS = Most Important Issue Increases After Pulse. All estimates use post-stratification survey weights to ensure representativeness. All covariates scaled between 0-1. 95% CIs displayed derived from HC2 robust standard errors.

B.3 Insensitivity to Truncation

B.3.1 Discussion

Online survey respondent inattentiveness produces low quality responses that attenuate associations of interest (Read et al., 2021). Attention is critical for question comprehension and retrieval of relevant information from memory to form a judgement (Krosnick and Alwin, 1987). Our design depends upon respondents cognitively making connections between violence against marginalized groups they observe in mass media and their policy preferences implicating said groups. Prior research suggests very quick and very slow survey response times are associated with lower attention and quality responses (Malhotra, 2008; Read et al., 2021). In TAPS, the minimum response time was 3 minutes, insufficient to process a ~ 250 item survey. Furthermore, the maximum response time is 34,586 minutes, raising the possibility some respondents were multi-tasking, distracted, or intermittently engaging the survey with low effort. Thus, in the absence of internal attention checks, we truncate the sample to respondents who completed the survey in a "reasonable duration" of time, defined as those who took between 15-60 minutes to complete the survey. Our truncation is consistent with the rule of thumb by Roßmann (2010), who suggest removing respondents below 60% the median completion time.

The final TAPS data contain N = 1142 respondents, with 682 (60%) interviewed before Pulse and 460 after (40%). Truncation is unlikely to undercut generalizability. There are limited differences between inattentive and attentive TAPS respondents (Figure B28, Panel A). Additionally, the truncated sample is compositionally similar to the full TAPS sample and the "gold standard" in election studies, the 2016 ANES (Figure B28, Panel B, Table B2). Although our truncation is arbitrary, we follow best practices (Greszki et al., 2015), and show the results are insensitive to using the initial raw data or alternative response time cut-offs for "reasonable duration (Figure B28, Panel C)."⁵⁵

 $^{^{55}}$ Another benefit of the truncated data is the reduction in imbalance between respondents interviewed before and after the massacre. The truncated sample is imbalanced on 1/20 baseline covariates, whereas the full sample is imbalanced on 3/20 covariates.

B.3.2 Analyses

| Covariate | TAPS Jun. '16 | ANES '16 | Diff. | T-test p-value |
|---------------|---------------|-----------------|-------|----------------|
| Woman | 0.51 | 0.51 | 0.01 | 0.73 |
| White | 0.78 | 0.78 | 0.01 | 0.71 |
| Age (18-29) | 0.20 | 0.18 | 0.03 | 0.03 |
| Age $(30-44)$ | 0.24 | 0.23 | 0.00 | 0.76 |
| Age $(45-59)$ | 0.29 | 0.32 | -0.02 | 0.13 |
| Age $(60+)$ | 0.26 | 0.27 | -0.01 | 0.55 |
| College | 0.31 | 0.29 | 0.02 | 0.16 |
| Liberal | 0.39 | 0.41 | -0.02 | 0.32 |
| California | 0.10 | 0.09 | 0.01 | 0.49 |
| New York | 0.05 | 0.04 | 0.01 | 0.36 |
| Florida | 0.05 | 0.06 | -0.01 | 0.22 |
| Pennsylvania | 0.05 | 0.05 | -0.00 | 0.93 |
| Texas | 0.07 | 0.08 | -0.01 | 0.18 |

Table B2: Comparison Between truncated TAPS June '16 Sample and ANES'16 Sample

Figure B28 displays estimates using different types of data truncation. The x-axis displays the kinds of respondents that are removed. For instance, >15, <60 means that respondents who took more than 15 minutes and less than 60 minutes are included in the sample, and those who took less than 15 minutes and more than 60 minutes are excluded from the sample.

The truncated estimates operate in a manner consistent with the notion that respondents who take the survey either too quickly or too slowly are less attentive. Respondents who take the survey too quickly may not have sufficient time to make cognitive connections between their political context and their expressed attitudes on particular issues. Respondents who take the survey for too long may be intermittently attentive to the survey or are not taking the survey as seriously as they otherwise should, again, undercutting cognitive connections between their political context and their expressed attitudes on particular issues (Malhotra, 2008; Read et al., 2021). We find that removing respondents who take too long to take the survey increases the size of the coefficient estimates, consistent with research on how inattention attenuates coefficient estimates. However, we do not find that removing speeders increases the size of the coefficient estimates (for example, respondents who take less than 15 minutes to take the survey, the threshold we use for the results in the main text).

We do not believe this to be a problem, given most speeders are not engaging in egregious levels of speeding and true speeders are a very small proportion of the sample, which would suggest speeders have an inconsequential effect on coefficient size. In the TAPS data, of the speeding population (that is, those who take the survey in less than 15 minutes), over 80% take the survey in more than 10 minutes. This is fast for a large survey, but not egregiously fast. The other 20% (only 21 respondents), took the survey in less than 10 minutes. Therefore, the number of serious speeders may not be large enough to affect post-Pulse coefficient estimates. But, dropping slow respondents based on our cutoff leads to 352 dropped respondents, who may be particularly inattentive to the survey given that the median "slow respondent" took 1421 minutes (24 hours) to respond to the survey.

Regardless, the results are *insensitive to truncation*. Across the different truncated estimates adjusting for controls on Figure B28, 15/16 are statistically significant at p < 0.10,



Figure B28: Truncated Estimates. Panel A displays balance between removed respondents (who finished the survey in less than 15 minutes, more than 60) and respondents that were not removed. Panel B displays balance between the full and the truncated sample. Panel C displays coefficients characterizing the influence of *post-Pulse* on *SSM support* based on various samples removing respondents who took more than or less than a particular number of minutes (defined on the x-axis). For Panels A-B, black coefficients are statistically significant, grey otherwise. For Panel C, black coefficients are derived from regression models including baseline control covariates (i.e. the balance covariates), grey coefficients are derived from regression modeels that do not include control covariates.

and 8/16 are statistically significant at p < 0.05. Importantly, the results hold without truncating the data at all at p < 0.10 (Figure B28, Panel C). The findings on temporal persistence are also the same without truncating the data (Figure B29). In addition, the *post-Pulse* effect may not be biased given the TAPS survey, when weighted, is composition-ally similar to the 2016 ANES, the gold standard in representative surveys. Prior evidence suggests the maintenance of a representative sample composition mitigates the prospect for coefficient effect bias after truncating data to attentive respondents (Alvarez et al., 2019).

Moreover, one might think the larger effects sizes we derive using the truncated sample may be due to cognitive difficulties or lifestyle factors. The one difference between the truncated and full sample is that the truncated sample includes *less* youth. Prior work shows younger people are less attentive (Alvarez et al., 2019), but they also tend to be more pro-LGBTQ+, so that should ostensibly attenuate effect estimates from a substantive basis but increase effect estimates from the basis of increasing attention. Likewise, if the problem The Influence of Pulse on SSM Support Attenuates Over Time



Figure B29: The Influence of Pulse on SSM Support Attenuates Over Time (Using Full TAPS Sample).

was cognition, then the truncated sample, which is older, should have smaller effects, given older people tend to be more likely to experience cognitive decline (Murman, 2015). We do not observe smaller effects using the older, truncated sample, suggesting cognitive decline may not bias our coefficient estimates. Therefore, we feel confident our truncation exercise is removing respondents inattentive to survey content.

B.4 Baseline Covariate Measurement

Age is a 4 category index from 0-3 characterizing respondents aged 18-29, 30-44, 45-59, 60+. Scaled between 0-1.

White is a binary indicator equal to 1 if the respondent indicates that "white" is a race they currently identify as.

Woman is a binary indicator equal to 1 if the respondent indicates they are "female" in response to a question asking if they are female or male.

Child is a binary indicator equal to 1 if the respondent indicates they have children in response to an item asking if they have biological or adopted children.

Non-religious is a binary indicator equal to 1 if the respondent indicates they are "not religious" in response to an item asking if they consider themself Christian, Jewish, Muslim, Buddhist, Hindu, or another religion.

Married is a binary indicator equal to 1 if the respondent did not indicate they were divorced, widowed, separated from their partner, or never married.

Income is a 0-5 scale of the respondents self-reported household income from < \$10,000, \$10-29,999, \$30-49,999, \$50-79,999, \$80-99,999, \$100,000 or more. Scaled between 0-1.

College is a binary indicator equal to 1 if the respondent reports the highest level of school they have completed is at or above a bachelor's degree.

Unemployment is a binary indicator equal to 1 if the respondent reports they are not working at a job for pay.

Union is a binary indicator equal to 1 if the respondent reports they or someone in their household is a member of a labor union.

Rent is a binary indicator equal to 1 if the respondent reports they rent when asked if they rent or own their home.

Internet Access is a binary indicator equal to 1 if the respondent reports they have household internet access.

Internet Mode is a binary indicator equal to 1 if the respondent was recruited via an online mechanism instead of mail, call-in, or outbound calls.

Liberal is a binary indicator equal to 1 if the respondent indicates they are "slightly liberal," "liberal," or "very liberal" in addition to indicating that they are "liberal if they had to choose" in an additional question conditional on indicating "don't know" or "moderate" in the initial question.

Metro is a binary indicator equal to 1 if the respondent lives in a zipcode that is a metropolitan area.

State indicators (Florida, Texas, California, New York, Pennsylvania) are equal to 1 if the respondent self-reports they live in the respective states.

B.5 SSM Support By Interview Date



Figure B30: Support for Same Sex Marriage (y-axis) Across Interview Dates (x-axis). Vertical line is the moment the Pulse nightclub shooting occurred. Loess models are fit on each side of the moment the Pulse shooting occurred and are weighted based on the interview date sample size. Larger circles denote more interviews on a given date. All covariates re-scaled between 0-1.

B.6 Outcome Item Non-response Balance

 Table B3: Outcome Item Non-response is Balanced Between Pre and Post-Pulse

 Periods

| | SSM Item Non-Response |
|----------------|-----------------------|
| Post-Pulse | 0.008 |
| | (0.005) |
| \mathbb{R}^2 | 0.003 |
| Ν | 1142 |

Note: ***p < 0.001, **p < 0.01, *p < 0.05. HC2 robust standard errors in parentheses.

B.7 Ruling Out Pre-Treatment Time Trends

Table B4: The Effect of Pulse On SSM Support is Not Driven by Pre-TreatmentTime Trends

| | SSM Support |
|--------------------|-------------|
| Post-Pulse Placebo | -0.035 |
| | (0.063) |
| \mathbb{R}^2 | 0.001 |
| N | 679 |

Note: *** p < 0.001, ** p < 0.01, *p < 0.05. HC2 robust standard errors in parentheses.

B.8 Regression Tables

B.8.1 Balance Plot

| Outcome | Post-Pulse Coef. | \mathbf{SE} | р | \mathbf{N} |
|-----------------|------------------|---------------|------|--------------|
| Age | -0.18 | 0.03 | 0.00 | 1142 |
| White | -0.06 | 0.04 | 0.17 | 1142 |
| Woman | 0.04 | 0.05 | 0.42 | 1142 |
| Non-religious | 0.06 | 0.04 | 0.16 | 1142 |
| Married | 0.04 | 0.03 | 0.13 | 1142 |
| Child | -0.06 | 0.05 | 0.18 | 1142 |
| Income | -0.02 | 0.03 | 0.48 | 1142 |
| College | -0.06 | 0.03 | 0.07 | 1140 |
| Unemployment | -0.05 | 0.05 | 0.27 | 1142 |
| Union | 0.01 | 0.03 | 0.69 | 1142 |
| Rent | 0.04 | 0.04 | 0.39 | 1142 |
| Internet Access | -0.01 | 0.04 | 0.79 | 1142 |
| Internet Mode | 0.01 | 0.05 | 0.80 | 1142 |
| Liberal | 0.06 | 0.05 | 0.18 | 1142 |
| Metro | 0.01 | 0.04 | 0.78 | 1142 |
| Florida | 0.01 | 0.02 | 0.62 | 1142 |
| Texas | 0.02 | 0.03 | 0.56 | 1142 |
| California | -0.03 | 0.03 | 0.18 | 1142 |
| New York | 0.03 | 0.02 | 0.17 | 1142 |
| Pennsylvania | -0.01 | 0.02 | 0.66 | 1142 |

Table B5: Balance Plot for TAPS data.

HC2 robust SEs displayed. Each coefficient is from a separate regression where the outcome is on the left hand side of the linear regression and the *post-Pulse* indicator is on the right hand side of the regression.

B.8.2 Post-Pulse Influence on SSM Support

| | SSM S | Support |
|-----------------|--------|---------------|
| | (1) | (2) |
| Post-Pulse | 0.13** | 0.10^{*} |
| | (0.05) | (0.04) |
| Age | | 0.02 |
| | | (0.07) |
| White | | 0.20^{***} |
| | | (0.05) |
| Woman | | 0.05 |
| | | (0.04) |
| Non-religious | | 0.25^{***} |
| | | (0.06) |
| Married | | -0.01 |
| C1 11 1 | | (0.05) |
| Child | | -0.09 |
| T | | (0.05) |
| Income | | (0.03) |
| Cullum. | | (0.07) |
| College | | $(0.01)^{-1}$ |
| Unomployed | | (0.04) |
| Unemployed | | -0.05 |
| Union | | (0.05) |
| Omon | | (0.05) |
| Renter | | 0.03 |
| rtenter | | (0.05) |
| Internet Access | | -0.02 |
| | | (0.05) |
| Internet Mode | | 0.01 |
| | | (0.04) |
| Liberal | | 0.38^{***} |
| | | (0.04) |
| Metro Area | | 0.06 |
| | | (0.05) |
| State FE | Ν | Y |
| \mathbb{R}^2 | 0.02 | 0.35 |
| Ν | 1134 | 1132 |

Table B6: Support for Same Sex Marriage Increases After Pulse

Note: $^{***}p < 0.001,\,^{**}p < 0.01,\,^{*}p < 0.05.$ HC2 robust standard errors in parentheses.

B.8.3 Falsification Tests

| | Increase Taxes (1) | Common Core (2) | Citizen Pathway (3) | Abortion (4) | Build Keystone (5) | Repeal ACA (6) | Cap Emissions (7) |
|---------------------|-----------------------|--------------------|------------------------|------------------|-----------------------|-------------------|----------------------|
| Post-Pulse | -0.02 (0.05) | $0.02 \\ (0.05)$ | $0.01 \\ (0.05)$ | $0.05 \\ (0.05)$ | -0.02 (0.05) | -0.06 (0.05) | -0.02 (0.05) |
| \mathbb{R}^2 N | $0.00 \\ 1135$ | $0.00 \\ 1138$ | $0.00 \\ 1137$ | $0.00 \\ 1132$ | $0.00 \\ 1136$ | $0.00 \\ 1137$ | $0.00 \\ 1135$ |

| Table B7: LGBTQ-Irrelevant | Attitudes | Do | Not | Change | Post-Pulse |
|------------------------------------|-----------|----|-----|--------|------------|
|------------------------------------|-----------|----|-----|--------|------------|

Note: ***p < 0.001, **p < 0.01, *p < 0.05. HC2 robust standard errors in parentheses.

B.8.4 Temporal Persistence

| Post-Pulse Coef. | SE | p-value | Ν | Days Cut | Controls |
|------------------|------|---------|------|----------|----------|
| 0.13 | 0.05 | 0.01 | 1111 | 1 | No |
| 0.14 | 0.05 | 0.01 | 1020 | 2 | No |
| 0.15 | 0.05 | 0.01 | 980 | 3 | No |
| 0.15 | 0.05 | 0.01 | 956 | 4 | No |
| 0.17 | 0.06 | 0.00 | 927 | 5 | No |
| 0.16 | 0.06 | 0.01 | 907 | 6 | No |
| 0.16 | 0.06 | 0.01 | 888 | 7 | No |
| 0.17 | 0.06 | 0.01 | 878 | 8 | No |
| 0.14 | 0.07 | 0.04 | 842 | 9 | No |
| 0.09 | 0.07 | 0.20 | 818 | 10 | No |
| 0.04 | 0.08 | 0.59 | 799 | 11 | No |
| 0.05 | 0.08 | 0.55 | 791 | 12 | No |
| 0.03 | 0.08 | 0.72 | 787 | 13 | No |
| 0.01 | 0.08 | 0.93 | 780 | 14 | No |
| -0.00 | 0.08 | 1.00 | 775 | 15 | No |
| -0.03 | 0.08 | 0.66 | 762 | 16 | No |
| 0.01 | 0.09 | 0.87 | 753 | 17 | No |
| 0.07 | 0.09 | 0.47 | 744 | 18 | No |
| 0.05 | 0.09 | 0.57 | 737 | 19 | No |
| 0.01 | 0.10 | 0.92 | 725 | 20 | No |
| 0.08 | 0.11 | 0.43 | 717 | 21 | No |
| 0.09 | 0.12 | 0.46 | 708 | 22 | No |
| 0.05 | 0.14 | 0.72 | 701 | 23 | No |
| -0.10 | 0.19 | 0.58 | 688 | 24 | No |
| 0.10 | 0.04 | 0.02 | 1109 | 1 | Yes |
| 0.12 | 0.05 | 0.01 | 1018 | 2 | Yes |
| 0.13 | 0.05 | 0.01 | 979 | 3 | Yes |
| 0.13 | 0.05 | 0.02 | 955 | 4 | Yes |
| 0.14 | 0.05 | 0.01 | 926 | 5 | Yes |
| 0.15 | 0.06 | 0.01 | 906 | 6 | Yes |
| 0.14 | 0.06 | 0.02 | 887 | 7 | Yes |
| 0.14 | 0.06 | 0.02 | 877 | 8 | Yes |
| 0.12 | 0.07 | 0.08 | 841 | 9 | Yes |
| 0.05 | 0.06 | 0.40 | 817 | 10 | Yes |
| -0.02 | 0.06 | 0.76 | 798 | 11 | Yes |
| -0.03 | 0.06 | 0.68 | 790 | 12 | Yes |
| -0.04 | 0.06 | 0.51 | 786 | 13 | Yes |
| -0.04 | 0.06 | 0.49 | 779 | 14 | Yes |
| -0.05 | 0.07 | 0.47 | 774 | 15 | Yes |
| -0.08 | 0.07 | 0.25 | 761 | 16 | Yes |
| -0.08 | 0.08 | 0.34 | 752 | 17 | Yes |
| 0.02 | 0.05 | 0.71 | 743 | 18 | Yes |
| 0.01 | 0.05 | 0.82 | 736 | 19 | Yes |
| 0.00 | 0.06 | 0.97 | 724 | 20 | Yes |
| 0.04 | 0.07 | 0.51 | 716 | 21 | Yes |
| 0.07 | 0.07 | 0.34 | 707 | 22 | Yes |
| 0.05 | 0.07 | 0.48 | 700 | 23 | Yes |
| 0.05 | 0.10 | 0.58 | 687 | 24 | Yes |

 Table B8: Table Characterizing Post-Pulse Coefficients Cutting Days Immediately After the Pulse Massacre

B.8.5 Temporal Persistence (Control Coefficients)

Table B9: Control Coefficients For Models Characterizing Temporal Durability of Post-Pulse Effect (Part 1)

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Table B10: Control Coefficients For Models Characterizing Temporal Durability of Post-Pulse Effect (Part 2)

| Renter Internet Access Internet Mode | | | | | |
|--|---|--|---|--|--|
| Internet Access Internet Mode | 0.05 | 0.07 | 0.47 | 8 | 877 |
| Internet Mode | -0.04 | 0.06 | 0.52 | 8 | 877 |
| Lib and | 0.03 | 0.04 | 0.55 | 8 | 877 |
| Liberal Motro Area | 0.35 | 0.05 | 0.00 | 8 | 811 |
| Florida | 0.19 | 0.13 | 0.14 | ŝ | 877 |
| Texas | -0.07 | 0.08 | 0.41 | 8 | 877 |
| California | 0.16 | 0.11 | 0.14 | 8 | 877 |
| New York | 0.18 | 0.10 | 0.06 | 8 | 877 |
| Pennsylvania | -0.01 | 0.14 | 0.94 | 8 | 877 |
| White | 0.21 | 0.07 | 0.00 | 9 | 841 |
| Woman | 0.08 | 0.05 | 0.11 | 9 | 841 |
| Nonreligious | 0.23 | 0.07 | 0.00 | 9 | 841 |
| Married | -0.00 | 0.06 | 0.99 | 9 | 841 |
| Incomo | -0.07 | 0.06 | 0.19 | 9 | 841 |
| College | 0.13 | 0.04 | 0.00 | 9 | 841 |
| Unemployed | -0.00 | 0.06 | 0.97 | 9 | 841 |
| Union | -0.09 | 0.06 | 0.13 | 9 | 841 |
| Renter | 0.04 | 0.07 | 0.56 | 9 | 841 |
| Internet Access | -0.04 | 0.06 | 0.53 | 9 | 841 |
| Internet Mode | 0.02 | 0.04 | 0.70 | 9 | 841 |
| Metro Area | 0.05 | 0.06 | 0.45 | 9 | 841 |
| Florida | 0.24 | 0.13 | 0.07 | 9 | 841 |
| Texas | -0.04 | 0.08 | 0.58 | 9 | 841 |
| California | 0.19 | 0.11 | 0.09 | 9 | 841 |
| New York Demonstration | 0.20 | 0.10 | 0.04 | 9 | 841 |
| Aco | 0.00 | 0.10 | 0.99 | 10 | 917 |
| White | 0.24 | 0.06 | 0.00 | 10 | 817 |
| Woman | 0.06 | 0.04 | 0.21 | 10 | 817 |
| Nonreligious | 0.22 | 0.08 | 0.00 | 10 | 817 |
| Married | -0.02 | 0.06 | 0.72 | 10 | 817 |
| Cmid | -0.09 | 0.06 | 0.14 | 10 | 817 |
| Colloro | 0.03 | 0.09 | 0.76 | 10 | 817 |
| Unemployed | 0.03 | 0.05 | 0.54 | 10 | 817 |
| Union | -0.09 | 0.06 | 0.13 | 10 | 817 |
| Renter | -0.00 | 0.06 | 1.00 | 10 | 817 |
| Internet Access | 0.01 | 0.05 | 0.78 | 10 | 817 |
| Internet Mode | 0.05 | 0.04 | 0.23 | 10 | 817 |
| Metro Area | 0.38 | 0.05 | 0.00 | 10 | 017 817 |
| Florida | 0.26 | 0.14 | 0.05 | 10 | 817 |
| Texas | -0.01 | 0.07 | 0.89 | 10 | 817 |
| California | 0.19 | 0.11 | 0.10 | 10 | 817 |
| New York | 0.23 | 0.09 | 0.01 | 10 | 817 |
| Pennsylvania | 0.03 | 0.16 | 0.87 | 10 | 817 |
| Age White | 0.03 | 0.08 | 0.74 | 11 | 798 |
| Woman | 0.24 | 0.06 | 0.00 | 11 | 708 |
| Nonreligious | 0.07 | 0.04 | 0.00 | 11 | 798 |
| Married | -0.03 | 0.06 | 0.55 | 11 | 798 |
| Child | -0.09 | 0.06 | 0.12 | 11 | 798 |
| Income | 0.05 | 0.09 | 0.57 | 11 | 798 |
| College | 0.13 | 0.04 | 0.00 | 11 | 798 |
| Unemployed | 0.03 | 0.05 | 0.62 | 11 | 798 |
| Union | -0.00 | 0.05 | 0.26 | 11 | 798 |
| Internet Access | 0.01 | 0.07 | 0.85 | 11 | 798 |
| Internet Mode | 0.05 | 0.04 | 0.25 | 11 | 798 |
| Liberal | 0.39 | 0.05 | 0.00 | 11 | 798 |
| Metro Area | 0.07 | 0.05 | 0.19 | 11 | 798 |
| Florida | 0.23 | 0.15 | 0.14 | 11 | 798 |
| Texas | -0.02 | 0.07 | 0.83 | 11 | 798 |
| Cantorma New York | 0.19 | 0.11 | 0.08 | 11 | 798 |
| New IOrk Donnerdannia | 0.20 | 0.10 | 0.01 | 11 | 798 |
| Age | 0.02 | 0.08 | 0.84 | 12 | 790 |
| White | 0.23 | 0.06 | 0.00 | 12 | 790 |
| Woman | 0.07 | 0.04 | 0.10 | 12 | 790 |
| Nonreligious | 0.22 | 0.08 | 0.00 | 12 | 790 |
| Child | -0.03 | 0.06 | 0.02 | 12 | 790 |
| Income | 0.05 | 0.09 | 0.56 | 12 | 790 |
| College | 0.14 | 0.04 | 0.00 | 12 | 790 |
| Unemployed | 0.02 | 0.05 | 0.65 | 12 | 790 |
| Union | -0.08 | 0.06 | 0.18 | 12 | 790 |
| Renter | 0.00 | 0.07 | 0.99 | 12 | 790 |
| Internet Mode | -0.01 | 0.03 | 0.00 | 12 | 790 |
| Liberal | 0.38 | 0.05 | 0.00 | 12 | 790 |
| Metro Area | 0.07 | 0.05 | 0.20 | 12 | 790 |
| Florida | 0.23 | 0.16 | 0.15 | 12 | 790 |
| Texas | -0.01 | 0.07 | 0.83 | 12 | 790 |
| California | 0.19 | 0.11 | 0.08 | 12 | 790 |
| New York Donneylynnia | 0.25 | 0.10 | 0.01 | 12 | 790 |
| Are | 0.04 | 0.08 | 0.57 | 13 | 786 |
| White | 0.23 | 0.06 | 0.00 | 13 | 786 |
| Woman | 0.07 | 0.04 | 0.10 | 13 | 796 |
| Nonreligious | 0.23 | 0.08 | 0.00 | 13 | |
| Married | -0.03 | 0.06 | 0.62 | 13 | 786 |
| Cmid | -0.12 | 0.06 | 0.04 | | 786 786 |
| College | 0.14 | 0.04 | 0.44 | 13 | 786 786 786 786 |
| | 0.00 | | 0.44 | 13 13 13 | 786 786 786 786 786 |
| Unemployed | 0.03 | 0.05 | 0.44 0.00 0.56 | 13 13 13 13 | 786 786 786 786 786 786 |
| Unemployed Union | -0.07 | 0.05 | 0.44 0.00 0.56 0.24 | 13 13 13 13 13 | 786 786 786 786 786 786 786 |
| Unemployed Union Renter | -0.03 -0.07 0.02 | 0.05 0.06 0.07 | 0.44 0.00 0.56 0.24 0.82 | 13 13 13 13 13 13 | 786 786 786 786 786 786 786 786 786 |
| Unemployed Union Renter Internet Access | -0.03 -0.07 -0.02 -0.02 | 0.05 0.06 0.07 0.05 | 0.44 0.00 0.56 0.24 0.82 0.71 | 13 13 13 13 13 13 13 | 786 786 786 786 786 786 786 786 786 786 |
| Unemployed Union Renter Internet Access Internet Mode Liberal | -0.03 -0.02 -0.02 -0.02 0.05 0.27 | 0.05 0.06 0.07 0.05 0.04 0.05 | 0.44 0.00 0.56 0.24 0.82 0.71 0.26 0.00 | 13 13 13 13 13 13 13 13 13 13 13 | 786 786 786 786 786 786 786 786 786 786 |
| Unemployed Union Renter Internet Access Internet Mode Liberal Metro Area | -0.07 0.02 -0.02 0.05 0.37 0.06 | 0.05 0.06 0.07 0.05 0.04 0.05 0.05 | 0.44 0.00 0.56 0.24 0.82 0.71 0.26 0.00 0.28 | 13 13 13 13 13 13 13 13 13 13 13 13 | 786 786 786 786 786 786 786 786 786 786 |
| Unemployed Union Renter Internet Access Internet Mode Liberal Metro Area Florida | -0.07 0.02 -0.02 0.05 0.37 0.06 0.23 | 0.05 0.06 0.07 0.05 0.04 0.05 0.05 0.05 0.15 | 0.44 0.00 0.56 0.24 0.82 0.71 0.26 0.00 0.28 0.14 | 13 13 13 13 13 13 13 13 13 13 13 13 13 | 786 786 786 786 786 786 786 786 786 786 |
| Unemployed Union Renter Internet Access Internet Mode Liberal Metro Area Florida Texas | -0.03 -0.02 -0.02 0.05 0.37 0.06 0.23 -0.00 | 0.05 0.06 0.07 0.05 0.04 0.05 0.05 0.15 0.07 | 0.44 0.00 0.56 0.24 0.71 0.26 0.00 0.28 0.14 0.97 | 13 13 13 13 13 13 13 13 13 13 13 13 13 1 | 786 786 786 786 786 786 786 786 786 786 |
| Unemployed Union Renter Internet Access Internet Mode Liberal Metro Area Florida Texas California Mem Vev [*] | -0.03 -0.07 -0.02 -0.02 0.05 0.37 0.06 0.23 -0.00 0.20 0.20 | 0.05 0.06 0.07 0.05 0.04 0.05 0.05 0.05 0.15 0.07 0.11 | 0.44 0.00 0.56 0.24 0.82 0.71 0.26 0.00 0.28 0.14 0.97 0.06 0.97 | 13 13 13 13 13 13 13 13 13 13 13 13 13 1 | 786 786 786 786 786 786 786 786 786 786 |
| Unemployed Union Renter Internet Access Internet Mode Liberal Metro Area Florida Texas California New York Peompedveret | -0.03 -0.07 0.02 0.05 0.37 0.06 0.23 -0.00 0.20 0.27 0.07 | 0.05 0.06 0.07 0.05 0.04 0.05 0.05 0.05 0.05 0.07 0.11 0.10 | 0.44 0.00 0.56 0.24 0.82 0.71 0.26 0.00 0.28 0.14 0.97 0.06 0.01 0.77 | 13 13 13 13 13 13 13 13 13 13 13 13 13 1 | 786 786 786 786 786 786 786 786 786 786 |
| Unemployed Union Renter Internet Access Internet Mode Liberal Metro Area Florida Texas California New York Pennsylvania | -0.03 -0.07 0.02 -0.02 0.05 0.37 0.06 0.23 -0.00 0.20 0.27 0.05 0.05 | 0.05 0.06 0.07 0.05 0.05 0.05 0.05 0.05 0.07 0.11 0.10 0.17 | 0.44 0.00 0.56 0.24 0.82 0.71 0.26 0.00 0.28 0.14 0.97 0.06 0.01 0.76 0.76 | 13 13 13 13 13 13 13 13 13 13 13 13 13 1 | 786 786 786 786 786 786 786 786 786 786 |
| Unemployed Union Renter Internet Access Internet Mode Liberal Metro Area Florida Texas California New York Pennsylvania Age White | -0.07 0.02 -0.02 0.05 0.37 0.06 0.23 -0.00 0.20 0.27 0.05 0.04 0.24 | 0.05 0.06 0.07 0.05 0.05 0.05 0.05 0.05 0.07 0.11 0.10 0.17 0.08 0.06 | 0.44 0.00 0.56 0.24 0.82 0.71 0.26 0.00 0.28 0.14 0.97 0.06 0.01 0.76 0.03 0.058 0.00 | 13 13 13 13 13 13 13 13 13 13 13 13 13 1 | 786 786 786 786 786 786 786 786 786 786 |
| Unemployed Union Renter Internet Access Internet Mode Liberal Metro Area Florida Texas California New York Pennsylvania Ags White Woman | -0.07 0.02 -0.02 0.05 0.37 0.06 0.23 -0.00 0.20 0.27 0.05 0.04 0.24 0.08 | 0.05 0.06 0.07 0.05 0.05 0.05 0.05 0.07 0.11 0.10 0.17 0.08 0.06 0.04 | 0.44 0.00 0.56 0.24 0.71 0.26 0.00 0.28 0.14 0.97 0.06 0.01 0.76 0.00 0.58 0.00 0.76 0.00 0.01 0.76 | 13 13 13 13 13 13 13 13 13 13 13 13 13 1 | 786 786 786 786 786 786 786 786 786 786 |
| Unemployed Union Renter Internet Access Internet Acde Liberal Metro Area Florida Texas California New York Pennsylvania Age White Woman Nomeligious | -0.07 0.02 -0.02 0.05 0.37 0.06 0.23 -0.00 0.20 0.27 0.05 0.04 0.24 0.08 0.22 | 0.05 0.06 0.07 0.05 0.05 0.05 0.05 0.05 0.07 0.11 0.10 0.17 0.08 0.06 0.04 0.04 0.08 | 0.44 0.00 0.56 0.24 0.71 0.26 0.02 0.28 0.14 0.97 0.06 0.01 0.76 0.01 0.76 0.00 0.07 0.00 | 13 13 13 13 13 13 13 13 13 13 13 13 13 1 | 786 786 786 786 786 786 786 786 786 786 |
| Unemployed Union Renter Internet Access Internet Access Liberal Metro Area Florida Texas California New York Pennsylvania Age White Woman Nonreligious Married | -0.03 -0.07 0.02 -0.02 0.05 0.037 0.06 0.23 -0.00 0.20 0.27 0.05 0.04 0.24 0.24 0.22 -0.03 | 0.05 0.06 0.07 0.05 0.05 0.05 0.05 0.05 0.07 0.11 0.10 0.17 0.08 0.06 0.04 0.08 0.06 | 0.44 0.00 0.56 0.24 0.71 0.26 0.00 0.28 0.14 0.97 0.06 0.01 0.76 0.00 0.07 0.00 0.07 0.00 0.07 0.00 0.05 | 13 13 13 13 13 13 13 13 13 13 13 13 13 1 | 786 786 786 786 786 786 786 786 786 786 |
| Unemployed Union Renter Internet Accesss Internet Mode Liberal Metro Area Florida Tecas California New York Pennsylvania Age White Opennsylvania Age White Moran Nonreligious Married Child Leave- | -0.03 -0.07 0.02 0.05 0.37 0.06 0.23 -0.00 0.20 0.27 0.05 0.04 0.24 0.24 0.22 -0.03 -0.03 -0.03 -0.03 -0.03 -0.04 0.22 -0.04 0.22 -0.05 | 0.05 0.06 0.07 0.05 0.05 0.05 0.05 0.05 0.07 0.11 0.10 0.17 0.08 0.06 0.04 0.08 0.06 0.06 0.06 | 0.44 0.00 0.56 0.24 0.71 0.26 0.00 0.28 0.14 0.97 0.06 0.01 0.71 0.71 0.06 0.01 0.71 0.71 0.06 0.00 0.07 0.06 0.00 0.07 0.06 0.00 0.05 0.00 0.28 0.14 0.97 0.06 0.00 0.05 0.05 | 13 13 13 13 13 13 13 13 13 13 13 13 13 1 | 786 786 786 786 786 786 786 786 786 786 |
| Unemployed Union Renter Internet Access Internet Mode Liberal Metro Area Florida Texas California New York Pennsylvania Age White Woman Nonreligious Married Child Income Collece | -0.03 -0.07 0.02 0.05 0.37 0.06 0.23 -0.00 0.20 0.27 0.05 0.04 0.24 0.24 0.22 -0.03 -0.03 -0.03 -0.03 -0.03 0.22 -0.03 -0.02 0.05 0.24 0.05 0.24 0.05 0.24 0.05 0.25 0.27 0.05 0.20 0.25 0.25 0.25 0.25 0.25 0.25 | 0.05 0.06 0.07 0.05 0.04 0.05 0.05 0.05 0.07 0.11 0.10 0.10 0.06 0.04 0.08 0.06 0.06 0.06 0.06 0.06 0.06 | 0.44 0.00 0.56 0.24 0.71 0.26 0.00 0.28 0.14 0.97 0.06 0.01 0.76 0.00 0.28 0.01 0.70 0.00 0.28 0.00 0.28 0.00 0.26 0.00 0.28 0.00 0.26 0.00 0.28 0.00 0.28 0.00 0.28 0.00 0.28 0.00 0.28 0.00 0.00 0.28 0.00 | 13 13 13 13 13 13 13 13 13 13 13 13 13 1 | 786 786 786 786 786 786 786 786 786 786 |
| Unemployed Union Renter Internet Accesss Internet Mode Liberal Metro Area Florida Texas California New York Pennsylvania Age Wointe Work Pennsylvania Age White Woman Noureligious Married Child Income College Unemployed | -0.03 -0.02 -0.02 0.05 0.37 0.06 0.23 -0.00 0.20 0.20 0.20 0.20 0.20 0.20 0.2 | 0.05 0.06 0.07 0.05 0.04 0.05 0.05 0.05 0.05 0.05 0.07 0.11 0.00 0.07 0.11 0.08 0.06 0.04 0.08 0.06 0.04 0.08 0.06 0.04 0.05 0.04 0.05 0.07 0.05 0.05 0.05 0.05 0.05 0.05 | 0.44 0.00 0.56 0.24 0.82 0.71 0.26 0.00 0.82 0.71 0.26 0.00 0.82 0.00 0.82 0.00 0.28 0.00 0.26 0.00 0.26 0.00 0.26 0.00 0.26 0.00 0.26 0.00 0.26 0.00 0.26 0.00 0.26 0.00 0.26 0.00 0.14 0.00 0.26 0.00 0.14 0.00 0.00 0.14 0.00 0.00 0.14 0.00 0.00 0.14 0.00 0.01 0.00 0.01 0.01 0.01 0.01 0.01 0.00 0.01 0.00 0.01 0.00 0.01 0.00 0.01 0.00 0.00 0.00 0.00 0.01 0.00 | 13 13 13 13 13 13 13 13 13 13 13 13 13 1 | 786 786 786 786 786 786 786 786 786 786 |
| Unemployed Union Renter Internet Access Internet Mode Liberal Metro Area Plorida Texas California New York Pennsylvania Age White Woman Nonreligious Married Child Income College Unemployed Union | 0.03 -0.07 0.02 0.05 0.37 0.06 0.23 0.23 0.23 0.20 0.20 0.27 0.05 0.27 0.05 0.27 0.05 0.27 0.05 0.27 0.05 0.27 0.02 0.27 0.02 0.27 0.02 0.02 0.23 0.23 0.23 0.22 0.02 0.23 0.02 0.23 0.02 0.23 0.02 0.02 | 0.05 0.06 0.07 0.05 0.05 0.05 0.05 0.05 0.05 0.05 | $\begin{array}{c} 0.44\\ 0.00\\ 0.56\\ 0.24\\ 0.82\\ 0.71\\ 0.26\\ 0.00\\ 0.82\\ 0.71\\ 0.26\\ 0.00\\ 0.34\\ 0.00\\ 0.65\\ 0.00\\ 0.65\\ 0.00\\ 0.65\\ 0.00\\ 0.65\\ 0.00\\ 0.63\\ 0.00\\ 0.63\\ 0.26\\ \end{array}$ | $\begin{array}{cccccccccccccccccccccccccccccccccccc$ | 786 786 786 786 786 786 786 786 786 786 |
| Unemployed Union Renter Internet Access Internet Access Internet Mode Liberal Herner Florida Tecas California New York Penmsylvania Age Penmsylvania Age Woman New York Penmsylvania Age Woman New York Penmsylvania Age Woman New York Penmsylvania Age Woman New York Penmsylvania Califor Informa Califor Informa Califor Union Renter Center Center Conter C | 0.033 -0.07 -0.02 0.022 0.05 0.37 0.06 0.23 0.05 0.37 0.06 0.23 0.05 0.37 0.06 0.23 0.05 0.37 0.06 0.23 0.20 0.27 0.05 0.08 0.22 -0.03 -0.01 0.07 0.13 -0.06 0.02 0.02 | 0.05 0.06 0.07 0.05 0.05 0.05 0.05 0.05 0.05 0.05 | 0.44 0.00 0.24 0.24 0.24 0.24 0.26 0.28 0.26 0.28 0.14 0.26 0.28 0.14 0.97 0.06 0.58 0.00 0.07 0.06 0.01 0.01 0.05 0.00 0.01 0.24 0.24 0.25 0.24 0.24 0.25 0.26 0.28 0.00 0.07 0.06 0.00 0.05 0.00 0.05 0.00 0.05 0.00 0.05 0.00 0.05 0.00 0.05 0.00 0.05 0.00 0.05 0.00 0.05 0.00 0.05 0.06 0.06 0.08 0.00 0.05 0.06 0.06 0.06 0.06 0.06 0.08 0.06 0.08 0.06 0.08 0.06 0.08 0.06 0.08 0.06 0.08 0.06 0.08 0.06 0.08 0.06 0.08 0.06 0.08 0.06 0.08 0.06 0.06 0.06 0.08 0.06 0.08 0.06 0.06 0.06 0.08 0.06 0.08 0.06 0.08 0.06 0.08 0 | $\begin{array}{cccccccccccccccccccccccccccccccccccc$ | 786 786 786 786 786 786 786 786 786 786 |
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| Usemployed Using Using Internet Mode Liberal Anne Person Person Versio | 0.003 -0.07 0.022 -0.02 0.035 0.037 -0.02 0.037 0.04 0.20 0.207 0.04 0.207 0.04 0.237 0.04 0.247 0.04 0.23 -0.01 0.06 0.23 -0.01 0.06 0.23 -0.02 0.04 0.24 0.04 0.05 0.04 0.05 0.04 0.02 0.04 0.04 0.02 0.04 0.02 0.04 0.02 0.04 0.02 0.04 0.02 0.04 0.02 0.04 0.02 0.04 0.02 0.04 0.02 0.04 0.04 0.02 0.04 0.02 0.04 0.02 0.04 0.02 0.04 0.04 0.02 0.04 0.04 0.02 0.04 0.04 0.02 0.04 0.04 0.04 0.02 0.04 0.04 0.04 0.04 0.04 0.02 0.04 0.022 0.05 0.04 0.02 0.04 0.02 0.04 0.04 0.02 0.04 0.04 0.02 0.04 0.02 0.04 0.04 0.04 0.02 0.04 0.04 0.02 0.04 0.04 0.02 0.04 0.04 0.02 0.04 0.04 0.02 0.04 0.04 0.02 0.04 0.02 0.04 0.02 0.04 0.02 0.04 0.04 0.02 0.04 0.02 0.04 0.02 0.04 0.02 0.04 0.04 0.02 0.04 0.06 0.02 0.02 0.02 0.02 0.02 0.02 0.02 0.02 0.02 0.02 0.02 0.02 0.02 0.02 0.02 0.02 0.02 0.05 0.05 0.04 0.06 0.05 0.05 0.05 0.04 0.06 0.05 0.05 0.05 0.05 0.04 0.05 0.05 0.05 0.05 0.04 0.05 0.05 0.05 0.05 0.04 0.05 0. | 0.05 0.06 0.07 0.05 0.05 0.05 0.05 0.05 0.05 0.05 | 0.44 0.44 0.00 0.564 0.24 0.711 0.26 0.82 0.711 0.26 0.82 0.711 0.26 0.82 0.711 0.26 0.82 0.000 0.28 0.000 0.28 0.001 0.001 0.001 0.011 0.768 0.000 0.026 0.000 0.655 0.000 0.655 0.000 0.655 0.000 0.655 0.000 0.655 0.000 0.655 0.000 0.655 0.000 0.655 0.000 0.655 0.000 0.655 0.000 0.655 0.000 0.655 0.000 0.000 0.655 0.0000 0.0000 0.0000 0.0000 0.0000 0.0000 0.00000 0.0000000000 | $\begin{array}{c} 13\\ 13\\ 13\\ 13\\ 13\\ 13\\ 13\\ 13\\ 13\\ 13\\$ | |
| Unemployed Union Union Internet Mode Liberal Liberal Harton Access Internet Mode California New York Peoploynamic New York Peoploynamic New York Peoploynamic New York New York California New York New Y | 0.003 -0.007 0.022 -0.022 0.035 0.037 -0.002 0.020 0.230 0.200 0.0000 0.0000 0.000 0.0000 0.0000 0.0000 0 | 0.05 0.06 0.06 0.05 0.05 0.05 0.05 0.05 | 0.44 0.44 0.60 0.56 0.24 0.71 0.82 0.71 0.92 0.00 0.00 0.00 0.00 0.14 0.97 0.00 0.14 0.97 0.00 0.14 0.97 0.00 0.14 0.97 0.00 0.15 0.00 0.01 0.05 0.00 0.00 0.05 0.00 | $\begin{smallmatrix} 13\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3$ | $\frac{1}{100} \frac{1}{100} \frac{1}$ |
| Unemployed Union Union Internet Mode Liberal Vertext Vertext Vertext New York White White White Woman Woman Woman Woman Woman Woman Woman Woman Union California Woman Woman Union Union California Marriool California Marriool California Marriool California Marroo California Metro Area Piotcha Tecas Metro Area Piotcha Tecas Metro Area Piotcha Metro Area Piotcha Metro Area Piotcha Marroo Marriool Metro Area Piotcha Metro Area Metro Area Piotcha Metro Area Piotcha Metr | 0.03 0.04 0.00 0.02 0.05 0.03 0.06 0.23 0.06 0.23 0.04 0.02 0.02 0.04 0.08 0.02 0.02 0.04 0.08 0.04 0.08 0.04 0.08 0.04 0.08 0.04 0.08 0.04 0.08 0.04 0.08 0.04 0.08 0.04 0.08 0.04 0.08 0.04 0.08 0.04 0.08 0.04 0.08 0.04 0.08 0.04 0.08 0.04 0.04 | 0.05 0.06 0.07 0.05 0.05 0.05 0.05 0.05 0.01 0.01 0.01 | 0.44 0.42 0.60 0.24 0.71 0.82 0.71 0.00 0.24 0.01 0.24 0.02 0.20 0.20 0.20 0.20 0.20 0.20 0.20 0.20 0.20 0.20 0.20 0.20 0.20 0.00 0.00 0.01 0.00 0.01 0.00 0.020 0.00 0.010 0.000 0.020 0.0000 0.0000 0.0000 0.0000 0.0000 0.0000 0.00000 0.0000 0.0 | $\begin{smallmatrix} 13\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3$ | |
| Usemployed Usens 200 Internet Access Internet Access Internet Mode Photola Texas Photola Texas Photola Texas Photola Texas Californet Murited Usens Californet Californet Californet Californet Usens Development Californet | 0.003 0.007 0.022 0.05 0.357 0.026 0.230 0.020 0.021 0.021 0.021 0.021 0.021 0.021 0.021 0.021 0.021 0.021 0.021 0.021 0.021 0.021 0.021 0.021 0.022 0.022 0.038 0.028 0.038 0.028 0.038 0.028 0.038 0.028 0.038 0.028 0.038 0.028 0.038 0.028 0.038 0.028 0.0388 0.0388 0.038 0.038 0.038 0.038 0 | 0.05 0.06 0.06 0.07 0.05 0.05 0.05 0.05 0.05 0.05 0.05 | 0.44 0.42 0.60 0.24 0.71 0.82 0.71 0.90 0.82 0.71 0.90 0.82 0.00 0.91 0.90 0.14 0.97 0.00 0.14 0.97 0.00 0.14 0.97 0.00 0.14 0.97 0.00 0.14 0.01 0.158 0.00 0.058 0.00 0.05 0.06 0.061 0.063 0.060 0.063 0.060 0.053 0.000 0.053 0.000 0.053 0.000 0.053 0.000 0.053 0.000 0.053 0.000 0.053 0.000 0.053 0.000 0.053 0.000 0.055 0.001 0.055 0.002 0.002 0.005 0.005 0.005 0.005 0.005 0.005 0.005 0.005 0.005 0.005 0.005 0.0000 0.0000 0.0000 0.0000 0.0000 0.00000 0.00000 0.0000000000 | $\begin{smallmatrix} 13\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3$ | 1786 786 7866 7866 7866 7866 7866 7866 7866 7866 7866 7866 7866 7866 7866 7866 7866 7866 7867 7866 7868 7866 7869 7866 7869 7866 7879 7799 7799 7799 7799 7799 7799 7799 7799 7799 7799 7799 7799 7799 7799 7799 7799 7799 7799 7799 7799 7799 7794 7744 774 774 774 774 774 774 |
| Usenspöyred Usens Tealan Internet Access Internet Access Inter | 0.03 0.007 0.02 0.05 0.37 0.02 0.05 0.37 0.02 0.05 0.23 0.06 0.23 0.06 0.23 0.06 0.23 0.04 0.02 0.05 0.24 0.02 0.05 0.24 0.02 0.05 0.24 0.02 0.05 0.02 0.05 0.04 0.02 0.02 0.05 0.04 0.02 0.02 0.05 0.04 0.02 0.02 0.05 0.02 0.04 0.02 0.02 0.05 0.04 0.02 0.05 0.02 0.05 0.04 0.02 0.02 0.05 0.04 0.02 0.05 0.04 0.02 0.05 0.04 0.02 0.05 0.04 0.02 0.05 0.04 0.02 0.05 0.02 0.05 0.04 0.02 0.05 0.02 0.03 0.03 0.03 0.03 0.03 0.03 0.03 0.03 0.02 0.03 0.03 0.02 0.03 0.02 0.03 0.03 0.02 0.03 0.02 0.03 0.02 0.03 0.03 0.02 0.03 0.02 0.03 0.03 0.03 0.02 0.03 0.05 0.04 0.02 0.03 0.03 0.02 0.03 0.05 0.04 0.02 0.03 0.03 0.05 0.02 0.05 0.04 0.02 0.03 0.03 0.02 0.05 0.02 0.03 0.02 0.05 0.02 0.03 0.05 0.02 0.05 0.02 0.05 0.02 0.03 0.05 0.02 0.05 0.02 0.05 0.02 0.05 0.02 0.05 0.02 0.05 0.02 0.05 0.02 0.05 0.05 0.02 0.05 0.02 0.05 0.05 0.02 0.05 0.03 0.02 0.05 0.02 0.05 0.02 0.05 0.02 0.05 0.02 0.05 0.02 0.05 0.02 0.05 0.02 0.05 0.02 0.05 0.02 0.05 0.02 0.05 0.02 0.05 0.02 0.05 0.02 0.05 0.02 0.05 0.02 0.05 0.02 0.05 0.05 0.02 0.05 0.05 0.02 0.05 | 0.05 0.06 0.06 0.07 0.05 0.05 0.05 0.05 0.05 0.05 0.05 | $\begin{array}{c} 0.44 \\ 0.40 \\ 0.50 \\ 0.24 \\ 0.71 \\ 0.82 \\ 0.71 \\ 0.97 \\ 0.00 \\ 0.98 \\ 0.14 \\ 0.97 \\ 0.00 \\ 0.016 \\ 0.016 \\ 0.00 \\ $ | $\begin{smallmatrix} 13\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3\\ 3$ | $\frac{1}{100} \frac{1}{100} \frac{1}$ |
| Usenspörydd Usensor Usensor Haterset Access Internest Access Internest Access Internest Access Internest Access Particles Particles Particles Particles Weinen Weinen Weinen Weinen Weinen Usenspörydd Usenspörydd Usenspörydd Usenspörydd Usenspörydd Haterset Access Internest Mode Internest Mode Internest Mode Internest Mode Internest Mode Internest Mode Internest Mode Internest Moder Namelydd Namely | 0.003 0.007 0.022 0.05 0.06 0.23 0.06 0.23 0.06 0.23 0.02 0.02 0.02 0.02 0.02 0.02 0.02 | 0.05 0.06 0.06 0.06 0.07 0.05 0.04 0.05 0.05 0.05 0.05 0.01 0.01 0.08 0.06 0.04 0.08 0.04 0.08 0.04 0.08 0.04 0.08 0.04 0.05 0.05 | 0.44 0.45 0.45 0.45 0.45 0.45 0.45 0.45 | $\begin{array}{c} 33\\ 33\\ 33\\ 33\\ 33\\ 33\\ 33\\ 33\\ 33\\ 33$ | |
| Usemployed Usenia Characteria Hateraet Access Internet Access Internet Mode Provide Texas Merco Area Prosta Texas Merco Area Prosta New York New York Proma Neuroitagions Metro Area Prosta New York Promy Yourk Neuroitagions Metro Area Prosta New York Neuroitagions Neuroitagion | 0.003 0.007 0.022 0.05 0.023 0.023 0.023 0.023 0.023 0.023 0.023 0.023 0.023 0.023 0.023 0.023 0.023 0.024 0.022 0.05 0.024 0.022 0.05 0.023 0.024 0.025 0.024 0.024 0.025 0.024 0.025 0.024 0.025 0.024 0.025 0.024 0.026 0.025 0.025 0.025 0.026 | 0.05 0.06 0.06 0.07 0.05 0.05 0.05 0.05 0.07 0.07 0.08 0.06 0.06 0.06 0.06 0.06 0.06 0.06 | 0.44 0.000 0.56 0.024 0.026 0.24 0.026 0.026 0.026 0.028 0.028 0.043 0.007 0.065 0.066 0.060 0.065 0.060 0.067 0.055 0.050 0.065 0.050 0.065 0.050 0.065 0.050 0.057 0.050 0.055 0.050 0.051 0.055 0.050 0.051 0.055 0.050 0.051 0.055 0.055 0.055 0.050 0.055 0.055 0.050 0.055 0.055 0.050 0.055 0.0 | $\begin{smallmatrix} 13&13&13&13&13&13&13&13&13&13&13&13&13&1$ | |
| Usenspöysel Usensor Harberset Access Internet | 0.003 0.007 0.022 0.05 0.023 0.023 0.023 0.023 0.022 0.024 0.024 0.024 0.022 0.027 0.02 0.027 0.02 0.02 0.024 0.024 0.024 0.022 0.024 0.025 0.02 | 0.05 0.06 0.06 0.07 0.05 0.05 0.05 0.05 0.05 0.07 0.01 0.07 0.01 0.06 0.06 0.07 0.01 0.07 0.08 0.06 0.06 0.09 0.04 0.08 0.06 0.08 0.00 0.00 0.00 0.00 0.00 | 0.44 0.000 0.56 0.024 0.822 0.71 0.26 0.029 0.026 0.028 0.028 0.028 0.028 0.028 0.028 0.028 0.026 0.028 0.026 0.02 | $\begin{smallmatrix} 13\\ 13\\ 13\\ 13\\ 13\\ 13\\ 13\\ 13\\ 13\\ 13\\$ | |
| Usemployed Usenar Datarent Access Interest Access Interest Access Interest Control Con | 0.003 0.007 0.022 0.05 0.023 0.023 0.023 0.023 0.023 0.023 0.023 0.024 0.023 0.024 0.023 0.024 0.025 0.024 0.024 0.025 0.024 0.024 0.025 0.024 0.025 0.024 0.025 0.024 0.025 0.024 0.025 0.025 0.024 0.025 0 | 0.05 0.06 0.06 0.07 0.05 0.05 0.05 0.05 0.05 0.07 0.01 0.00 0.00 0.00 0.00 0.00 0.00 | 0.44 0.42 0.62 0.620 | $\begin{smallmatrix} 13 \\ 13 \\ 13 \\ 13 \\ 13 \\ 13 \\ 13 \\ 13 $ | |

Table B11: Control Coefficients For Models Characterizing Temporal Durability of Post-Pulse Effect (Part 3)



Table B12: Control Coefficients For Models Characterizing Temporal Durability of Post-Pulse Effect (Part 4)



B.9 Temporal Placebo Tests



Figure B31: The Effect of Pulse is Unique to 2016. The x-axis is the survey at use. The y-axis is the coefficient for a binary indicator if the respondent was interviewed the calendar day after the Pulse massacre in 2012, 2013, 2016, and 2017 respectively. The outcome for all studies/models is support for same sex marriage. Color denotes the inclusion/exclusion of adjustment for baseline covariates between respondents interviewed before and after the calendar day of the Pulse massacre. All covariates rescaled between 0-1. 95% CIs displayed from HC2 robust standard errors.

B.9.1 Temporal Placebo Test Survey Information

Pew 2012: The 2012 Pew Voter Attitude Survey obtained telephone interviews with a nationally representative sample of N = 2013 adults living in the United States. The interviews were conducted by Princeton Survey Research Associates International between June 7, 2012 to June 17, 2012. The margin of sampling error for the complete set of weighted data is ± 2.6 percentage points. The same sex marriage outcome asks respondents if they "strongly favor, favor, oppose or strongly oppose allowing gays and lesbians to marry legally." The outcome is coded 1 if the respondent indicates strongly favor or favor, 0 otherwise.

CNN 2013: The 2013 CNN poll is a nationally representative survey using landline and cell phone sampling (N = 1014). The poll was in the field between June 11, 2013 and June 13, 2013. The same sex marriage outcome asks respondents if they "think marriages between gay and lesbian couples should or should not be recognized by the law as valid, with the same rights as traditional marriages?" The outcome is coded 1 if the respondent indicates gay and lesbian couples should be recognized by the law, and 0 otherwise.

Pew 2017: The 2017 Pew Political Landscape Survey was in the field between June 8, 2017 and June 18, 2017. It is a nationally representative survey of 2504 respondents. Interviews were conducted via landline and cell phone. The survey was conducted by Princeton Survey Research Associates International. The margin of error is \pm 1.6 percentage points. The same sex marriage outcome asks respondents if they "strongly favor, favor, oppose or strongly oppose allowing gays and lesbians to marry legally." The outcome is coded 1 if the respondent indicates strongly favor or favor, 0 otherwise.

B.10 Alternative Bandwidths

B.11 Ordinal Outcome Re-estimation

| | SSM S (1) | upport (Ordinal) (2) |
|----------------|------------------------|------------------------------|
| Post-Pulse | 0.102^{*} (0.044) | 0.068^{\dagger} (0.038) |
| \mathbb{R}^2 | 0.012 | 0.351 |
| Controls | N | Y |

Table B13: Findings Are Robust To Using Ordinal Outcome

Note: ***p < 0.001, **p < 0.01, *p < 0.05, †p < 0.1. All covariates re-scaled between 0-1. HC2 robust standard errors in parentheses.



Figure B32: The Effect of Pulse is Robust to Alternate Bandwidths. The x-axis is the bandwidth (in days) for the pre and post Pulse period. The y-axis is the coefficient for a binary indicator if the respondent was interviewed after the Pulse nightclub shooting. Color denotes the inclusion/exclusion of control covariates adjusting for covariate imbalance between respondents interviewed before and after the Pulse nightclub shooting. Annotations denote sample size for each estimate in addition to the number of imbalanced covariates. All covariates re-scaled between 0-1. 95% CIs displayed from HC2 robust standard errors.

B.12 Insensitivity to Weighting

| | SSM S | Support |
|-----------------------|--------|------------------|
| | (1) | (2) |
| Post-Pulse | 0.07* | 0.05^{\dagger} |
| | (0.03) | (0.03) |
| Age | . , | -0.09^{*} |
| - | | (0.04) |
| White | | 0.18^{***} |
| | | (0.04) |
| Woman | | 0.08^{**} |
| | | (0.03) |
| Non-religious | | 0.21^{***} |
| | | (0.03) |
| Married | | -0.01 |
| | | (0.03) |
| Child | | -0.08^{*} |
| | | (0.03) |
| Income | | 0.09^{\dagger} |
| | | (0.05) |
| College | | 0.10^{***} |
| ·· , , | | (0.03) |
| Unemployed | | -0.01 |
| | | (0.03) |
| Union | | 0.02 |
| D | | (0.03) |
| Renter | | -0.00 |
| T A | | (0.04) |
| Internet Access | | (0.04) |
| Test surget Marda | | (0.03) |
| Internet Mode | | (0.00) |
| Liboral | | (0.02) |
| Liberal | | (0.44) |
| Motro Aroa | | (0.03) 0.00** |
| metro Area | | (0.09) |
| B ² | 0.00 | 0.38 |
| N | 1134 | 1139 |
| T N | 1104 | 1104 |

Table B14: Support for Same Sex Marriage Increases After Pulse (UnweightedEstimates)

Note: ***p < 0.001, **p < 0.01, *p < 0.05, †p < 0.1. All covariates re-scaled between 0-1. HC2 robust standard errors in parentheses.

| | | SSN | A Suppo | ort | |
|--------------------------------|---------------|---------|---------|---------|---------|
| | (1) | (2) | (3) | (4) | (5) |
| Post-Pulse | 0.08 | 0.14** | 0.10 | 0.11 | 0.13 |
| | (0.04) | (0.05) | (0.05) | (0.12) | (0.07) |
| Post-Pulse x Non-White | 0.07 | | | | |
| Dest Dulas Warner | (0.10) | 0.00 | | | |
| Post-Pulse x woman | | -0.08 | | | |
| Post-Pulse x Liberal | | (0.08) | -0.01 | | |
| | | | (0.08) | | |
| Post-Pulse x % LGBTQ (State) | | | (0.00) | -0.09 | |
| - 、 / | | | | (0.58) | |
| Post-Pulse x SS Couple Density | | | | | -0.19 |
| | | | | | (0.34) |
| Non-White | -0.24^{***} | | | | |
| Woman | (0.07) | 0.00 | 0.05 | 0.05 | 0.05 |
| woman | (0.03) | (0.09) | (0.03) | (0.03) | (0.03) |
| Liberal | 0.38*** | 0.38*** | 0.38*** | 0.38*** | 0.38*** |
| | (0.04) | (0.04) | (0.05) | (0.04) | (0.04) |
| % LGBT (State) | ~ / | · · · | | 0.31 | × / |
| | | | | (0.40) | |
| SS Couple Per Capita (County) | | | | | 0.22 |
| | | | | | (0.17) |
| \mathbb{R}^2 | 0.35 | 0.36 | 0.35 | 0.36 | 0.36 |
| Num. obs. | 1132 | 1132 | 1132 | 1132 | 1132 |
| N Clusters | | | | 50 | 585 |

B.13 Evaluating Individual-Level Heterogeneity

Table B15: Assessing Heterogeneous Influence of Post-Pulse (Study 1)

Note: ***p < 0.001; **p < 0.01; *p < 0.05. All models adjust for age, white (if not assessing heterogeneity by non-white), woman, religiosity, marital status, parental status, income, college education, unemployed status, union member, renter status, internet access, internet mode, liberal, metropolitan residence and Florida, Texas, California, New York, and Pennsylvania residence. HC2 robust SEs in parentheses but clustered at state and county-level for Models 4-5.

B.14 Evaluating Mechanisms

Table B16: Evaluating Different Mechanisms That Motivate SSM Support Post-
Pulse

| | SSM Support | | | | | | | |
|---------------------------------|----------------|-----------------|-----------------|-----------------|-----------------|----------------|-----------------|---|
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Post-Pulse x Black | 0.11 (0.12) | | | | | | | |
| Post-Pulse x Latino | () | -0.03 (0.11) | | | | | | |
| Post-Pulse x Woman | | | -0.09 (0.08) | | | | | |
| Post-Pulse x % LGBTQ (State) | | | | -0.02 (0.77) | | | | |
| Post-Pulse x SS Couple Density | | | | | -0.20 (0.34) | | | |
| Post-Pulse x Political Interest | | | | | . , | 0.03 (0.08) | | |
| Post-Pulse x News Freq. | | | | | | . , | -0.02 (0.08) | |
| Post-Pulse x Liberal | | | | | | | () | $\begin{array}{c} 0.01 \\ (0.08) \end{array}$ |
| \mathbb{R}^2 | 0.37 | 0.36 | 0.36 | 0.36 | 0.36 | 0.36 | 0.36 | 0.35 |
| Ν | 1132 | 1132 | 1132 | 1132 | 1132 | 1132 | 1132 | 1132 |

Note: ***p < 0.001, *p < 0.01, *p < 0.05. All models are fully specified, this table only presents the interaction between the *post-pulse* indicator and mechanisms that may explain the adoption of SSM support. HC2 robust standard errors in parentheses. SEs in Models 4, 5 are clustered at the state and county-level respectively. All variables are scaled between 0-1. All estimates are population-weighted.

B.15 Heterogeneity By Political Interest and Media Consumption

B.15.1 Measurement of Political Interest and Media Consumption

Political Interest: How interested would you say you are in politics and current affairs? 1) very interested, 2) somewhat interested, 3) not very interested, 4) not at all interested. Coded as a binary indicator equal to 1 if respondent puts "very interested," 0 otherwise (45% say "very interested," 55% say otherwise).

News Consumption: How frequently do you pay attention to news about national and international issues? 1) every day, 2) several times a week, 3) once a week, 4) several times a month, 5) once a month, 6) less often, 7) never. Coded as a binary indicator equal to 1 if respondent puts "every day," 0 otherwise (59% say "every day", 41% say otherwise).

Interest Scale: The *interest scale* is an additive index from 0-2 of the news consumption and *political interest* measures discussed above (0 = 32% of the sample, 1 = 32% of the sample, 2 = 36% of the sample).

B.15.2 Results

| | | | SSM S | upport | | |
|---------------------------------|------------------|------------------|-----------------|-----------------|-----------------|---|
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Post-Pulse x Political Interest | $0.00 \\ (0.09)$ | $0.03 \\ (0.08)$ | | | | |
| Post-Pulse x News Consumption | | | -0.11 (0.09) | -0.02 (0.08) | | |
| Post-Pulse x Interest Scale | | | | | -0.04 (0.06) | $\begin{array}{c} 0.00 \\ (0.05) \end{array}$ |
| Controls | Ν | Y | Ν | Υ | Ν | Υ |
| \mathbb{R}^2 | 0.02 | 0.36 | 0.03 | 0.36 | 0.02 | 0.36 |
| Ν | 1134 | 1132 | 1134 | 1132 | 1134 | 1132 |

Table B17: Evaluating Heterogenous Influence of Post-Pulse Conditional onPolitical Interest and News Consumption

Note: ***p < 0.001, **p < 0.01, *p < 0.05. Models alternate between excluding/including control covariates. This table only presents the interaction between the *post-pulse* indicator and *political interest*, *news consumption*, and the *interest scale*. HC2 robust standard errors in parentheses. All covariates are scaled between 0-1. All estimates are population-weighted.

Here we assess the heterogenous influence of being interviewed *post-Pulse* on *SSM support* among TAPS respondents conditional on political interest and news consumption. We conduct this test to assess if those who are attuned to media and politics are differentially more likely to support the rights of segments of the LGBTQ+ community in response to exposure to violence against LGBTQ+ (Reny and Newman, 2021).

Table B17 demonstrates that the influence of being interviewed *post-Pulse* on *SSM support* is not heterogeneous with respect to *political interest*, *news consumption* levels, or the *interest scale*.

We do not think the absence of heterogeneity poses a problem for the validity of our results. Consistent with prior research, the *political interest* and *news consumption* measures capture a *general* disposition towards consuming media and politics that is relatively stable (Prior, 2010). But that general disposition may be abrogated in the context of high-salience events. This is to say, even those segments of the mass public who do not necessarily pay attention to salient political/media events may have internalized information about the Pulse massacre. This is corroborated by our evidence on Figure A1, which demonstrates that 86% (Kaiser Poll, June 15-21, 2016) to 89% (CBS News Poll, June 13-14, 2016) of the mass public was closely following the shooting. Therefore, nearly all of the mass public was closely following the shooting. Therefore, nearly all of the mass public was closely following the pulse massacre, implying high levels of potential treatment reception regardless of one's generalized *political interest* or level of *news consumption*.

B.16 Balance Tests After Removing Days After Pulse Event

| Days Cut | # Imbalanced Covariates (out of 20) | Imbalanced Covariates |
|----------|-------------------------------------|--|
| 1 | 2/20 | Age, College |
| 2 | 1/20 | Age |
| 3 | 2/20 | Age, Child |
| 4 | 1/20 | Age |
| 5 | 2/20 | Age, Child |
| 6 | 1/20 | Age |
| 7 | 2/20 | Age, College |
| 8 | 2/20 | Age, College |
| 9 | 3/20 | Age, Married, College |
| 10 | 3/20 | Age, Child, College |
| 11 | 3/20 | Age, Child, College |
| 12 | 2/20 | Age, College |
| 13 | 3/20 | Age, Child, College |
| 14 | 1/20 | Age |
| 15 | 2/20 | Age, Child |
| 16 | 1/20 | Age |
| 17 | 2/20 | Age, Metro |
| 18 | 3/20 | Age, Married, Metro |
| 19 | 2/20 | Age, Married |
| 20 | 2/20 | Union, Internet Access |
| 21 | 1/20 | Internet Access |
| 22 | 1/20 | Internet Access |
| 23 | 3/20 | Married, Internet Access, Florida |
| 24 | 5/20 | Non-religious, College, Internet Access, Florida, California |

 Table B18: Covariate Balance Tests After Cutting Days Immediately After Pulse

 Massacre

B.17 Heterogeneity by Conservatism

B.17.1 Full Sample

Table B19: Heterogenous Effect of $Post\mathchar`Pulse$ on SSM Support Conditional on Conservatism

| | SSM Support | | |
|---------------------------|---------------|------------------|--|
| | (1) | (2) | |
| Post-Pulse x Conservative | 0.01 | 0.00 | |
| | (0.09) | (0.08) | |
| Post-Pulse | 0.10 | 0.10^{\dagger} | |
| | (0.06) | (0.06) | |
| Age | | 0.02 | |
| | | (0.07) | |
| White | | 0.20^{***} | |
| | | (0.05) | |
| Woman | | 0.04 | |
| | | (0.04) | |
| Non-religious | | 0.24^{***} | |
| | | (0.06) | |
| Married | | -0.02 | |
| | | (0.05) | |
| Child | | -0.09' | |
| | | (0.05) | |
| Income | | 0.06 | |
| Callana | | (0.07) | |
| College | | (0.04) | |
| Unomployed | | (0.04) | |
| Unemployed | | -0.03 | |
| Union | | (0.05) | |
| emon | | (0.05) | |
| Benter | | 0.01 | |
| | | (0.06) | |
| Internet Access | | -0.03 | |
| | | (0.05) | |
| Internet Mode | | 0.02 | |
| | | (0.04) | |
| Conservative | -0.39^{***} | -0.15^{*} | |
| | (0.05) | (0.06) | |
| Liberal | . , | 0.28*** | |
| | | (0.06) | |
| Metro Area | | 0.05 | |
| | | (0.05) | |
| \mathbb{R}^2 | 0.16 | 0.37 | |
| Ν | 1134 | 1132 | |

*** $p < 0.001; ** p < 0.01; * p < 0.05; \dagger p < 0.1$

B.17.2 After Cutting Days Immediately Post-Pulse

| Table B20: | Heterogenous | Effects of | f Post-Pulse | Conditional | on | Conservatism |
|--------------|---------------|------------|--------------|-------------|----|--------------|
| For Differen | t Samples Whe | ere Days I | Post-Pulse A | re Cut | | |

| Variable | Coefficient | SE | p-value | Days Cut | N | Controls |
|---------------------------|-------------|------|---------|----------|-----|----------|
| Post-Pulse | 0.02 | 0.08 | 0.85 | 10 | 817 | Yes |
| Post-Pulse x Conservative | 0.10 | 0.13 | 0.44 | 10 | 817 | Yes |
| Post-Pulse | -0.03 | 0.08 | 0.69 | 11 | 798 | Yes |
| Post-Pulse x Conservative | 0.04 | 0.12 | 0.75 | 11 | 798 | Yes |
| Post-Pulse | -0.05 | 0.08 | 0.53 | 12 | 790 | Yes |
| Post-Pulse x Conservative | 0.08 | 0.12 | 0.51 | 12 | 790 | Yes |
| Post-Pulse | -0.07 | 0.08 | 0.39 | 13 | 786 | Yes |
| Post-Pulse x Conservative | 0.09 | 0.12 | 0.46 | 13 | 786 | Yes |
| Post-Pulse | -0.08 | 0.09 | 0.38 | 14 | 779 | Yes |
| Post-Pulse x Conservative | 0.09 | 0.12 | 0.46 | 14 | 779 | Yes |
| Post-Pulse | -0.09 | 0.09 | 0.33 | 15 | 774 | Yes |
| Post-Pulse x Conservative | 0.11 | 0.13 | 0.39 | 15 | 774 | Yes |
| Post-Pulse | -0.11 | 0.09 | 0.23 | 16 | 761 | Yes |
| Post-Pulse x Conservative | 0.06 | 0.13 | 0.64 | 16 | 761 | Yes |
| Post-Pulse | -0.09 | 0.10 | 0.39 | 17 | 752 | Yes |
| Post-Pulse x Conservative | 0.03 | 0.14 | 0.81 | 17 | 752 | Yes |
| Post-Pulse | 0.04 | 0.07 | 0.59 | 18 | 743 | Yes |
| Post-Pulse x Conservative | -0.08 | 0.12 | 0.48 | 18 | 743 | Yes |
| Post-Pulse | 0.01 | 0.07 | 0.87 | 19 | 736 | Yes |
| Post-Pulse x Conservative | -0.03 | 0.12 | 0.79 | 19 | 736 | Yes |
| Post-Pulse | 0.00 | 0.07 | 0.95 | 20 | 724 | Yes |
| Post-Pulse x Conservative | -0.03 | 0.13 | 0.82 | 20 | 724 | Yes |
| Post-Pulse | 0.08 | 0.08 | 0.32 | 21 | 716 | Yes |
| Post-Pulse x Conservative | -0.09 | 0.14 | 0.55 | 21 | 716 | Yes |
| Post-Pulse | 0.10 | 0.09 | 0.23 | 22 | 707 | Yes |
| Post-Pulse x Conservative | -0.09 | 0.16 | 0.59 | 22 | 707 | Yes |
| Post-Pulse | 0.11 | 0.09 | 0.21 | 23 | 700 | Yes |
| Post-Pulse x Conservative | -0.20 | 0.13 | 0.13 | 23 | 700 | Yes |
| Post-Pulse | 0.15 | 0.17 | 0.38 | 24 | 687 | Yes |
| Post-Pulse x Conservative | -0.28 | 0.20 | 0.16 | 24 | 687 | Yes |
B.18 Heterogeneity by Gun Control Opposition

B.18.1 Full Sample

| | (1) | (2) |
|---------------------------------|---------------|-------------------|
| Post-Pulse x Oppose Gun Control | -0.04 | -0.07 |
| | (0.09) | (0.08) |
| Post-Pulse | 0.12^{*} | 0.12^{*} |
| | (0.06) | (0.06) |
| Oppose Gun Control | -0.30^{***} | -0.08 |
| | (0.06) | (0.05) |
| Age | | 0.02 |
| | | (0.07) |
| White | | 0.21^{***} |
| | | (0.05) |
| Woman | | 0.03 |
| | | (0.04) |
| Non-religious | | 0.24^{***} |
| | | (0.06) |
| Married | | -0.01 |
| | | (0.05) |
| Child | | -0.09^{\dagger} |
| _ | | (0.05) |
| Income | | 0.04 |
| | | (0.07) |
| College | | 0.10^{*} |
| | | (0.04) |
| Unemployed | | -0.03 |
| TT • | | (0.05) |
| Union | | -0.04 |
| D (| | (0.05) |
| Renter | | (0.02) |
| Internet Access | | (0.06) |
| Internet Access | | -0.03 |
| Internet Mede | | (0.04) |
| Internet Mode | | (0.01) |
| Liboral | | (0.04) |
| Liberai | | (0.05) |
| Metro Area | | 0.05 |
| meno mea | | (0.05) |
| - 2 | | (0.00) |
| R ² | 0.12 | 0.36 |
| Num. obs. | 1134 | 1132 |

| Table B21: | Heterogenous | Effect of | Post-Pulse | on | SSM | Support | Conditional | on |
|------------|---------------|-----------|------------|----|-----|---------|-------------|----|
| Gun Contre | ol Opposition | | | | | | | |

***p < 0.001; **p < 0.01; *p < 0.05; †p < 0.1

B.18.2 After Cutting Days Immediately Post-Pulse

Table B22: Heterogenous Effects of Post-Pulse Conditional on Gun ControlOpposition For Different Samples Where Days Post-Pulse Are Cut

| Variable | Coefficient | \mathbf{SE} | p-value | Days Cut | Ν | Controls |
|---------------------------------|-------------|---------------|---------|----------|-----|----------|
| Post-Pulse | 0.05 | 0.09 | 0.55 | 10 | 817 | Yes |
| Post-Pulse x Oppose Gun Control | -0.02 | 0.12 | 0.83 | 10 | 817 | Yes |
| Post-Pulse | -0.04 | 0.09 | 0.66 | 11 | 798 | Yes |
| Post-Pulse x Oppose Gun Control | 0.05 | 0.12 | 0.67 | 11 | 798 | Yes |
| Post-Pulse | -0.06 | 0.09 | 0.49 | 12 | 790 | Yes |
| Post-Pulse x Oppose Gun Control | 0.09 | 0.12 | 0.45 | 12 | 790 | Yes |
| Post-Pulse | -0.09 | 0.09 | 0.35 | 13 | 786 | Yes |
| Post-Pulse x Oppose Gun Control | 0.11 | 0.12 | 0.36 | 13 | 786 | Yes |
| Post-Pulse | -0.09 | 0.09 | 0.35 | 14 | 779 | Yes |
| Post-Pulse x Oppose Gun Control | 0.11 | 0.13 | 0.39 | 14 | 779 | Yes |
| Post-Pulse | -0.10 | 0.10 | 0.30 | 15 | 774 | Yes |
| Post-Pulse x Oppose Gun Control | 0.13 | 0.13 | 0.32 | 15 | 774 | Yes |
| Post-Pulse | -0.13 | 0.10 | 0.21 | 16 | 761 | Yes |
| Post-Pulse x Oppose Gun Control | 0.12 | 0.13 | 0.37 | 16 | 761 | Yes |
| Post-Pulse | -0.10 | 0.12 | 0.40 | 17 | 752 | Yes |
| Post-Pulse x Oppose Gun Control | 0.06 | 0.15 | 0.67 | 17 | 752 | Yes |
| Post-Pulse | 0.03 | 0.07 | 0.72 | 18 | 743 | Yes |
| Post-Pulse x Oppose Gun Control | -0.02 | 0.11 | 0.88 | 18 | 743 | Yes |
| Post-Pulse | 0.01 | 0.07 | 0.90 | 19 | 736 | Yes |
| Post-Pulse x Oppose Gun Control | 0.01 | 0.11 | 0.94 | 19 | 736 | Yes |
| Post-Pulse | -0.04 | 0.08 | 0.63 | 20 | 724 | Yes |
| Post-Pulse x Oppose Gun Control | 0.09 | 0.12 | 0.44 | 20 | 724 | Yes |
| Post-Pulse | -0.01 | 0.07 | 0.93 | 21 | 716 | Yes |
| Post-Pulse x Oppose Gun Control | 0.13 | 0.15 | 0.39 | 21 | 716 | Yes |
| Post-Pulse | 0.00 | 0.08 | 0.98 | 22 | 707 | Yes |
| Post-Pulse x Oppose Gun Control | 0.16 | 0.16 | 0.31 | 22 | 707 | Yes |
| Post-Pulse | 0.05 | 0.08 | 0.55 | 23 | 700 | Yes |
| Post-Pulse x Oppose Gun Control | -0.02 | 0.17 | 0.92 | 23 | 700 | Yes |
| Post-Pulse | 0.07 | 0.11 | 0.50 | 24 | 687 | Yes |
| Post-Pulse x Oppose Gun Control | -0.05 | 0.23 | 0.83 | 24 | 687 | Yes |

B.19 Heterogeneity by Predicted SSM Support

Table B23: Assessing The Heterogenous Effect of Post-Pulse Conditional On Predicted Support For Same-Sex Marriage If Post-Pulse Indicator Is Equal To 0

| | SSM Support (1) |
|--------------------------------------|-----------------|
| Post-Pulse x SSM Support (Predicted) | -0.03 |
| De et Dellee | (0.11) |
| Post-Pulse | (0.06) |
| SSM Support (Predicted) | 1.01*** |
| | (0.06) |
| Controls? | Ν |
| R^2 | 0.36 |
| Ν | 1132 |

***p < 0.001; **p < 0.01; *p < 0.05; †p < 0.1

C Study 2: PI S-IAT

C.1 Representativeness Discussion

The PI data are not population representative. The sample contains more youth (68% aged 18-29 vs. 18%), women (65% vs. 51%), liberals (57% vs. 41%), college educated (44% vs. 29%), and non-whites (36% vs. 22%) than TAPS. However, although the PI sample is disproportionately composed of respondent attributes associated with pro-LGBTQ+ attitudes, the empirical conclusions we draw from the PI sample may translate to a representative population. Prior research demonstrates non-representative internet samples respond similarly to external stimuli as representative samples (Coppock, 2019). If Study 2 corroborates results from a nationally representative sample (Study 1), we may have confidence Study 2's findings are generalizable.

C.2 Baseline Covariate Measurement

Age: Self-reported age, rescaled between 0-1.

Woman: 1 if respondent indicates they are "female," 0 otherwise.

White: 1 if respondent indicates they are "white," 0 otherwise.

College: 1 if respondent indicates the highest level of education they have is a "bachelor's degree," "some graduate school," a "master's degree," a "J.D.," a "M.D.," a "PhD," an other "advanced degree" or a "M.B.A." 0 otherwise.

Liberal: 1 if respondent indicates their political identity is "slightly liberal," "moderately liberal," or "strongly liberal." 0 otherwise.

Religious: 1 if respondent indicates they are not "not at all religious," 0 otherwise

Non-Metro: 1 if respondent is not from a "nonmetropolitian area," 0 otherwise.

California/Pennsylvania/New York/Florida/Illinois: 1 if respondent indicates their state of residence is California/Pennsylvania/New York/Florida/Illinois, 0 otherwise.

C.3 D-Score Details

The S-IAT acquires respondents mean compatible response latency (CRL) and incompatible response latency (IRL) (in milliseconds). The CRL is an average of how quickly a respondent associates "good" (e.g. happy, terrific) and "bad" (e.g. evil, rotten) words in addition to "gay" (e.g. homosexual, woman/woman image) or "straight" (e.g. heterosexual, man/woman image) words/images to a left or right-sided bin that characterize associations designed to be easy for people who prefer straight to gay people (e.g. gay/bad, straight/good). The IRL measures the same thing but where the left or right-sided bins characterize associations designed to be difficult for people who prefer straight to gay people (e.g. gay/good, straight/bad). The S-IAT assumes implicitly biased respondents will be faster making congruent than incongruent associations. Consequently, the *D*-score is the IRL - CRL difference divided by the within-individual standard deviation of response latencies calculated across the compatible and incompatible trials. The *D*-score ranges from -2-2, with higher values suggesting implicit bias against gay people .

C.4 Anti-Gay Attitudes Over Time



Figure C33: Anti-Gay Attitudes (y-axis) Over Time (x-axis, in days) Between 2016-01-01 and 2016-09-07. Dashed vertical line is the moment the Pulse nightclub massacre occurred. Loess models are fit on each side of the moment Pulse occurred. All covariates re-scaled between 0-1.

C.5 Balance Tests



Figure C34: Balance on IAT Taker Composition Before and After the Massacre. Each panel characterizes covariate balance for different bandwidths (see plot title, with sample size). The x-axis is the *post-Pulse* coefficient derived from separate regression models regressing a baseline covariate (y-axis) on *post-Pulse*. Black coefficients are statistically significant, grey otherwise. See Section C.7.1 for regression tables characterizing these balance plots.

C.6 Explaining Coefficients = Meaningful

Figure 5 displays *post-Pulse* intent-to-treat coefficients where the outcome is the daily *D*-score, straight bias, and heterocentrism. Honing in on 15 and 20-day sample bandwidth estimates, respondents surveyed *post-Pulse* have a lower *D*-score (-0.01, p < 0.10) and heterocentrism (-0.01, p < 0.01), equivalent to roughly 7% and 8% of the respective outcome standard deviations pre-Pulse. Although these coefficients are small, they are reasonable and likely underestimated. First, prior research shows affective attitudes toward marginalized groups tend to be stable, so small attitudinal shifts may be meaningful (Sears, 1993; Vuletich and Payne, 2019).

Second, the *D*-score is indirectly measured, so it is less subject to impression management. Thus, small coefficients may be meaningful because the mass public may have difficulties shifting automatic attitudes toward LGBTQ+ community segments.

Third, conversely, *heterocentrism* is an explicit outcome asking respondents to indicate they favor straight to gay people. Thus, the measure may be subject to impression management where individuals who would otherwise adopt genuinely more prosocial beliefs toward LGBTQ+ group members may already be self-reporting disingenuous prosocial beliefs prior to the massacre on the basis of social desirability. These dynamics may generate ceiling effects on external stimuli that would otherwise motivate prosocial attitudes.

Fourth, coefficients may be smaller since we are estimating an intent-to-treat effect with a relatively youthful PI S-IAT sample relative to TAPS in Study 1. Youth pay less attention to media (Neundorf et al., 2013). Therefore, they may be less likely to shift their attitudes in response to media context changes, which could attenuate ITT effects. In summary, the "true" ITT effect may be much larger than what we identify in Study 2 if we had a representative adult population.

Fifth, we cannot truncate to attentive respondents in Study 2 like Study 1 due to the absence of auxiliary interview length data.⁵⁶ Inattentiveness may produce underestimates of the *post-Pulse* coefficient. Fifth, the coefficients are still meaningful from a relative basis. The *post-Pulse* coefficients for the *D-score* and *heterocentrism* outcomes are roughly 10% of the political liberalism coefficient, one of the most prognostic covariates determining prosocial attitudes toward LGBTQ+ people (Flores, 2014).

Sixth, Studies 1 and 3 suggest that violence against LGBTQ+ people can motivate relatively large ITT effects on prosocial attitudes toward LGBTQ+ group members (20% of the outcome standard deviation for Study 1, a 10 percentage point increase in support samesex marriage; 20% of the outcome standard deviation for Study 3, a 10 percentage point decrease in reporting homosexuality is immoral). Statistically, multiple testing of the same hypothesis will generate variation in effects (Gelman, 2015). This means that the small effect in Study 2 may be a function of statistical and/or sampling variation instead of the "true" effect if we had a survey of the entire US adult population. Therefore, on balance, we have two studies with relatively large effects, and one study with relatively small effects. In the aggregate, we believe these findings imply that violence against LGBTQ+ people can have a meaningful initial impact on mass attitudes. We also believe Study 2, even in identifying smaller effect sizes, is still important because it provides additional evidence congruent with

⁵⁶Study 3's coefficients, which are substantively larger, are from a telephone survey, where respondents are typically more attentive.

Studies 1 and 3.

Seventh, even if prosocial attitudinal shifts *post-Pulse* are small in Study 2, our target population is all US adult Americans, which could suggest substantively important effects. For instance, there are roughly 260 million adult Americans. If *heterocentrism* decreases by 0.01 on a scale from 0-1 (the *post-Pulse* effect size we identify), that could mean nearly 3 million Americans go from the maximum level of *heterocentrism* to the lowest level of *heterocentrism* while the other 257 million Americans do not shift their attitudes, a meaningful effect as far as the adult mass public is concerned. More reasonably, it could also mean *heterocentrism* decreases by one-fifth of the 0-1 scale for 15 million Americans while staying constant for the other 245 million Americans, again, a substantively meaningful effect given 15 million people are holding measurably lower levels of *heterocentrism*.

Eighth, we caution against demands for large effects in political science research. Often, large effect sizes are a function of limited statistical power, which could result in Type 1 errors. Small effect sizes are likely more reasonable, replicable, realistic, and externally valid in helping to explain human behavior (Arel-Bundock et al., 2022). Indeed, it is no surprise smaller effect sizes stem from Study 2 (N = 3674, 4956) relative to Studies 1 (N = 1132) and 3 (N = 2052) since Study 2 has a larger sample size.

C.7 Regression Tables

C.7.1 Balance Tests

| Outcome | Post-Pulse Coef. | \mathbf{SE} | р | Bandwidth | Ν |
|-----------|------------------|---------------|------|-----------|-------|
| Age | -0.01 | 0.01 | 0.42 | 5 days | 1501 |
| Age | -0.01 | 0.01 | 0.35 | 10 days | 2665 |
| Age | -0.00 | 0.00 | 0.46 | 15 days | 3674 |
| Age | -0.00 | 0.00 | 0.92 | 20 days | 4956 |
| Age | 0.00 | 0.00 | 0.56 | 25 days | 5991 |
| Age | 0.01 | 0.00 | 0.00 | 30 days | 7778 |
| Age | 0.01 | 0.00 | 0.00 | 35 days | 9419 |
| Age | 0.01 | 0.00 | 0.00 | 40 days | 10857 |
| Age | 0.02 | 0.00 | 0.00 | 45 days | 12198 |
| Age | 0.03 | 0.00 | 0.00 | 50 days | 14209 |
| Woman | 0.01 | 0.02 | 0.66 | 5 days | 1501 |
| Woman | 0.02 | 0.02 | 0.25 | 10 days | 2665 |
| Woman | 0.03 | 0.02 | 0.06 | 15 days | 3674 |
| Woman | 0.02 | 0.01 | 0.14 | 20 days | 4956 |
| Woman | 0.02 | 0.01 | 0.12 | 25 days | 5991 |
| Woman | 0.02 | 0.01 | 0.14 | 30 days | 7778 |
| Woman | -0.00 | 0.01 | 0.92 | 35 days | 9419 |
| Woman | 0.00 | 0.01 | 0.98 | 40 days | 10857 |
| Woman | -0.00 | 0.01 | 0.98 | 45 days | 12198 |
| Woman | -0.01 | 0.01 | 0.31 | 50 days | 14209 |
| White | -0.06 | 0.02 | 0.02 | 5 days | 1501 |
| White | -0.05 | 0.02 | 0.00 | 10 days | 2665 |
| White | -0.05 | 0.02 | 0.00 | 15 days | 3674 |
| White | -0.03 | 0.01 | 0.02 | 20 days | 4956 |
| White | -0.04 | 0.01 | 0.00 | 25 days | 5991 |
| White | -0.00 | 0.01 | 0.64 | 30 days | 7778 |
| White | -0.01 | 0.01 | 0.52 | 35 days | 9419 |
| White | -0.01 | 0.01 | 0.32 | 40 days | 10857 |
| White | -0.01 | 0.01 | 0.25 | 45 days | 12198 |
| Collore | 0.00 | 0.01 | 0.89 | 50 days | 14209 |
| College | -0.01 | 0.03 | 0.04 | 10 days | 2665 |
| College | -0.02 | 0.02 | 0.55 | 10 days | 2005 |
| College | -0.02 | 0.02 | 0.10 | 20 days | 4956 |
| College | 0.01 | 0.01 | 0.02 | 25 days | 5991 |
| College | 0.06 | 0.01 | 0.00 | 30 days | 7778 |
| College | 0.00 | 0.01 | 0.00 | 35 days | 9419 |
| College | 0.13 | 0.01 | 0.00 | 40 days | 10857 |
| College | 0.15 | 0.01 | 0.00 | 45 days | 12198 |
| College | 0.19 | 0.01 | 0.00 | 50 days | 14209 |
| Liberal | 0.02 | 0.03 | 0.43 | 5 days | 1501 |
| Liberal | 0.03 | 0.02 | 0.15 | 10 days | 2665 |
| Liberal | 0.02 | 0.02 | 0.20 | 15 days | 3674 |
| Liberal | 0.02 | 0.01 | 0.20 | 20 days | 4956 |
| Liberal | 0.03 | 0.01 | 0.03 | 25 days | 5991 |
| Liberal | 0.05 | 0.01 | 0.00 | 30 days | 7778 |
| Liberal | 0.07 | 0.01 | 0.00 | 35 days | 9419 |
| Liberal | 0.07 | 0.01 | 0.00 | 40 days | 10857 |
| Liberal | 0.07 | 0.01 | 0.00 | 45 days | 12198 |
| Liberal | 0.09 | 0.01 | 0.00 | 50 days | 14209 |
| Religious | 0.00 | 0.02 | 0.88 | 5 days | 1501 |
| Religious | -0.00 | 0.02 | 0.94 | 10 days | 2665 |
| Religious | 0.00 | 0.02 | 0.90 | 15 days | 3674 |
| Religious | -0.01 | 0.01 | 0.59 | 20 days | 4956 |
| Religious | -0.01 | 0.01 | 0.67 | 25 days | 5991 |
| Religious | -0.01 | 0.01 | 0.35 | 30 days | 7778 |
| Religious | -0.02 | 0.01 | 0.09 | 35 days | 9419 |
| Religious | -0.02 | 0.01 | 0.02 | 40 days | 10857 |
| Religious | -0.02 | 0.01 | 0.03 | 45 days | 12198 |
| Religious | -0.03 | 0.01 | 0.00 | 50 days | 14209 |

Table C24: Balance Tests (Part 1)

| Outcome | Post-Pulse Coef. | \mathbf{SE} | \mathbf{p} | Bandwidth | Ν |
|--------------|------------------|---------------|--------------|-----------|-------|
| Pennsylvania | 0.00 | 0.01 | 0.88 | 5 days | 1501 |
| Pennsylvania | 0.00 | 0.01 | 0.54 | 10 days | 2665 |
| Pennsylvania | 0.00 | 0.01 | 0.77 | 15 days | 3674 |
| Pennsylvania | 0.01 | 0.01 | 0.29 | 20 days | 4956 |
| Pennsylvania | 0.00 | 0.00 | 0.68 | 25 days | 5991 |
| Pennsylvania | -0.00 | 0.00 | 0.89 | 30 days | 7778 |
| Pennsylvania | -0.01 | 0.00 | 0.09 | 35 days | 9419 |
| Pennsylvania | -0.00 | 0.00 | 0.26 | 40 days | 10857 |
| Pennsylvania | -0.00 | 0.00 | 0.29 | 45 days | 12198 |
| Pennsylvania | -0.00 | 0.00 | 0.77 | 50 days | 14209 |
| New York | 0.01 | 0.01 | 0.40 | 5 days | 1501 |
| New York | 0.00 | 0.01 | 0.78 | 10 days | 2665 |
| New York | 0.00 | 0.01 | 0.86 | 15 days | 3674 |
| New York | -0.00 | 0.01 | 0.96 | 20 days | 4956 |
| New York | 0.00 | 0.01 | 0.80 | 25 days | 5991 |
| New York | -0.00 | 0.00 | 0.99 | 30 days | 7778 |
| New York | 0.00 | 0.00 | 0.29 | 35 days | 9419 |
| New York | 0.01 | 0.00 | 0.08 | 40 days | 10857 |
| New York | 0.01 | 0.00 | 0.20 | 45 days | 12198 |
| New York | 0.01 | 0.00 | 0.14 | 50 days | 14209 |
| Florida | 0.03 | 0.01 | 0.01 | 5 days | 1501 |
| Florida | 0.03 | 0.01 | 0.00 | 10 days | 2665 |
| Florida | 0.01 | 0.01 | 0.19 | 15 days | 3674 |
| Florida | 0.01 | 0.01 | 0.39 | 20 days | 4956 |
| Florida | 0.01 | 0.01 | 0.05 | 25 days | 5991 |
| Florida | 0.01 | 0.00 | 0.10 | 30 days | 7778 |
| Florida | 0.01 | 0.00 | 0.01 | 35 days | 9419 |
| Florida | 0.01 | 0.00 | 0.01 | 40 days | 10857 |
| Florida | 0.01 | 0.00 | 0.00 | 45 days | 12198 |
| Florida | 0.01 | 0.00 | 0.00 | 50 days | 14209 |
| Illinois | -0.01 | 0.01 | 0.56 | 5 days | 1501 |
| Illinois | 0.00 | 0.01 | 0.95 | 10 days | 2665 |
| Illinois | 0.00 | 0.01 | 0.62 | 15 days | 3674 |
| Illinois | -0.00 | 0.01 | 0.88 | 20 days | 4956 |
| Illinois | -0.00 | 0.00 | 0.81 | 25 days | 5991 |
| Illinois | 0.00 | 0.00 | 0.61 | 30 days | 7778 |
| Illinois | 0.00 | 0.00 | 0.65 | 35 days | 9419 |
| Illinois | -0.01 | 0.00 | 0.08 | 40 days | 10857 |
| Illinois | -0.01 | 0.00 | 0.03 | 45 days | 12198 |
| Illinois | -0.01 | 0.00 | 0.01 | 50 days | 14209 |

Table C25: Balance Tests (Part 2)

C.7.2 Influence of Pulse on Anti-Gay Attitudes

| A. D-Score -0.003 0.007 0.710 5 1487 No A. D-Score -0.009 0.006 0.089 10 2639 No A. D-Score -0.008 0.004 0.035 20 4907 No A. D-Score -0.008 0.004 0.037 25 5925 No A. D-Score -0.011 0.003 0.000 35 9313 No A. D-Score -0.017 0.003 0.000 40 10735 No A. D-Score -0.018 0.003 0.000 41651 No A. D-Score -0.022 0.002 0.000 Full 41900 No A. D-Score -0.002 0.007 0.757 5 1487 Yes A. D-Score -0.007 0.004 0.055 15 3638 Yes A. D-Score -0.007 0.003 0.006 3033 Yes |
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| A. D-Score -0.009 0.006 0.089 10 2639 No A. D-Score -0.008 0.004 0.037 25 5925 No A. D-Score -0.011 0.003 0.001 30 7689 No A. D-Score -0.017 0.003 0.000 35 9313 No A. D-Score -0.017 0.003 0.000 40 10735 No A. D-Score -0.018 0.002 0.000 50 14051 No A. D-Score -0.022 0.002 0.000 50 14051 No A. D-Score -0.002 0.007 0.757 5 1487 Yes A. D-Score -0.009 0.004 0.055 15 3638 Yes A. D-Score -0.007 0.004 0.056 20 4907 Yes A. D-Score -0.007 0.003 0.098 25 5925 Yes A. D-Score -0.012 0.003 0.000 35 9313 Yes A. D-Score -0 |
| A. D-Score -0.009 0.005 0.051 15 3638 No A. D-Score -0.008 0.004 0.035 20 4907 No A. D-Score -0.011 0.003 0.001 30 7689 No A. D-Score -0.017 0.003 0.000 35 9313 No A. D-Score -0.017 0.003 0.000 40 10735 No A. D-Score -0.012 0.002 0.000 50 14051 No A. D-Score -0.020 0.007 7.75 5 1487 Yes A. D-Score -0.009 0.005 0.093 10 2639 Yes A. D-Score -0.007 0.004 0.055 15 3638 Yes A. D-Score -0.007 0.003 0.026 30 7689 Yes A. D-Score -0.007 0.003 0.026 30 7689 Yes A. D-Score -0.012 0.003 0.000 45 12057 Yes A. D-Score - |
| A. D-Score -0.008 0.004 0.035 20 4907 No A. D-Score -0.008 0.004 0.037 25 5925 No A. D-Score -0.017 0.003 0.000 35 9313 No A. D-Score -0.017 0.003 0.000 40 10735 No A. D-Score -0.018 0.003 0.000 45 12057 No A. D-Score -0.022 0.002 0.000 Full 41900 No A. D-Score -0.020 0.002 0.000 Full 41900 No A. D-Score -0.009 0.005 0.933 10 2639 Yes A. D-Score -0.007 0.004 0.055 15 3638 Yes A. D-Score -0.007 0.003 0.002 30 7689 Yes A. D-Score -0.012 0.003 0.006 35 9313 Yes A. D-Score -0.012 0.002 0.000 40 10735 Yes A. D-Score |
| A. D-Score -0.008 0.004 0.037 25 5925 No A. D-Score -0.011 0.003 0.000 35 9313 No A. D-Score -0.017 0.003 0.000 40 10735 No A. D-Score -0.018 0.000 40 10735 No A. D-Score -0.022 0.002 0.000 Full 41900 No A. D-Score -0.020 0.002 0.000 Full 41900 No A. D-Score -0.009 0.004 0.055 14051 No A. D-Score -0.009 0.004 0.055 1487 Yes A. D-Score -0.007 0.004 0.055 15 3638 Yes A. D-Score -0.007 0.003 0.002 30 7689 Yes A. D-Score -0.012 0.003 0.000 40 10735 Yes A. D-Score -0.012 0.003 0.000 50 14051 Yes A. D-Score -0.014 0.002 0.000< |
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| A. D-Score -0.022 0.002 0.000 Full 41901 No A. D-Score -0.020 0.002 0.000 Full 41900 No A. D-Score -0.009 0.005 0.933 10 2639 Yes A. D-Score -0.009 0.004 0.055 15 3638 Yes A. D-Score -0.007 0.004 0.056 20 4907 Yes A. D-Score -0.006 0.003 0.098 25 5925 Yes A. D-Score -0.012 0.003 0.000 350 7689 Yes A. D-Score -0.012 0.003 0.000 40 1735 Yes A. D-Score -0.012 0.002 0.000 451 12057 Yes A. D-Score -0.014 0.002 0.000 Full 41900 Yes B. Straight Bias 0.010 0.098 0.73 10 2254 No B. Straight Bias -0.011 0.002 |
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| D. Straight Bias -0.017 0.004 0.000 45 1015 100 B. Straight Bias -0.017 0.004 0.000 45 11827 No B. Straight Bias -0.022 0.003 0.000 50 13780 No B. Straight Bias -0.021 0.002 0.000 Full 42738 No B. Straight Bias -0.005 0.007 0.439 10 2584 Yes B. Straight Bias -0.005 0.007 0.439 10 2584 Yes B. Straight Bias -0.006 0.006 0.332 15 3562 Yes B. Straight Bias -0.005 0.005 0.310 20 4799 Yes B. Straight Bias -0.008 0.004 0.012 30 7511 Yes B. Straight Bias -0.012 0.003 0.001 35 9111 Yes B. Straight Bias -0.011 0.003 0.002 45 11827 Yes B. |
| D. Straight Bias -0.022 0.003 0.000 50 13780 No B. Straight Bias -0.022 0.003 0.000 50 13780 No B. Straight Bias -0.021 0.002 0.000 Full 42738 No B. Straight Bias -0.005 0.007 0.439 10 2584 Yes B. Straight Bias -0.005 0.006 0.332 15 3562 Yes B. Straight Bias -0.005 0.005 0.310 20 4799 Yes B. Straight Bias -0.008 0.005 0.107 25 5794 Yes B. Straight Bias -0.008 0.004 0.042 30 7511 Yes B. Straight Bias -0.012 0.004 0.001 35 9111 Yes B. Straight Bias -0.011 0.003 0.002 45 11827 Yes B. Straight Bias -0.012 0.003 0.000 50 13780 Yes B. Straigh |
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| B. Straight Bias -0.008 0.004 0.042 30 7511 Yes B. Straight Bias -0.012 0.004 0.001 35 9111 Yes B. Straight Bias -0.011 0.003 0.001 40 10519 Yes B. Straight Bias -0.010 0.003 0.002 45 11827 Yes B. Straight Bias -0.012 0.003 0.000 50 13780 Yes B. Straight Bias -0.012 0.002 0.000 Full 42738 Yes |
| B. Straight Bias -0.012 0.004 0.001 35 9111 Yes B. Straight Bias -0.011 0.003 0.001 40 10519 Yes B. Straight Bias -0.010 0.003 0.002 45 11827 Yes B. Straight Bias -0.012 0.003 0.000 50 13780 Yes B. Straight Bias -0.012 0.002 0.000 Full 42738 Yes |
| B. Straight Bias -0.011 0.003 0.001 40 10519 Yes B. Straight Bias -0.010 0.003 0.002 45 11827 Yes B. Straight Bias -0.012 0.003 0.000 50 13780 Yes B. Straight Bias -0.012 0.002 0.000 Full 42738 Yes |
| B. Straight Bias -0.010 0.003 0.002 45 11827 Yes B. Straight Bias -0.012 0.003 0.000 50 13780 Yes B. Straight Bias -0.012 0.002 0.000 Full 42738 Yes |
| B. Straight Bias -0.012 0.003 0.000 50 13780 Yes B. Straight Bias -0.012 0.002 0.000 Full 42738 Yes |
| D. Straight Dias -0.012 0.002 0.000 Full 42738 fes |
| C Heterocontrigm 0.011 0.007 0.125 5 1480 No |
| C. Heterocentrism -0.011 0.007 0.125 0 1489 No |
| C. Heterocentrism -0.013 0.003 0.003 10 2043 No |
| C Heterocentrism -0.014 0.004 0.002 19 3049 No |
| C. Heterocentrism -0.013 0.003 0.000 25 5946 No |
| C. Heterocentrism -0.016 0.003 0.000 30 7720 No |
| C. Heterocentrism -0.020 0.003 0.000 35 9342 No |
| C. Heterocentrism -0.020 0.002 0.000 40 10772 No |
| C. Heterocentrism -0.019 0.002 0.000 45 12106 No |
| C. Heterocentrism -0.022 0.002 0.000 50 14093 No |
| C. Heterocentrism -0.021 0.001 0.000 Full 43639 No |
| C. Heterocentrism -0.010 0.007 0.129 5 1489 Yes |
| C. Heterocentrism -0.013 0.005 0.008 10 2643 Yes |
| C. Heterocentrism -0.012 0.004 0.002 15 3645 Yes |
| C. Heterocentrism -0.011 0.003 0.001 20 4920 Yes |
| U. Heterocentrism -0.010 0.003 0.001 25 5946 Yes |
| C. Heterocentrism -0.012 0.000 30 7720 Yes |
| C Heterocentrism 0.014 0.002 0.000 40 10779 Vec |
| C. Heterocentrism -0.013 0.002 0.000 40 10772 Yes |
| C Heterocentrism -0.014 0.002 0.000 50 14003 Ves |
| C. Heterocentrism -0.013 0.001 0.000 Full 43639 Yes |

Table C26: Influence of Pulse on Anti-Gay Attitudes

C.7.3 Influence of Control Covariates on Anti-Gay Attitudes

Table C27: Influence of Control Covariates on Heterocentrism (Part 1)

| Control | Control Coef | SE | n | Bandwidth | Outcome | N |
|--------------|--------------|-------|-------|-----------|----------------|------|
| Liberal | -0.082 | 0.007 | 0.000 | 5 000 | Heterocentrism | 1489 |
| Age | 0.002 | 0.001 | 0.447 | 5.000 | Heterocentrism | 1489 |
| White | -0.010 | 0.007 | 0.170 | 5.000 | Heterocentrism | 1/80 |
| Woman | -0.010 | 0.007 | 0.583 | 5.000 | Heterocentrism | 1489 |
| College | -0.018 | 0.008 | 0.017 | 5.000 | Heterocentrism | 1489 |
| Beligious | 0.042 | 0.008 | 0.000 | 5.000 | Heterocentrism | 1489 |
| Nonmetro | 0.004 | 0.014 | 0.748 | 5.000 | Heterocentrism | 1489 |
| California | -0.000 | 0.011 | 0.969 | 5.000 | Heterocentrism | 1489 |
| Pennsylvania | 0.020 | 0.016 | 0.232 | 5.000 | Heterocentrism | 1489 |
| New York | 0.026 | 0.013 | 0.649 | 5.000 | Heterocentrism | 1489 |
| Florida | -0.007 | 0.014 | 0.617 | 5.000 | Heterocentrism | 1489 |
| Illinois | -0.022 | 0.021 | 0.289 | 5.000 | Heterocentrism | 1489 |
| Liberal | -0.088 | 0.005 | 0.000 | 10,000 | Heterocentrism | 2643 |
| Age | 0.000 | 0.018 | 0.010 | 10.000 | Heterocentrism | 2643 |
| White | -0.012 | 0.015 | 0.010 | 10.000 | Heterocentrism | 2643 |
| Woman | -0.012 | 0.005 | 0.021 | 10.000 | Heterocentrism | 2643 |
| College | -0.004 | 0.005 | 0.400 | 10.000 | Heterocentrism | 2643 |
| Beligious | -0.000 | 0.000 | 0.000 | 10.000 | Heterocentrism | 2643 |
| Nonmotro | 0.040 | 0.000 | 0.000 | 10.000 | Hotorocontrism | 2643 |
| California | 0.000 | 0.009 | 0.400 | 10.000 | Hotorocontrism | 2043 |
| Popperlyania | -0.005 | 0.008 | 0.125 | 10.000 | Hotorocontrism | 2043 |
| Now Vork | 0.018 | 0.013 | 0.117 | 10.000 | Hotorocontrism | 2043 |
| Florida | 0.003 | 0.010 | 0.415 | 10.000 | Hotorocontrism | 2043 |
| Illinois | -0.011 | 0.010 | 0.207 | 10.000 | Hotorocontrism | 2043 |
| Liboral | -0.010 | 0.013 | 0.285 | 15.000 | Hotorocontrism | 2645 |
| Ago | -0.084 | 0.004 | 0.000 | 15.000 | Hotorocontrism | 3645 |
| Mbito | 0.048 | 0.015 | 0.001 | 15.000 | Hotorocontrism | 3645 |
| Woman | -0.009 | 0.005 | 0.004 | 15.000 | Hotorocontrism | 3645 |
| Collogo | -0.005 | 0.004 | 0.221 | 15.000 | Hotorocontrism | 3645 |
| Deligious | -0.008 | 0.005 | 0.079 | 15.000 | Heterocentrism | 2645 |
| Nonmotro | 0.044 | 0.005 | 0.000 | 15.000 | Heterocentrism | 2645 |
| California | 0.007 | 0.007 | 0.555 | 15.000 | Heterocentrism | 2645 |
| Donnauluania | -0.002 | 0.000 | 0.708 | 15.000 | Heterocentrism | 2645 |
| Your Vonly | 0.015 | 0.010 | 0.202 | 15.000 | Heterocentrism | 2645 |
| Florida | 0.005 | 0.009 | 0.002 | 15.000 | Heterocentrism | 3045 |
| Florida | 0.001 | 0.008 | 0.927 | 15.000 | Heterocentrism | 2645 |
| Liboral | -0.013 | 0.012 | 0.209 | 20,000 | Heterocentrism | 4020 |
| Ago | -0.065 | 0.004 | 0.000 | 20.000 | Hotorocontrism | 4920 |
| Mbito | 0.045 | 0.013 | 0.001 | 20.000 | Hotorocontrism | 4920 |
| Woman | -0.009 | 0.004 | 0.022 | 20.000 | Hotorocontrism | 4920 |
| Colloro | -0.008 | 0.004 | 0.030 | 20.000 | Heterocentrism | 4920 |
| Deligious | -0.004 | 0.004 | 0.275 | 20.000 | Heterocentrism | 4920 |
| Nonmotro | 0.044 | 0.004 | 0.000 | 20.000 | Heterocentrism | 4920 |
| California | -0.002 | 0.000 | 0.702 | 20.000 | Heterocentrism | 4920 |
| Donnauluania | -0.009 | 0.005 | 0.065 | 20.000 | Heterocentrism | 4920 |
| Feinsyivania | 0.000 | 0.009 | 0.519 | 20.000 | Heterocentrism | 4920 |
| Flowide | -0.005 | 0.008 | 0.000 | 20.000 | Heterocentrism | 4920 |
| F IOFICIA | 0.004 | 0.007 | 0.024 | 20.000 | Heterocentrism | 4920 |
| Tilinois | -0.009 | 0.009 | 0.338 | 20.000 | Heterocentrism | 4920 |
| Liberai | -0.081 | 0.005 | 0.000 | 25.000 | Heterocentrism | 5940 |
| Age | 0.045 | 0.011 | 0.000 | 25.000 | Heterocentrism | 5940 |
| white | -0.012 | 0.004 | 0.000 | 25.000 | Heterocentrism | 5940 |
| woman | -0.008 | 0.003 | 0.021 | 25.000 | Heterocentrism | 5946 |
| College | -0.003 | 0.004 | 0.341 | 25.000 | Heterocentrism | 5946 |
| Rengious | 0.043 | 0.004 | 0.000 | 25.000 | Heterocentrism | 5946 |
| Nonmetro | 0.001 | 0.006 | 0.859 | 25.000 | Heterocentrism | 5946 |
| California | -0.010 | 0.005 | 0.029 | 25.000 | Heterocentrism | 5946 |
| Pennsylvania | 0.001 | 0.008 | 0.946 | 25.000 | neterocentrism | 5946 |
| New York | -0.011 | 0.007 | 0.138 | 25.000 | Heterocentrism | 5946 |
| Florida | 0.004 | 0.007 | 0.591 | 25.000 | Heterocentrism | 5946 |
| Illinois | -0.003 | 0.008 | 0.706 | 25.000 | neterocentrism | 5946 |
| Liberal | -0.081 | 0.003 | 0.000 | 30.000 | neterocentrism | 7720 |
| Age | 0.046 | 0.010 | 0.000 | 30.000 | Heterocentrism | 7720 |
| White | -0.016 | 0.003 | 0.000 | 30.000 | Heterocentrism | 7720 |
| Woman | -0.012 | 0.003 | 0.000 | 30.000 | Heterocentrism | 7720 |
| College | -0.003 | 0.003 | 0.273 | 30.000 | Heterocentrism | 7720 |
| Religious | 0.042 | 0.003 | 0.000 | 30.000 | Heterocentrism | 7720 |

Table C28: Influence of Control Covariates on Heterocentrism (Part 2)

| Control | Control Coof | CL | n | Dandwidth | Outcomo | N |
|------------------|--------------|-------|-------|-----------|----------------|-------|
| Deligious | 0.042 | 0.002 | p | 20.000 | Hotorocontriam | 7790 |
| Nengious | 0.042 | 0.005 | 0.000 | 30.000 | Heterocentrism | 7720 |
| Nonmetro | -0.001 | 0.005 | 0.864 | 30.000 | Heterocentrism | 7720 |
| California | -0.011 | 0.004 | 0.008 | 30.000 | Heterocentrism | 7720 |
| Pennsylvania | -0.004 | 0.007 | 0.596 | 30.000 | Heterocentrism | 7720 |
| New York | -0.014 | 0.006 | 0.026 | 30.000 | Heterocentrism | 7720 |
| Florida | 0.001 | 0.007 | 0.828 | 30.000 | Heterocentrism | 7720 |
| Illinois | -0.009 | 0.008 | 0.257 | 30.000 | Heterocentrism | 7720 |
| Liberal | -0.078 | 0.003 | 0.000 | 35.000 | Heterocentrism | 9342 |
| Age | 0.045 | 0.009 | 0.000 | 35.000 | Heterocentrism | 9342 |
| White | -0.016 | 0.003 | 0.000 | 35.000 | Heterocentrism | 9342 |
| Woman | -0.012 | 0.003 | 0.000 | 35.000 | Heterocentrism | 9342 |
| College | -0.003 | 0.003 | 0.213 | 35.000 | Heterocentrism | 9342 |
| Religious | 0.043 | 0.003 | 0.000 | 35.000 | Heterocentrism | 9342 |
| Nonmetro | 0.001 | 0.005 | 0.829 | 35.000 | Heterocentrism | 9342 |
| California | -0.012 | 0.004 | 0.002 | 35.000 | Heterocentrism | 9342 |
| Pennsylvania | -0.004 | 0.006 | 0.482 | 35.000 | Heterocentrism | 9342 |
| New York | -0.016 | 0.006 | 0.004 | 35.000 | Heterocentrism | 9342 |
| Florida | 0.001 | 0.006 | 0.823 | 35.000 | Heterocentrism | 9342 |
| Illinois | -0.010 | 0.007 | 0.139 | 35.000 | Heterocentrism | 9342 |
| Liberal | -0.077 | 0.002 | 0.000 | 40.000 | Heterocentrism | 10772 |
| Age | 0.045 | 0.008 | 0.000 | 40.000 | Heterocentrism | 10772 |
| White | -0.016 | 0.003 | 0.000 | 40.000 | Heterocentrism | 10772 |
| Woman | -0.013 | 0.002 | 0.000 | 40.000 | Heterocentrism | 10772 |
| College | -0.004 | 0.003 | 0.086 | 40.000 | Heterocentrism | 10772 |
| Religious | 0.043 | 0.003 | 0.000 | 40.000 | Heterocentrism | 10772 |
| Nonmetro | 0.006 | 0.004 | 0.199 | 40.000 | Heterocentrism | 10772 |
| California | -0.011 | 0.004 | 0.003 | 40.000 | Heterocentrism | 10772 |
| Pennsylvania | -0.005 | 0.006 | 0.421 | 40.000 | Heterocentrism | 10772 |
| New York | -0.016 | 0.005 | 0.001 | 40,000 | Heterocentrism | 10772 |
| Florida | 0.001 | 0.006 | 0.897 | 40.000 | Heterocentrism | 10772 |
| Illinois | -0.001 | 0.006 | 0.225 | 40.000 | Heterocentrism | 10772 |
| Liberal | -0.077 | 0.000 | 0.000 | 45.000 | Heterocentrism | 12106 |
| Age | 0.045 | 0.002 | 0.000 | 45.000 | Heterocentrism | 12100 |
| White | -0.016 | 0.000 | 0.000 | 45.000 | Heterocentrism | 12100 |
| Women | -0.010 | 0.002 | 0.000 | 45.000 | Heterocentrism | 12100 |
| Collogo | -0.013 | 0.002 | 0.000 | 45.000 | Heterocentrism | 12100 |
| Religious | -0.004 | 0.002 | 0.124 | 45.000 | Heterocentrism | 12100 |
| Nonmotro | 0.044 | 0.003 | 0.000 | 45.000 | Heterocentrism | 12100 |
| California | 0.005 | 0.004 | 0.204 | 45.000 | Heterocentrism | 12100 |
| Damornia Dama | -0.012 | 0.005 | 0.001 | 45.000 | Heterocentrism | 12100 |
| Pennsylvania | -0.004 | 0.000 | 0.519 | 45.000 | Heterocentrism | 12100 |
| New York | -0.018 | 0.005 | 0.000 | 45.000 | Heterocentrism | 12100 |
| Florida | 0.004 | 0.006 | 0.513 | 45.000 | Heterocentrism | 12100 |
| Tilinois | -0.010 | 0.006 | 0.102 | 45.000 | Heterocentrism | 12100 |
| Liberal | -0.078 | 0.002 | 0.000 | 50.000 | Heterocentrism | 14093 |
| Age | 0.047 | 0.007 | 0.000 | 50.000 | Heterocentrism | 14093 |
| White | -0.017 | 0.002 | 0.000 | 50.000 | Heterocentrism | 14093 |
| Woman | -0.013 | 0.002 | 0.000 | 50.000 | Heterocentrism | 14093 |
| College | -0.005 | 0.002 | 0.044 | 50.000 | Heterocentrism | 14093 |
| Religious | 0.043 | 0.002 | 0.000 | 50.000 | Heterocentrism | 14093 |
| Nonmetro | 0.004 | 0.004 | 0.257 | 50.000 | Heterocentrism | 14093 |
| California | -0.011 | 0.003 | 0.001 | 50.000 | Heterocentrism | 14093 |
| Pennsylvania | -0.002 | 0.006 | 0.769 | 50.000 | Heterocentrism | 14093 |
| New York | -0.017 | 0.004 | 0.000 | 50.000 | Heterocentrism | 14093 |
| Florida | 0.002 | 0.005 | 0.654 | 50.000 | Heterocentrism | 14093 |
| Illinois | -0.008 | 0.005 | 0.166 | 50.000 | Heterocentrism | 14093 |
| Liberal | -0.080 | 0.001 | 0.000 | 200.000 | Heterocentrism | 43639 |
| Age | 0.030 | 0.005 | 0.000 | 200.000 | Heterocentrism | 43639 |
| White | -0.012 | 0.001 | 0.000 | 200.000 | Heterocentrism | 43639 |
| Woman | -0.015 | 0.001 | 0.000 | 200.000 | Heterocentrism | 43639 |
| College | -0.005 | 0.001 | 0.000 | 200.000 | Heterocentrism | 43639 |
| Religious | 0.047 | 0.001 | 0.000 | 200.000 | Heterocentrism | 43639 |
| Nonmetro | 0.012 | 0.002 | 0.000 | 200.000 | Heterocentrism | 43639 |
| California | -0.012 | 0.002 | 0.000 | 200.000 | Heterocentrism | 43639 |
| Pennsylvania | -0.001 | 0.003 | 0.714 | 200.000 | Heterocentrism | 43639 |
| New York | -0.012 | 0.002 | 0.000 | 200.000 | Heterocentrism | 43639 |
| Florida | -0.002 | 0.003 | 0.592 | 200.000 | Heterocentrism | 43639 |
| Illinois | -0.002 | 0.003 | 0.553 | 200.000 | Heterocentrism | 43639 |

Table C29: Influence of Control Covariates on Straight Bias (Part 1)

| Control | Control Coof | SF | n | Bandwidth | Outcomo | N |
|---------------------------|--------------|-------|------------|-----------|---------------|--------------|
| Liberal | 0.106 | 0.011 | 0.000 P | 5 000 | Straight Biog | 1453 |
| Age | -0.100 | 0.011 | 0.000 | 5.000 | Straight Bias | 1453 |
| White | -0.011 | 0.033 | 0.229 | 5.000 | Straight Bias | 1453 |
| Women | -0.011 | 0.011 | 0.233 | 5.000 | Straight Bias | 1453 |
| College | -0.027 | 0.011 | 0.107 | 5.000 | Straight Bias | 1453 |
| Beligious | 0.072 | 0.011 | 0.101 | 5.000 | Straight Bias | 1453 |
| Nonmetro | 0.072 | 0.012 | 0.000 | 5.000 | Straight Bias | 1453 |
| California | 0.025 | 0.022 | 0.043 | 5.000 | Straight Bias | 1453 |
| Popperlyania | 0.001 | 0.010 | 0.806 | 5.000 | Straight Bias | 1453 |
| New Vork | -0.003 | 0.025 | 0.558 | 5.000 | Straight Bias | 1453 |
| Florida | -0.010 | 0.017 | 0.338 | 5.000 | Straight Bias | 1453 |
| Illinois | -0.020 | 0.021 | 0.944 | 5.000 | Straight Bias | 1453 |
| Liboral | -0.020 | 0.023 | 0.201 | 10,000 | Straight Bias | 2584 |
| Age | 0.074 | 0.000 | 0.000 | 10.000 | Straight Bias | 2584 |
| White | 0.000 | 0.020 | 0.000 | 10.000 | Straight Bias | 2584 |
| Women | -0.009 | 0.008 | 0.252 | 10.000 | Straight Bias | 2584 |
| Collogo | -0.050 | 0.008 | 0.000 | 10.000 | Straight Bias | 2584 |
| Religious | -0.000 | 0.008 | 0.401 | 10.000 | Straight Bias | 2584 |
| Nonmotro | 0.015 | 0.009 | 0.000 | 10.000 | Straight Bias | 2584 |
| California | 0.013 | 0.014 | 0.200 | 10.000 | Straight Dias | 2504 |
| Donnauluania | 0.002 | 0.011 | 0.855 | 10.000 | Straight Dias | 2004 |
| Feinisyivania New Yeah | -0.005 | 0.019 | 0.794 | 10.000 | Straight Dias | 2004 |
| Flowido | -0.012 | 0.015 | 0.330 | 10.000 | Straight Dias | 2004 |
| Florida | -0.031 | 0.010 | 0.040 | 10.000 | Straight Dias | 2004 |
| Tillions | -0.005 | 0.019 | 0.802 | 15,000 | Straight Dias | 2004 |
| Liberai | -0.105 | 0.000 | 0.000 | 15.000 | Straight Dias | 3502 |
| Age | 0.000 | 0.025 | 0.007 | 15.000 | Straight Dias | 3502 |
| White | -0.011 | 0.007 | 0.097 | 15.000 | Straight Dias | 3502 |
| Collogo | -0.034 | 0.007 | 0.000 | 15.000 | Straight Dias | 3302 2569 |
| Dolligious | -0.004 | 0.007 | 0.000 | 15.000 | Straight Dias | 3302 2569 |
| Neuroretere | 0.075 | 0.007 | 0.000 | 15.000 | Straight Dias | 2562 |
| California | 0.017 | 0.011 | 0.152 | 15.000 | Straight Dias | 3502 |
| Damornia Damardanaia | 0.000 | 0.009 | 0.981 | 15.000 | Straight Dias | 3502 |
| New Verla | -0.000 | 0.015 | 0.095 | 15.000 | Straight Dias | 3502 |
| Flow fork | -0.014 | 0.012 | 0.223 | 15.000 | Straight Dias | 3502 |
| F IOI Ida | -0.010 | 0.015 | 0.450 | 15.000 | Straight Dias | 2562 |
| Tilinois | -0.010 | 0.017 | 0.009 | 15.000 | Straight Dias | 3302 |
| Ago | -0.108 | 0.000 | 0.000 | 20.000 | Straight Dias | 4799 |
| Age | 0.052 | 0.020 | 0.008 | 20.000 | Straight Dias | 4799 |
| Waman | -0.011 | 0.000 | 0.000 | 20.000 | Straight Dias | 4799 |
| Callana | -0.055 | 0.000 | 0.000 | 20.000 | Straight Dias | 4799 |
| D ali ai ana | 0.001 | 0.000 | 0.804 | 20.000 | Straight Dias | 4799 |
| Neuroretee | 0.074 | 0.000 | 0.000 | 20.000 | Straight Dias | 4799 |
| California | 0.009 | 0.010 | 0.332 | 20.000 | Straight Dias | 4799 |
| Damornia Damardanaia | -0.001 | 0.008 | 0.949 | 20.000 | Straight Dias | 4799 |
| New Verla | -0.012 | 0.013 | 0.049 | 20.000 | Straight Dias | 4799 |
| New York | -0.022 | 0.011 | 0.042 | 20.000 | Straight Dias | 4799 |
| Florida | 0.004 | 0.011 | 0.701 | 20.000 | Straight Blas | 4799 |
| 1111nois | -0.014 | 0.015 | 0.324 | 20.000 | Straight Blas | 4799 |
| Liberal | -0.111 | 0.005 | 0.000 | 25.000 | Straight Blas | 5794 |
| Age | 0.048 | 0.018 | 0.007 | 25.000 | Straight Dias | 5794 |
| white | -0.014 | 0.005 | 0.007 | 25.000 | Straight Dias | 5794 |
| Woman | -0.031 | 0.005 | 0.000 | 25.000 | Straight Bias | 5794 |
| College | 0.002 | 0.005 | 0.666 | 25.000 | Straight Bias | 5794 |
| Religious | 0.076 | 0.006 | 0.000 | 25.000 | Straight Blas | 5794 |
| Nonmetro | 0.008 | 0.009 | 0.370 | 25.000 | Straight Bias | 5794 |
| California | -0.003 | 0.007 | 0.641 | 25.000 | Straight Bias | 5794 |
| Pennsylvania | -0.002 | 0.013 | 0.878 | 25.000 | Straight Bias | 5794 |
| New York | -0.024 | 0.010 | 0.019 | 25.000 | Straight Bias | 5794 |
| r lorida | 0.002 | 0.010 | 0.828 | 25.000 | Straight Bias | 5794 |
| Illinois | -0.003 | 0.013 | 0.831 | 25.000 | Straight Bias | 5794 |
| Liberal | -0.111 | 0.004 | 0.000 | 30.000 | Straight Bias | 7511 |
| Age | 0.052 | 0.015 | 0.001 | 30.000 | Straight Bias | 7511 |
| White | -0.016 | 0.005 | 0.001 | 30.000 | Straight Bias | 7511 |
| Woman | -0.037 | 0.005 | 0.000 | 30.000 | Straight Bias | 7511 |
| College | 0.003 | 0.005 | 0.477 | 30.000 | Straight Bias | 7511 |
| Religious | 0.072 | 0.005 | 0.000 | 30.000 | Straight Bias | 7511 |

Table C30: Influence of Control Covariates on Straight Bias (Part 2)

| Control | Control Coef. | SE | р | Bandwidth | Outcome | Ν |
|--------------|---------------|-------|-------|-----------|---------------|-------|
| Religious | 0.072 | 0.005 | 0.000 | 30.000 | Straight Bias | 7511 |
| Nonmetro | 0.004 | 0.008 | 0.605 | 30.000 | Straight Bias | 7511 |
| California | -0.004 | 0.006 | 0.532 | 30.000 | Straight Bias | 7511 |
| Pennsylvania | -0.005 | 0.011 | 0.658 | 30.000 | Straight Bias | 7511 |
| New York | -0.026 | 0.009 | 0.004 | 30.000 | Straight Bias | 7511 |
| Florida | 0.002 | 0.009 | 0.803 | 30.000 | Straight Bias | 7511 |
| Illinois | -0.003 | 0.012 | 0.812 | 30.000 | Straight Bias | 7511 |
| Liberal | -0.109 | 0.004 | 0.000 | 35.000 | Straight Bias | 9111 |
| Age | 0.051 | 0.014 | 0.000 | 35.000 | Straight Bias | 9111 |
| White | -0.016 | 0.004 | 0.000 | 35.000 | Straight Bias | 9111 |
| Woman | -0.037 | 0.004 | 0.000 | 35.000 | Straight Bias | 9111 |
| College | 0.002 | 0.004 | 0.643 | 35.000 | Straight Bias | 9111 |
| Religious | 0.073 | 0.004 | 0.000 | 35.000 | Straight Bias | 9111 |
| Nonmetro | 0.008 | 0.007 | 0.238 | 35.000 | Straight Bias | 9111 |
| California | -0.003 | 0.006 | 0.613 | 35.000 | Straight Bias | 9111 |
| Pennsylvania | -0.005 | 0.010 | 0.642 | 35.000 | Straight Bias | 9111 |
| New York | -0.029 | 0.008 | 0.001 | 35.000 | Straight Bias | 9111 |
| Florida | 0.002 | 0.009 | 0.805 | 35.000 | Straight Bias | 9111 |
| Illinois | -0.004 | 0.011 | 0.699 | 35.000 | Straight Bias | 9111 |
| Liberal | -0.109 | 0.004 | 0.000 | 40.000 | Straight Bias | 10519 |
| Age | 0.057 | 0.013 | 0.000 | 40.000 | Straight Bias | 10519 |
| White | -0.017 | 0.004 | 0.000 | 40.000 | Straight Bias | 10519 |
| Woman | -0.037 | 0.004 | 0.000 | 40.000 | Straight Bias | 10519 |
| College | 0.002 | 0.004 | 0.681 | 40.000 | Straight Bias | 10519 |
| Religious | 0.074 | 0.004 | 0.000 | 40.000 | Straight Bias | 10519 |
| Nonmetro | 0.014 | 0.006 | 0.031 | 40.000 | Straight Bias | 10519 |
| California | -0.003 | 0.005 | 0.631 | 40.000 | Straight Bias | 10519 |
| Pennsylvania | -0.003 | 0.009 | 0.704 | 40.000 | Straight Bias | 10519 |
| New York | -0.028 | 0.008 | 0.000 | 40.000 | Straight Bias | 10519 |
| Florida | 0.002 | 0.008 | 0.817 | 40.000 | Straight Bias | 10519 |
| Illinois | -0.002 | 0.010 | 0.875 | 40.000 | Straight Bias | 10519 |
| Liberal | -0.111 | 0.004 | 0.000 | 45.000 | Straight Bias | 11827 |
| Age | 0.057 | 0.012 | 0.000 | 45.000 | Straight Bias | 11827 |
| White | -0.017 | 0.004 | 0.000 | 45.000 | Straight Bias | 11827 |
| Woman | -0.038 | 0.004 | 0.000 | 45.000 | Straight Bias | 11827 |
| College | 0.004 | 0.004 | 0.332 | 45.000 | Straight Bias | 11827 |
| Religious | 0.074 | 0.004 | 0.000 | 45.000 | Straight Bias | 11827 |
| Nonmetro | 0.012 | 0.006 | 0.041 | 45.000 | Straight Bias | 11827 |
| California | -0.007 | 0.005 | 0.191 | 45.000 | Straight Bias | 11827 |
| Pennsylvania | -0.005 | 0.009 | 0.594 | 45.000 | Straight Bias | 11827 |
| New York | -0.032 | 0.007 | 0.000 | 45.000 | Straight Bias | 11827 |
| Florida | 0.005 | 0.008 | 0.549 | 45.000 | Straight Bias | 11827 |
| Illinois | -0.009 | 0.009 | 0.315 | 45.000 | Straight Bias | 11827 |
| Liberal | -0.115 | 0.003 | 0.000 | 50.000 | Straight Bias | 13780 |
| Age | 0.064 | 0.011 | 0.000 | 50.000 | Straight Bias | 13780 |
| White | -0.018 | 0.003 | 0.000 | 50.000 | Straight Bias | 13780 |
| Woman | -0.039 | 0.003 | 0.000 | 50.000 | Straight Bias | 13780 |
| College | 0.001 | 0.003 | 0.838 | 50.000 | Straight Bias | 13780 |
| Religious | 0.070 | 0.004 | 0.000 | 50.000 | Straight Bias | 13780 |
| Nonmetro | 0.010 | 0.006 | 0.081 | 50.000 | Straight Bias | 13780 |
| California | -0.007 | 0.005 | 0.129 | 50.000 | Straight Bias | 13780 |
| Pennsylvania | -0.002 | 0.008 | 0.837 | 50.000 | Straight Bias | 13780 |
| New York | -0.031 | 0.007 | 0.000 | 50.000 | Straight Bias | 13780 |
| Florida | 0.003 | 0.008 | 0.747 | 50.000 | Straight Bias | 13780 |
| Illinois | -0.004 | 0.009 | 0.620 | 50.000 | Straight Bias | 13780 |
| Liberal | -0.113 | 0.002 | 0.000 | 200.000 | Straight Bias | 42738 |
| Age | 0.056 | 0.007 | 0.000 | 200.000 | Straight Bias | 42738 |
| White | -0.013 | 0.002 | 0.000 | 200.000 | Straight Bias | 42738 |
| Woman | -0.045 | 0.002 | 0.000 | 200.000 | Straight Bias | 42738 |
| College | -0.001 | 0.002 | 0.563 | 200.000 | Straight Bias | 42738 |
| Religious | 0.076 | 0.002 | 0.000 | 200.000 | Straight Bias | 42738 |
| Nonmetro | 0.010 | 0.003 | 0.001 | 200.000 | Straight Bias | 42738 |
| California | -0.013 | 0.003 | 0.000 | 200.000 | Straight Bias | 42738 |
| Pennsylvania | -0.002 | 0.004 | 0.589 | 200.000 | Straight Bias | 42738 |
| New York | -0.023 | 0.004 | 0.000 | 200.000 | Straight Bias | 42738 |
| Florida | -0.005 | 0.005 | 0.332 | 200.000 | Straight Bias | 42738 |
| Illinois | -0.005 | 0.004 | 0.271 | 200.000 | Straight Bias | 42738 |

Table C31: Influence of Control Covariates on D-Score (Part 1)

| Control | Control Coef | SE | n | Bandwidth | Outcome | N |
|---------------|--------------|-------|----------|-----------|---------|--------------|
| Liberal | 0.071 | 0.008 | <u>p</u> | 5 000 | D Score | 1487 |
| | -0.071 | 0.008 | 0.000 | 5.000 | D-Score | 1407 |
| White | -0.019 | 0.020 | 0.013 | 5.000 | D-Score | 1487 |
| Woman | -0.021 | 0.008 | 0.010 | 5.000 | D-Score | 1487 |
| College | -0.017 | 0.008 | 0.033 | 5.000 | D-Score | 1487 |
| Beligious | 0.033 | 0.008 | 0.000 | 5.000 | D-Score | 1487 |
| Nonmetro | 0.005 | 0.013 | 0.274 | 5.000 | D-Score | 1487 |
| California | -0.025 | 0.010 | 0.040 | 5.000 | D-Score | 1487 |
| Pennevlyania | 0.017 | 0.012 | 0.388 | 5.000 | D-Score | 1487 |
| New York | -0.006 | 0.013 | 0.500 | 5.000 | D-Score | 1487 |
| Florida | -0.005 | 0.014 | 0.005 | 5.000 | D-Score | 1487 |
| Illinois | 0.014 | 0.014 | 0.720 | 5.000 | D-Score | 1487 |
| Liberal | -0.071 | 0.020 | 0.001 | 10.000 | D-Score | 2630 |
| Age | 0.056 | 0.000 | 0.000 | 10.000 | D-Score | 2639 |
| White | -0.020 | 0.020 | 0.005 | 10.000 | D-Score | 2630 |
| Woman | -0.020 | 0.000 | 0.001 | 10.000 | D-Score | 2039 |
| College | -0.021 | 0.000 | 0.000 | 10.000 | D-Score | 2039 |
| Beligious | 0.035 | 0.000 | 0.020 | 10.000 | D-Score | 2630 |
| Nonmetro | 0.035 | 0.000 | 0.000 | 10.000 | D-Score | 2630 |
| California | 0.012 | 0.010 | 0.234 | 10.000 | D-Score | 2630 |
| Poppeylyonio | -0.010 | 0.009 | 0.073 | 10.000 | D-Score | 2039 |
| Now Vork | 0.013 | 0.013 | 0.372 | 10.000 | D-Score | 2039 |
| Florida | -0.012 | 0.011 | 0.200 | 10.000 | D-Score | 2039 |
| Illinoia | 0.018 | 0.011 | 0.105 | 10.000 | D-Score | 2039 |
| Liboral | 0.013 | 0.015 | 0.000 | 15.000 | D-Score | 2009 |
| Ago | -0.071 | 0.005 | 0.000 | 15.000 | D-Score | 2620 |
| Age White | 0.000 | 0.017 | 0.000 | 15.000 | D-Score | 2620 |
| Women | -0.020 | 0.005 | 0.000 | 15.000 | D-Score | 2620 |
| Collogo | -0.019 | 0.005 | 0.000 | 15.000 | D-Score | 2620 |
| Roligious | -0.015 | 0.005 | 0.003 | 15.000 | D-Score | 3638 |
| Nonmotro | 0.030 | 0.005 | 0.000 | 15.000 | D-Score | 2620 |
| Colifornio | 0.009 | 0.008 | 0.280 | 15.000 | D-Score | 2028 |
| Dopportionia | -0.015 | 0.007 | 0.041 | 15.000 | D-Score | 2028 |
| New Verla | 0.015 | 0.012 | 0.198 | 15.000 | D-Score | 2028 |
| Flowido | -0.010 | 0.010 | 0.324 | 15.000 | D-Score | 2028 |
| r ionda | 0.010 | 0.009 | 0.069 | 15.000 | D-Score | 3030 |
| Tillnois | 0.001 | 0.014 | 0.955 | 15.000 | D-Score | 3038 |
| Ago | -0.071 | 0.004 | 0.000 | 20.000 | D-Score | 4907 |
| Age White | 0.002 | 0.014 | 0.000 | 20.000 | D-Score | 4907 |
| Women | -0.015 | 0.004 | 0.000 | 20.000 | D-Score | 4907 |
| Callana | -0.018 | 0.004 | 0.000 | 20.000 | D-Score | 4907 |
| D ali ai anna | -0.014 | 0.004 | 0.001 | 20.000 | D-Score | 4907 |
| Religious | 0.038 | 0.004 | 0.000 | 20.000 | D-Score | 4907 |
| C-lif-mi- | 0.004 | 0.007 | 0.383 | 20.000 | D-Score | 4907 |
| Dopportionia | -0.000 | 0.000 | 0.351 | 20.000 | D-Score | 4907 |
| New Verla | 0.017 | 0.011 | 0.109 | 20.000 | D-Score | 4907 |
| New YORK | -0.011 | 0.009 | 0.180 | 20.000 | D-Score | 4907 |
| r Iorida | 0.014 | 0.008 | 0.077 | 20.000 | D-Score | 4907 |
| Tillnois | -0.005 | 0.011 | 0.034 | 20.000 | D-Score | 4907 |
| A | -0.072 | 0.004 | 0.000 | 25.000 | D-Score | 5925 F025 |
| Age | 0.000 | 0.015 | 0.000 | 25.000 | D-Score | 5925 F025 |
| white | -0.017 | 0.004 | 0.000 | 25.000 | D-Score | 5925 |
| woman | -0.019 | 0.004 | 0.000 | 25.000 | D-Score | 5925 |
| College | -0.011 | 0.004 | 0.004 | 25.000 | D-Score | 5925 |
| Religious | 0.038 | 0.004 | 0.000 | 25.000 | D-Score | 5925 |
| Nonmetro | 0.010 | 0.007 | 0.120 | 25.000 | D-Score | 5925 |
| California | -0.004 | 0.005 | 0.517 | 25.000 | D-Score | 5925 |
| rennsylvania | 0.016 | 0.010 | 0.111 | 25.000 | D-Score | 5925 5025 |
| new York | -0.009 | 0.008 | 0.277 | 25.000 | D-Score | 5925 5025 |
| r lorida | 0.012 | 0.007 | 0.119 | 25.000 | D-Score | 5925 |
| 11linois | 0.002 | 0.009 | 0.849 | 25.000 | D-Score | 5925 |
| Liberal | -0.073 | 0.003 | 0.000 | 30.000 | D-Score | 7689 |
| Age | 0.068 | 0.011 | 0.000 | 30.000 | D-Score | 7689 |
| white | -0.016 | 0.003 | 0.000 | 30.000 | D-Score | 7689 |
| woman | -0.021 | 0.003 | 0.000 | 30.000 | D-Score | 7689 |
| College | -0.011 | 0.003 | 0.001 | 30.000 | D-Score | 7689 |
| Keligious | 0.037 | 0.003 | 0.000 | 30.000 | D-Score | 7689 |

Table C32: Influence of Control Covariates on D-Score (Part 2)

| Control | Control Coef. | SE | р | Bandwidth | Outcome | Ν |
|--------------|---------------|-------|-------|-----------|---------|-------|
| Religious | 0.072 | 0.005 | 0.000 | 30.000 | D-Score | 7511 |
| Nonmetro | 0.004 | 0.008 | 0.605 | 30.000 | D-Score | 7511 |
| California | -0.004 | 0.006 | 0.532 | 30.000 | D-Score | 7511 |
| Pennsylvania | -0.005 | 0.011 | 0.658 | 30.000 | D-Score | 7511 |
| New York | -0.026 | 0.009 | 0.004 | 30.000 | D-Score | 7511 |
| Florida | 0.002 | 0.009 | 0.803 | 30.000 | D-Score | 7511 |
| Illinois | -0.003 | 0.012 | 0.812 | 30.000 | D-Score | 7511 |
| Liberal | -0.109 | 0.004 | 0.000 | 35.000 | D-Score | 9111 |
| Age | 0.051 | 0.014 | 0.000 | 35.000 | D-Score | 9111 |
| White | -0.016 | 0.004 | 0.000 | 35.000 | D-Score | 9111 |
| Woman | -0.037 | 0.004 | 0.000 | 35.000 | D-Score | 9111 |
| College | 0.002 | 0.004 | 0.643 | 35.000 | D-Score | 9111 |
| Religious | 0.073 | 0.004 | 0.000 | 35.000 | D-Score | 9111 |
| Nonmetro | 0.008 | 0.007 | 0.238 | 35.000 | D-Score | 9111 |
| Damornia | -0.003 | 0.006 | 0.613 | 35.000 | D-Score | 9111 |
| Pennsylvania | -0.005 | 0.010 | 0.642 | 35.000 | D-Score | 9111 |
| New York | -0.029 | 0.008 | 0.001 | 35.000 | D-Score | 9111 |
| Florida | 0.002 | 0.009 | 0.805 | 35.000 | D-Score | 9111 |
| Tillnois | -0.004 | 0.011 | 0.099 | 35.000 | D-Score | 9111 |
| Are | -0.109 | 0.004 | 0.000 | 40.000 | D-Score | 10519 |
| nge White | -0.017 | 0.013 | 0.000 | 40.000 | D-Score | 10519 |
| Woman | -0.017 | 0.004 | 0.000 | 40.000 | D-Score | 10519 |
| College | -0.037 | 0.004 | 0.000 | 40.000 | D-Score | 10510 |
| Religious | 0.002 | 0.004 | 0.001 | 40.000 | D-Score | 10510 |
| Nonmetro | 0.014 | 0.004 | 0.031 | 40.000 | D-Score | 10510 |
| California | -0.003 | 0.000 | 0.631 | 40.000 | D-Score | 10519 |
| Pennewlyania | -0.003 | 0.000 | 0.001 | 40.000 | D-Score | 10510 |
| New Vork | -0.005 | 0.005 | 0.000 | 40.000 | D-Score | 10510 |
| Florida | 0.002 | 0.008 | 0.817 | 40.000 | D-Score | 10519 |
| Illinois | -0.002 | 0.000 | 0.875 | 40.000 | D-Score | 10519 |
| Liberal | -0.111 | 0.004 | 0.000 | 45.000 | D-Score | 11827 |
| Age | 0.057 | 0.012 | 0.000 | 45 000 | D-Score | 11827 |
| White | -0.017 | 0.004 | 0.000 | 45.000 | D-Score | 11827 |
| Woman | -0.038 | 0.004 | 0.000 | 45 000 | D-Score | 11827 |
| College | 0.004 | 0.004 | 0.332 | 45.000 | D-Score | 11827 |
| Religious | 0.074 | 0.004 | 0.000 | 45.000 | D-Score | 11827 |
| Nonmetro | 0.012 | 0.006 | 0.041 | 45.000 | D-Score | 11827 |
| California | -0.007 | 0.005 | 0.191 | 45.000 | D-Score | 11827 |
| Pennsylvania | -0.005 | 0.009 | 0.594 | 45.000 | D-Score | 11827 |
| New York | -0.032 | 0.007 | 0.000 | 45.000 | D-Score | 11827 |
| Florida | 0.005 | 0.008 | 0.549 | 45.000 | D-Score | 11827 |
| Illinois | -0.009 | 0.009 | 0.315 | 45.000 | D-Score | 11827 |
| Liberal | -0.115 | 0.003 | 0.000 | 50.000 | D-Score | 13780 |
| Age | 0.064 | 0.011 | 0.000 | 50.000 | D-Score | 13780 |
| White | -0.018 | 0.003 | 0.000 | 50.000 | D-Score | 13780 |
| Woman | -0.039 | 0.003 | 0.000 | 50.000 | D-Score | 13780 |
| College | 0.001 | 0.003 | 0.838 | 50.000 | D-Score | 13780 |
| Religious | 0.070 | 0.004 | 0.000 | 50.000 | D-Score | 13780 |
| Nonmetro | 0.010 | 0.006 | 0.081 | 50.000 | D-Score | 13780 |
| California | -0.007 | 0.005 | 0.129 | 50.000 | D-Score | 13780 |
| Pennsylvania | -0.002 | 0.008 | 0.837 | 50.000 | D-Score | 13780 |
| New York | -0.031 | 0.007 | 0.000 | 50.000 | D-Score | 13780 |
| Florida | 0.003 | 0.008 | 0.747 | 50.000 | D-Score | 13780 |
| Illinois | -0.004 | 0.009 | 0.620 | 50.000 | D-Score | 13780 |
| Liberal | -0.113 | 0.002 | 0.000 | 200.000 | D-Score | 42738 |
| Age | 0.056 | 0.007 | 0.000 | 200.000 | D-Score | 42738 |
| White | -0.013 | 0.002 | 0.000 | 200.000 | D-Score | 42738 |
| Woman | -0.045 | 0.002 | 0.000 | 200.000 | D-Score | 42738 |
| College | -0.001 | 0.002 | 0.563 | 200.000 | D-Score | 42738 |
| Religious | 0.076 | 0.002 | 0.000 | 200.000 | D-Score | 42738 |
| Nonmetro | 0.010 | 0.003 | 0.001 | 200.000 | D-Score | 42738 |
| California | -0.013 | 0.003 | 0.000 | 200.000 | D-Score | 42738 |
| Pennsylvania | -0.002 | 0.004 | 0.589 | 200.000 | D-Score | 42738 |
| New York | -0.023 | 0.004 | 0.000 | 200.000 | D-Score | 42738 |
| Florido | | | | | | |
| FIOLICIA | -0.005 | 0.005 | 0.332 | 200.000 | D-Score | 42738 |

C.7.4 Temporal Durability (D-Score, No Controls)

| Table C33: Coefficients Characterizing Temporal Durability of Post-Pulse Effe | ect |
|---|-----|
| (D-Score, No Controls) | |

| Post-Pulse Coef. | SE | p-value | Outcome | Controls? | Days From Pulse Event | Ν |
|------------------|------|---------|------------|-----------|-----------------------|--------------|
| -0.00 | 0.01 | 0.94 | A. D-Score | No | 1.00 | 2030 |
| -0.01 | 0.01 | 0.46 | A. D-Score | No | 2.00 | 2195 |
| -0.01 | 0.01 | 0.08 | A. D-Score | No | 3.00 | 2198 |
| -0.01 | 0.01 | 0.03 | A. D-Score | No | 4.00 | 2275 |
| -0.02 | 0.01 | 0.01 | A. D-Score | No | 5.00 6.00 | 2440 2567 |
| -0.02 | 0.01 | 0.00 | A D-Score | No | 7.00 | 2609 |
| -0.02 | 0.01 | 0.00 | A. D-Score | No | 8.00 | 2790 |
| -0.02 | 0.01 | 0.00 | A. D-Score | No | 9.00 | 2862 |
| -0.01 | 0.01 | 0.04 | A. D-Score | No | 10.00 | 2821 |
| -0.01 | 0.01 | 0.02 | A. D-Score | No | 11.00 | 2768 |
| -0.01 | 0.01 | 0.06 | A. D-Score | No | 12.00 | 2871 |
| -0.02 | 0.01 | 0.00 | A. D-Score | No | 13.00 | 3198 |
| -0.02 | 0.00 | 0.00 | A. D-Score | No | 14.00 | 3507 |
| -0.02 | 0.00 | 0.00 | A. D-Score | No Na | 15.00 | 3662 |
| -0.02 | 0.00 | 0.00 | A. D-Score | No | 17.00 | 3791 |
| -0.02 | 0.00 | 0.00 | A. D-Score | No | 18.00 | 3833 |
| -0.02 | 0.00 | 0.00 | A. D-Score | No | 19.00 | 3845 |
| -0.02 | 0.00 | 0.00 | A. D-Score | No | 20.00 | 3818 |
| -0.02 | 0.00 | 0.00 | A. D-Score | No | 21.00 | 3856 |
| -0.02 | 0.00 | 0.00 | A. D-Score | No | 22.00 | 3934 |
| -0.02 | 0.00 | 0.00 | A. D-Score | No | 23.00 | 3989 |
| -0.03 | 0.00 | 0.00 | A. D-Score | No | 24.00 | 3982 |
| -0.03 | 0.00 | 0.00 | A. D-Score | NO | 25.00 | 3948 |
| -0.03 | 0.00 | 0.00 | A. D-Score | NO No | 20.00 | 3780 |
| -0.02 | 0.00 | 0.00 | A D-Score | No | 27.00 | 3532 |
| -0.02 | 0.00 | 0.00 | A. D-Score | No | 29.00 | 3445 |
| -0.02 | 0.00 | 0.00 | A. D-Score | No | 30.00 | 3299 |
| -0.03 | 0.00 | 0.00 | A. D-Score | No | 31.00 | 3327 |
| -0.03 | 0.00 | 0.00 | A. D-Score | No | 32.00 | 3389 |
| -0.03 | 0.00 | 0.00 | A. D-Score | No | 33.00 | 3383 |
| -0.03 | 0.00 | 0.00 | A. D-Score | No | 34.00 | 3334 |
| -0.03 | 0.00 | 0.00 | A. D-Score | No | 35.00 | 3371 |
| -0.03 | 0.00 | 0.00 | A. D-Score | NO No | 30.00 | 3388 |
| -0.03 | 0.00 | 0.00 | A D-Score | No | 38.00 | 3357 |
| -0.03 | 0.00 | 0.00 | A. D-Score | No | 39.00 | 3400 |
| -0.03 | 0.00 | 0.00 | A. D-Score | No | 40.00 | 3379 |
| -0.03 | 0.00 | 0.00 | A. D-Score | No | 41.00 | 3367 |
| -0.03 | 0.01 | 0.00 | A. D-Score | No | 42.00 | 3343 |
| -0.03 | 0.00 | 0.00 | A. D-Score | No | 43.00 | 3379 |
| -0.03 | 0.01 | 0.00 | A. D-Score | No | 44.00 | 3368 |
| -0.03 | 0.01 | 0.00 | A. D-Score | No | 45.00 | 3356 |
| -0.03 | 0.01 | 0.00 | A. D-Score | No | 40.00 | 3129 2864 |
| -0.02 | 0.01 | 0.00 | A D-Score | No | 48.00 | 2705 |
| -0.01 | 0.01 | 0.01 | A. D-Score | No | 49.00 | 2635 |
| -0.01 | 0.01 | 0.01 | A. D-Score | No | 50.00 | 2548 |
| -0.02 | 0.01 | 0.01 | A. D-Score | No | 51.00 | 2449 |
| -0.01 | 0.01 | 0.02 | A. D-Score | No | 52.00 | 2344 |
| -0.01 | 0.01 | 0.04 | A. D-Score | No | 53.00 | 2236 |
| -0.02 | 0.01 | 0.02 | A. D-Score | No | 54.00 | 2068 |
| -0.02 | 0.01 | 0.04 | A. D-Score | NO No | 55.00 56.00 | 2000 |
| -0.02 | 0.01 | 0.05 | A D-Score | No | 57.00 | 1970 |
| -0.02 | 0.01 | 0.07 | A. D-Score | No | 58.00 | 1917 |
| -0.01 | 0.01 | 0.21 | A. D-Score | No | 59.00 | 1919 |
| -0.00 | 0.01 | 0.61 | A. D-Score | No | 60.00 | 1943 |
| -0.01 | 0.01 | 0.42 | A. D-Score | No | 61.00 | 2036 |
| -0.01 | 0.01 | 0.30 | A. D-Score | No | 62.00 | 2098 |
| -0.00 | 0.01 | 0.54 | A. D-Score | No | 63.00 | 2192 |
| -0.01 | 0.01 | 0.37 | A. D.Score | No | 64.00 | 2351 |
| -0.01 | 0.01 | 0.09 | A. D.Score | No | 66.00 | 2515 2672 |
| -0.01 | 0.01 | 0.07 | A D-Score | No | 67.00 | 2012 |
| -0.01 | 0.01 | 0.14 | A. D-Score | No | 68.00 | 3007 |
| -0.01 | 0.01 | 0.23 | A. D-Score | No | 69.00 | 3137 |
| -0.00 | 0.00 | 0.41 | A. D-Score | No | 70.00 | 3351 |
| -0.00 | 0.00 | 0.55 | A. D-Score | No | 71.00 | 3598 |
| -0.00 | 0.00 | 0.41 | A. D-Score | No | 72.00 | 3975 |

C.7.5 Temporal Durability (D-Score, With Controls)

Table C34: Coefficients Characterizing Temporal Durability of Post-Pulse Effect (D-Score, With Controls)

| Post-Pulse Coef. | SE | p-value | Outcome | Controls? | Days From Pulse Event | Ν |
|------------------|------|---------|------------|-------------|-----------------------|--------------|
| 0.00 | 0.01 | 0.72 | A. D-Score | Yes | 1.00 | 2030 |
| -0.00 | 0.01 | 0.68 | A. D-Score | Yes | 2.00 | 2195 |
| -0.01 | 0.01 | 0.08 | A. D-Score | Yes | 3.00 | 2198 |
| -0.01 | 0.01 | 0.05 | A. D-Score | Yes | 4.00 | 2275 |
| -0.01 | 0.01 | 0.03 | A. D-Score | Yes | 5.00 | 2440 |
| -0.02 | 0.01 | 0.01 | A. D-Score | Yes | 0.00 | 2007 |
| -0.01 | 0.01 | 0.02 | A. D-Score | Ves | 7.00 | 2009 |
| -0.01 | 0.01 | 0.01 | A. D-Score | Yes | 9.00 | 2862 |
| -0.01 | 0.01 | 0.08 | A. D-Score | Yes | 10.00 | 2821 |
| -0.01 | 0.01 | 0.02 | A. D-Score | Yes | 11.00 | 2768 |
| -0.01 | 0.01 | 0.10 | A. D-Score | Yes | 12.00 | 2871 |
| -0.01 | 0.00 | 0.01 | A. D-Score | Yes | 13.00 | 3198 |
| -0.01 | 0.00 | 0.00 | A. D-Score | Yes | 14.00 | 3507 |
| -0.01 | 0.00 | 0.01 | A. D-Score | Yes | 15.00 | 3662 |
| -0.01 | 0.00 | 0.01 | A. D-Score | Yes | 16.00 | 3747 |
| -0.01 | 0.00 | 0.00 | A. D-Score | Voe | 17.00 | 3833 |
| -0.02 | 0.00 | 0.00 | A D-Score | Ves | 19.00 | 3845 |
| -0.02 | 0.00 | 0.00 | A. D-Score | Yes | 20.00 | 3818 |
| -0.02 | 0.00 | 0.00 | A. D-Score | Yes | 21.00 | 3856 |
| -0.02 | 0.00 | 0.00 | A. D-Score | Yes | 22.00 | 3934 |
| -0.02 | 0.00 | 0.00 | A. D-Score | Yes | 23.00 | 3989 |
| -0.02 | 0.00 | 0.00 | A. D-Score | Yes | 24.00 | 3982 |
| -0.02 | 0.00 | 0.00 | A. D-Score | Yes | 25.00 | 3948 |
| -0.02 | 0.00 | 0.00 | A. D-Score | Yes | 26.00 | 3785 |
| -0.02 | 0.00 | 0.00 | A. D-Score | Yes | 27.00 | 3638 |
| -0.02 | 0.00 | 0.00 | A. D-Score | Yes | 28.00 | 3032 3445 |
| -0.02 | 0.00 | 0.00 | A D-Score | Ves | 29.00 | 3200 |
| -0.02 | 0.00 | 0.00 | A. D-Score | Yes | 31.00 | 3327 |
| -0.02 | 0.00 | 0.00 | A. D-Score | Yes | 32.00 | 3389 |
| -0.02 | 0.00 | 0.00 | A. D-Score | Yes | 33.00 | 3383 |
| -0.02 | 0.00 | 0.00 | A. D-Score | Yes | 34.00 | 3334 |
| -0.02 | 0.00 | 0.00 | A. D-Score | Yes | 35.00 | 3371 |
| -0.02 | 0.00 | 0.00 | A. D-Score | Yes | 36.00 | 3388 |
| -0.02 | 0.00 | 0.00 | A. D-Score | Yes | 37.00 | 3361 |
| -0.02 | 0.00 | 0.00 | A. D-Score | Yes | 38.00 | 3357 |
| -0.02 | 0.00 | 0.00 | A D-Score | Ves | 40.00 | 3400 |
| -0.02 | 0.00 | 0.00 | A. D-Score | Yes | 41.00 | 3367 |
| -0.02 | 0.00 | 0.00 | A. D-Score | Yes | 42.00 | 3343 |
| -0.02 | 0.00 | 0.00 | A. D-Score | Yes | 43.00 | 3379 |
| -0.02 | 0.00 | 0.00 | A. D-Score | Yes | 44.00 | 3368 |
| -0.02 | 0.00 | 0.00 | A. D-Score | Yes | 45.00 | 3356 |
| -0.02 | 0.01 | 0.00 | A. D-Score | Yes | 46.00 | 3129 |
| -0.01 | 0.01 | 0.04 | A. D-Score | Yes | 47.00 | 2864 |
| -0.01 | 0.01 | 0.13 | A. D-Score | Yes | 48.00 | 2705 |
| -0.01 | 0.01 | 0.21 | A. D-Score | Voe | 49.00 | 2050 |
| -0.01 | 0.01 | 0.15 | A. D-Score | Yes | 51.00 | 2449 |
| -0.01 | 0.01 | 0.23 | A. D-Score | Yes | 52.00 | 2344 |
| -0.01 | 0.01 | 0.21 | A. D-Score | Yes | 53.00 | 2236 |
| -0.01 | 0.01 | 0.10 | A. D-Score | Yes | 54.00 | 2068 |
| -0.01 | 0.01 | 0.14 | A. D-Score | Yes | 55.00 | 2000 |
| -0.01 | 0.01 | 0.12 | A. D-Score | Yes | 56.00 | 1970 |
| -0.02 | 0.01 | 0.09 | A. D-Score | Yes | 57.00 | 1934 |
| -0.01 | 0.01 | 0.14 | A. D-Score | Yes | 58.00 | 1917 |
| -0.01 | 0.01 | 0.43 | A. D-Score | Yes | 59.00 60.00 | 1919 |
| -0.00 | 0.01 | 0.00 | A D-Score | Yes | 61.00 | 2036 |
| -0.00 | 0.01 | 0.64 | A. D-Score | Yes | 62.00 | 2098 |
| -0.00 | 0.01 | 0.97 | A. D-Score | Yes | 63.00 | 2192 |
| -0.00 | 0.01 | 0.70 | A. D-Score | Yes | 64.00 | 2351 |
| -0.01 | 0.01 | 0.28 | A. D-Score | Yes | 65.00 | 2515 |
| -0.01 | 0.01 | 0.29 | A. D-Score | Yes | 66.00 | 2672 |
| -0.01 | 0.01 | 0.32 | A. D-Score | Yes | 67.00 | 2829 |
| -0.01 | 0.00 | 0.30 | A. D-Score | Yes | 68.00 | 3007 |
| -0.00 | 0.00 | 0.36 | A. D.Score | Yes | 69.00 | 3137 |
| -0.00 | 0.00 | 0.32 | A. D.Score | 1 es Voc | 70.00 | 3508 3508 |
| -0.01 | 0.00 | 0.41 | A. D-Score | Yes | 72.00 | 3975 |

C.7.6 Temporal Durability (Heterocentrism, No Controls)

Table C35: Coefficients Characterizing Temporal Durability of Post-Pulse Effect(Heterocentrism, No Controls)

| Post-Pulse Coef. | SE | p-value | Outcome | Controls? | Days From Pulse Event | Ν |
|------------------|------|---------|--|-----------|-----------------------|--------------|
| -0.01 | 0.01 | 0.09 | B. Heterocentrism | No | 1.00 | 2025 |
| -0.02 | 0.01 | 0.00 | B. Heterocentrism | No | 2.00 | 2191 |
| -0.02 | 0.01 | 0.00 | B. Heterocentrism | No | 3.00 | 2196 |
| -0.02 | 0.01 | 0.01 | B. Heterocentrism D. Heterocentrism | No Na | 4.00 | 2276 |
| -0.02 | 0.01 | 0.01 | B. Heterocentrism | No | 5.00 | 2440 2572 |
| -0.02 | 0.01 | 0.00 | B Heterocentrism | No | 7.00 | 2613 |
| -0.01 | 0.01 | 0.01 | B. Heterocentrism | No | 8.00 | 2792 |
| -0.01 | 0.00 | 0.00 | B. Heterocentrism | No | 9.00 | 2864 |
| -0.01 | 0.00 | 0.01 | B. Heterocentrism | No | 10.00 | 2822 |
| -0.01 | 0.00 | 0.01 | B. Heterocentrism | No | 11.00 | 2770 |
| -0.01 | 0.00 | 0.03 | B. Heterocentrism | No | 12.00 | 2874 |
| -0.01 | 0.00 | 0.00 | B. Heterocentrism | No | 13.00 | 3198 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism D. Heterocentrism | NO Nu | 14.00 | 3513 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 15.00 | 3754 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 17.00 | 3802 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 18.00 | 3843 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 19.00 | 3854 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 20.00 | 3827 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 21.00 | 3864 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 22.00 | 3941 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 23.00 | 3995 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism B. Heterocentrism | No | 24.00 | 3988 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 25.00 | 3788 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 27.00 | 3641 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 28.00 | 3533 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 29.00 | 3443 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 30.00 | 3296 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 31.00 | 3327 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 32.00 | 3382 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 33.00 | 3373 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No Na | 34.00 | 3323 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 36.00 36.00 | 3376 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 37.00 | 3347 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 38.00 | 3345 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 39.00 | 3386 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 40.00 | 3365 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 41.00 | 3352 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 42.00 | 3329 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No Na | 43.00 | 3364 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 44.00 | 2248 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 45.00 | 3913 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 47.00 | 3022 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 48.00 | 2912 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 49.00 | 2885 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 50.00 | 2900 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 51.00 | 2915 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | NO | 52.00 | 2921 |
| -0.02 | 0.00 | 0.00 | D. Heterocentrism | No | 53.00 54.00 | 2901 2802 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 55.00 | 2002 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 56.00 | 2801 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 57.00 | 2882 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 58.00 | 2917 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 59.00 | 2920 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 60.00 | 2936 |
| -0.01 | 0.00 | 0.00 | B. Heterocentrism | No | 61.00 | 2936 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | No | 62.00 | 2925 |
| -0.01 | 0.00 | 0.00 | B Heterocentrism | No | 64.00 | 2974 |
| -0.01 | 0.00 | 0.01 | B. Heterocentrism | No | 65.00 | 3156 |
| -0.01 | 0.00 | 0.02 | B. Heterocentrism | No | 66.00 | 3201 |
| -0.01 | 0.00 | 0.05 | B. Heterocentrism | No | 67.00 | 3255 |
| -0.01 | 0.00 | 0.04 | B. Heterocentrism | No | 68.00 | 3344 |
| -0.01 | 0.00 | 0.15 | B. Heterocentrism | No | 69.00 | 3408 |
| -0.00 | 0.00 | 0.40 | B. Heterocentrism | No | 70.00 | 3587 |
| -0.00 | 0.00 | 0.48 | B. Heterocentrism | NO | 71.00 | 3776 |
| -0.00 | 0.00 | 0.87 | D. Heterocentrism | INO | 72.00 | 4037 |

C.7.7 Temporal Durability (Heterocentrism, With Controls)

Table C36: Coefficients Characterizing Temporal Durability of Post-Pulse Effect(Heterocentrism, With Controls)

| Post-Pulse Coef. | SE | p-value | Outcome | Controls? | Days From Pulse Event | Ν |
|------------------|------|---------|--|-------------|-----------------------|--------------|
| -0.01 | 0.01 | 0.20 | B. Heterocentrism | Yes | 1.00 | 2025 |
| -0.01 | 0.01 | 0.01 | B. Heterocentrism | Yes | 2.00 | 2191 |
| -0.02 | 0.01 | 0.01 | B. Heterocentrism | Yes | 3.00 | 2196 |
| -0.01 | 0.01 | 0.02 | B. Heterocentrism | Yes | 4.00 | 2276 |
| -0.01 | 0.01 | 0.04 | B. Heterocentrism B. Heterocentrism | Yes | 5.00 | 2440 |
| -0.01 | 0.01 | 0.01 | B. Heterocentrism | Tes Voc | 0.00 | 2012 |
| -0.01 | 0.00 | 0.01 | B. Heterocentrism | Ves | 8.00 | 2013 |
| -0.01 | 0.00 | 0.01 | B. Heterocentrism B. Heterocentrism | Yes | 9.00 | 2864 |
| -0.01 | 0.00 | 0.04 | B. Heterocentrism | Yes | 10.00 | 2822 |
| -0.01 | 0.00 | 0.01 | B. Heterocentrism | Yes | 11.00 | 2770 |
| -0.01 | 0.00 | 0.05 | B. Heterocentrism | Yes | 12.00 | 2874 |
| -0.01 | 0.00 | 0.02 | B. Heterocentrism | Yes | 13.00 | 3198 |
| -0.01 | 0.00 | 0.00 | B. Heterocentrism | Yes | 14.00 | 3513 |
| -0.01 | 0.00 | 0.00 | B. Heterocentrism | Yes | 15.00 | 3670 |
| -0.01 | 0.00 | 0.00 | B. Heterocentrism | Yes | 16.00 | 3754 |
| -0.01 | 0.00 | 0.00 | B. Heterocentrism | Yes | 17.00 | 3802 |
| -0.01 | 0.00 | 0.00 | B. Heterocentrism | Tes Voc | 10.00 | 2054 |
| -0.01 | 0.00 | 0.00 | B. Heterocentrism | Voe | 20.00 | 3897 |
| -0.01 | 0.00 | 0.00 | B Heterocentrism | Ves | 20.00 | 3864 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | Yes | 22.00 | 3941 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | Yes | 23.00 | 3995 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | Yes | 24.00 | 3988 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | Yes | 25.00 | 3950 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | Yes | 26.00 | 3788 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | Yes | 27.00 | 3641 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism | Yes | 28.00 | 3533 |
| -0.01 | 0.00 | 0.00 | B. Heterocentrism | Yes | 29.00 | 3443 |
| -0.01 | 0.00 | 0.00 | B. Heterocentrism | Yes | 30.00 | 3296 |
| -0.01 | 0.00 | 0.00 | B. Heterocentrism | Yes | 31.00 | 3327 |
| -0.02 | 0.00 | 0.00 | B. Heterocentrism B. Heterocentrism | Yes | 32.00 | 3382 |
| -0.01 | 0.00 | 0.00 | B. Heterocentrism | Voe | 34.00 | 3393 |
| -0.02 | 0.00 | 0.00 | B Heterocentrism | Ves | 35.00 | 3360 |
| -0.01 | 0.00 | 0.00 | B. Heterocentrism B. Heterocentrism | Yes | 36.00 | 3376 |
| -0.01 | 0.00 | 0.00 | B. Heterocentrism | Yes | 37.00 | 3347 |
| -0.01 | 0.00 | 0.00 | B. Heterocentrism | Yes | 38.00 | 3345 |
| -0.01 | 0.00 | 0.02 | B. Heterocentrism | Yes | 39.00 | 3386 |
| -0.01 | 0.00 | 0.02 | B. Heterocentrism | Yes | 40.00 | 3365 |
| -0.01 | 0.00 | 0.01 | B. Heterocentrism | Yes | 41.00 | 3352 |
| -0.01 | 0.00 | 0.01 | B. Heterocentrism | Yes | 42.00 | 3329 |
| -0.01 | 0.00 | 0.01 | B. Heterocentrism | Yes | 43.00 | 3364 |
| -0.01 | 0.00 | 0.01 | B. Heterocentrism | Yes | 44.00 | 3354 |
| -0.01 | 0.00 | 0.00 | B. Heterocentrism | Yes | 45.00 | 3348 |
| -0.01 | 0.00 | 0.01 | B. Heterocentrism | Yes | 46.00 | 3213 |
| -0.01 | 0.00 | 0.10 | B. Heterocentrism | Voe | 47.00 | 3022 2012 |
| -0.01 | 0.00 | 0.12 | B Heterocentrism | Ves | 49.00 | 2885 |
| -0.01 | 0.00 | 0.10 | B Heterocentrism | Yes | 50.00 | 2900 |
| -0.01 | 0.00 | 0.14 | B. Heterocentrism | Yes | 51.00 | 2915 |
| -0.01 | 0.00 | 0.07 | B. Heterocentrism | Yes | 52.00 | 2921 |
| -0.01 | 0.00 | 0.10 | B. Heterocentrism | Yes | 53.00 | 2901 |
| -0.01 | 0.00 | 0.02 | B. Heterocentrism | Yes | 54.00 | 2802 |
| -0.01 | 0.00 | 0.02 | B. Heterocentrism | Yes | 55.00 | 2770 |
| -0.01 | 0.00 | 0.02 | B. Heterocentrism | Yes | 56.00 | 2801 |
| -0.01 | 0.00 | 0.03 | B. Heterocentrism | Yes | 57.00 | 2882 |
| -0.01 | 0.00 | 0.05 | B. Heterocentrism | Yes | 58.00 | 2917 |
| -0.01 | 0.00 | 0.03 | B. Heterocentrism | Yes | 59.00 | 2920 |
| -0.01 | 0.00 | 0.09 | B. Heterocentrism | Yes | 60.00 | 2936 |
| -0.01 | 0.00 | 0.17 | B. Heterocentrism | 1 es Voc | 62.00 | 2930 2025 |
| -0.01 | 0.00 | 0.05 | B Heterocentrism | Yes | 63.00 | 2920 |
| -0.01 | 0.00 | 0.07 | B. Heterocentrism | Yes | 64.00 | 3092 |
| -0.01 | 0.00 | 0.15 | B. Heterocentrism | Yes | 65.00 | 3156 |
| -0.00 | 0.00 | 0.27 | B. Heterocentrism | Yes | 66.00 | 3201 |
| -0.00 | 0.00 | 0.34 | B. Heterocentrism | Yes | 67.00 | 3255 |
| -0.01 | 0.00 | 0.13 | B. Heterocentrism | Yes | 68.00 | 3344 |
| -0.00 | 0.00 | 0.25 | B. Heterocentrism | Yes | 69.00 | 3408 |
| -0.00 | 0.00 | 0.34 | B. Heterocentrism | Yes | 70.00 | 3587 |
| -0.00 | 0.00 | 0.31 | B. Heterocentrism | Yes | 71.00 | 3776 |
| -0.00 | 0.00 | 0.37 | B. Heterocentrism | Yes | 72.00 | 4037 |

C.7.8 Temporal Durability (D-Score Outcome, Control Coefficients)

Table C37: Control Coefficients For Models Characterizing Temporal Durability of Post-Pulse Effect (D-Score, With Controls, Part 1)



Table C38: Control Coefficients For Models Characterizing Temporal Durability of Post-Pulse Effect (D-Score, With Controls, Part 2)



Table C39: Control Coefficients For Models Characterizing Temporal Durability of Post-Pulse Effect (D-Score, With Controls, Part 3)



Table C40: Control Coefficients For Models Characterizing Temporal Durability of Post-Pulse Effect (D-Score, With Controls, Part 4)



Table C41: Control Coefficients For Models Characterizing Temporal Durability of Post-Pulse Effect (D-Score, With Controls, Part 5)



Table C42: Control Coefficients For Models Characterizing Temporal Durability of Post-
Pulse Effect (D-Score, With Controls, Part 6)

| covar_name | est | se 0.01 | pv 0.78 | nobs | days_cut |
|--|---|--|--|--|--|
| California | -0.01 | 0.01 | 0.18 | 2192 2192 | 63 |
| Pennsylvania | 0.05 | 0.02 | 0.00 | 2192 | 63 |
| New York | -0.01 | 0.01 | 0.41 | 2192 | 63 |
| Florida Illinoio | -0.01 | 0.01 | 0.65 | 2192 | 63 |
| Liberal | -0.01 | 0.02 | 0.02 | 2351 | 64 |
| Age | 0.03 | 0.02 | 0.11 | 2351 | 64 |
| White | -0.01 | 0.01 | 0.08 | 2351 | 64 |
| Woman | -0.02 | 0.01 | 0.00 | 2351 | 64 |
| College | -0.02 | 0.01 | 0.01 | 2351 | 64 |
| Non-Metro | -0.00 | 0.01 | 0.00 | 2351 | 64 |
| California | -0.01 | 0.01 | 0.26 | 2351 | 64 |
| Pennsylvania | 0.04 | 0.02 | 0.04 | 2351 | 64 |
| New York | -0.01 | 0.01 | 0.63 | 2351 | 64 |
| Florida | -0.01 | 0.01 | 0.50 | 2351 | 64 |
| Illinois | -0.01 | 0.02 | 0.74 | 2351 | 64 |
| Liberal | -0.08 | 0.01 | 0.00 | 2515 | 65 |
| Age White | -0.04 | 0.02 | 0.03 | 2515 | 65 |
| Woman | -0.02 | 0.01 | 0.00 | 2515 | 65 |
| College | -0.02 | 0.01 | 0.00 | 2515 | 65 |
| Religious | 0.03 | 0.01 | 0.00 | 2515 | 65 |
| Non-Metro | -0.00 | 0.01 | 1.00 | 2515 | 65 |
| California | -0.01 | 0.01 | 0.43 | 2515 | 65 |
| Pennsylvania | 0.03 | 0.02 | 0.07 | 2515 | 65 |
| New York | -0.01 | 0.01 | 0.66 | 2515 | 65 |
| Florida | -0.01 | 0.01 | 0.40 | 2515 | 65 |
| Illinois L'hand | -0.01 | 0.02 | 0.72 | 2515 | 65 |
| Am | -0.08 | 0.01 | 0.00 | 2672 | 66 |
| White | -0.01 | 0.01 | 0.04 | 2672 | 66 |
| Woman | -0.02 | 0.01 | 0.00 | 2672 | 66 |
| College | -0.02 | 0.01 | 0.00 | 2672 | 66 |
| Religious | 0.03 | 0.01 | 0.00 | 2672 | 66 |
| Non-Metro | 0.00 | 0.01 | 0.71 | 2672 | 66 |
| California | -0.01 | 0.01 | 0.39 | 2672 | 66 |
| Pennsylvania | 0.04 | 0.02 | 0.03 | 2672 | 66 |
| New York Florido | -0.01 | 0.01 | 0.54 | 2672 | 66 |
| r 10110a Illinois | -0.01 | 0.01 | 0.34 | 2072 | 00 66 |
| Liberal | -0.08 | 0.02 | 0.00 | 2829 | 67 |
| Age | 0.06 | 0.02 | 0.00 | 2829 | 67 |
| White | -0.01 | 0.01 | 0.06 | 2829 | 67 |
| Woman | -0.02 | 0.01 | 0.00 | 2829 | 67 |
| College | -0.02 | 0.01 | 0.00 | 2829 | 67 |
| Religious | 0.03 | 0.01 | 0.00 | 2829 | 67 |
| Non-Metro | 0.00 | 0.01 | 0.87 | 2829 | 67 |
| California | -0.01 | 0.01 | 0.40 | 2829 | 67 |
| Pennsylvania New Verb | 0.04 | 0.02 | 0.03 | 2829 | 67 |
| New TOTK Florido | -0.01 | 0.01 | 0.40 | 2829 | 67 |
| Illinoio | 0.01 | 0.01 | 0.35 | 2828 | 67 |
| Liberal | -0.01 | 0.02 | 0.00 | 3007 | 68 |
| Age | 0.06 | 0.02 | 0.00 | 3007 | 68 |
| White | -0.01 | 0.01 | 0.07 | 3007 | 68 |
| Woman | -0.02 | 0.01 | 0.00 | 3007 | 68 |
| College | -0.02 | 0.01 | 0.00 | 3007 | 68 |
| Religious | 0.03 | 0.01 | 0.00 | 3007 | 68 |
| Non-Metro | 0.01 | 0.01 | 0.46 | 3007 | 68 |
| California | -0.01 | 0.01 | 0.37 | 3007 | 68 |
| Pennsylvania Nom Vork | 0.04 | 0.02 | 0.02 | 3007 | 68 |
| Florida | -0.01 | 0.01 | 0.34 | 3007 | 68 |
| Illinois | -0.01 | 0.01 | 0.53 | 3007 | 68 |
| Liberal | -0.08 | 0.01 | 0.00 | 3137 | 69 |
| Age | 0.06 | 0.02 | 0.00 | 3137 | 69 |
| White | -0.01 | 0.01 | 0.06 | 3137 | 69 |
| Woman | -0.02 | 0.01 | 0.00 | 3137 | 69 |
| College | -0.02 | 0.01 | 0.00 | 3137 | 69 |
| Religious | 0.03 | 0.01 | 0.00 | 3137 | 69 |
| Non-Metro | 0.01 | 0.01 | 0.42 | 3137 | 69 |
| Camornia Pennsylvania | -0.01 | 0.01 | 0.24 | 3137 | 60 09 |
| New York | -0.04 | 0.02 | 0.45 | 3137 | 0.5 |
| Florida | -0.01 | 0.01 | 10.1413 | 3137 | 60 |
| Illinois | -0.01 | 0.01 | 0.26 | 3137 3137 | 69 69 |
| 1.1 | -0.01 -0.01 | 0.01 0.01 | $0.26 \\ 0.42$ | 3137 3137 3137 | 69 69 69 |
| Laberai | -0.01 -0.01 -0.08 | 0.01 0.01 0.00 | $\begin{array}{c} 0.26 \\ 0.42 \\ 0.00 \end{array}$ | 3137 3137 3137 3351 | 69 69 69 70 |
| Age | -0.01 -0.01 -0.08 0.05 | 0.01 0.01 0.00 0.02 | $\begin{array}{c} 0.26 \\ 0.42 \\ 0.00 \\ 0.00 \end{array}$ | 3137 3137 3137 3351 3351 | 69 69 70 70 |
| Age White | -0.01 -0.01 -0.08 0.05 -0.01 | 0.01 0.01 0.00 0.02 0.00 | 0.26 0.42 0.00 0.00 0.06 | 3137 3137 3137 3351 3351 3351 3351 | 69 69 70 70 70 |
| Age White Woman | -0.01 -0.08 0.05 -0.01 -0.02 0.05 | 0.01 0.01 0.00 0.02 0.00 0.00 0.00 | 0.26 0.42 0.00 0.00 0.06 0.00 0.00 | 3137 3137 3137 3351 3351 3351 3351 3351 | 69 69 70 70 70 70 70 |
| Age White Woman College Religious | -0.01 -0.01 -0.08 0.05 -0.01 -0.02 -0.02 0.03 | 0.01 0.00 0.02 0.00 0.00 0.00 0.00 0.01 | 0.26 0.42 0.00 0.00 0.06 0.00 0.00 0.00 0.00 | 3137 3137 3351 3351 3351 3351 3351 3351 | 69 69 70 70 70 70 70 70 70 70 |
| Age White Woman College Religious Non-Metro | -0.01 -0.08 0.05 -0.01 -0.02 -0.02 0.03 0.01 | 0.01 0.01 0.00 0.02 0.00 0.00 0.00 0.00 0.01 0.01 | 0.26 0.42 0.00 0.00 0.06 0.00 0.00 0.00 0.33 | 3137 3137 3137 3351 3351 3351 3351 3351 | 69 69 70 70 70 70 70 70 70 70 70 |
| Age White Woman College Religious Non-Metro California | -0.01 -0.08 0.05 -0.01 -0.02 -0.02 0.03 0.01 -0.01 | 0.01 0.01 0.00 0.02 0.00 0.00 0.00 0.01 0.01 0.01 | $\begin{array}{c} 0.26 \\ 0.42 \\ 0.00 \\ 0.00 \\ 0.00 \\ 0.00 \\ 0.00 \\ 0.33 \\ 0.19 \end{array}$ | 3137 3137 3351 3351 3351 3351 3351 3351 | 69 69 70 70 70 70 70 70 70 70 70 70 |
| Agerai Age White Woman College Religious Non-Metro California Pennsylvania | -0.01 -0.08 0.05 -0.01 -0.02 -0.02 0.03 0.01 -0.01 0.03 | 0.01 0.01 0.00 0.02 0.00 0.00 0.00 0.01 0.01 0.01 0.02 | $\begin{array}{c} 0.26 \\ 0.42 \\ 0.00 \\ 0.06 \\ 0.00 \\ 0.00 \\ 0.00 \\ 0.33 \\ 0.19 \\ 0.04 \end{array}$ | 3137 3137 3137 3351 3351 3351 3351 3351 | 69 69 70 70 70 70 70 70 70 70 70 70 70 |
| Aberai Age White Woman College Religious Non-Metro California Pennsylvania New York | -0.01 -0.01 -0.08 0.05 -0.01 -0.02 0.03 0.01 -0.01 0.03 -0.01 | 0.01 0.01 0.02 0.00 0.00 0.00 0.01 0.01 0.01 0.02 0.01 0.02 0.01 | $\begin{array}{c} 0.26 \\ 0.42 \\ 0.00 \\ 0.00 \\ 0.00 \\ 0.00 \\ 0.00 \\ 0.00 \\ 0.33 \\ 0.19 \\ 0.04 \\ 0.58 \end{array}$ | 3137 3137 3137 3351 3351 3351 3351 3351 | 69 69 70 70 70 70 70 70 70 70 70 70 70 70 |
| Age White Woman College Religious Non-Metro California Pennsylvania New York Florida | -0.01 -0.01 -0.08 0.05 -0.02 -0.02 0.03 0.01 -0.01 0.03 -0.01 -0.01 | 0.01 0.01 0.02 0.00 0.00 0.00 0.01 0.01 0.01 0.02 0.01 0.01 0.02 0.01 0.02 0.01 0.01 0.01 0.00 | $\begin{array}{c} 0.26\\ 0.42\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.33\\ 0.19\\ 0.04\\ 0.58\\ 0.21\\ \end{array}$ | 3137 3137 3137 3351 3351 3351 3351 3351 | 69 69 70 70 70 70 70 70 70 70 70 70 70 70 |
| Age White Woman College Religious Non-Metro California Pennsylvania New York Florida Illinois Ubergi | -0.01 -0.01 -0.08 0.05 -0.01 -0.02 -0.02 0.03 0.01 -0.01 -0.01 -0.01 -0.01 -0.01 | 0.01 0.01 0.00 0.02 0.00 0.00 0.00 0.01 0.01 0.01 0.01 0.01 0.01 0.02 0.01 0.02 0.00 0.01 | 0.26 0.42 0.00 0.00 0.00 0.00 0.00 0.00 0.33 0.19 0.04 0.58 0.21 0.37 0.27 | 3137 3137 3137 3351 3351 3351 3351 3351 | 69 69 70 70 70 70 70 70 70 70 70 70 70 |
| Liberal Age White Woman College Religious Non-Metro California Pennsylvania New York Florida Illinois Liberal | -0.01 -0.01 -0.08 0.05 -0.01 -0.02 -0.02 0.03 0.01 -0.01 0.03 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 | 0.01 0.01 0.00 0.02 0.00 0.00 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.02 | 0.26 0.42 0.00 0.00 0.00 0.00 0.00 0.33 0.19 0.04 0.58 0.21 0.37 0.00 0.37 | 3137 3137 3137 3351 3351 3351 3351 3351 | 69 69 70 70 70 70 70 70 70 70 70 70 70 70 70 |
| Liberal Age White Woman College Religious Non-Metro California Pennsylvania New York Florida Illinois Liberal Age White | -0.01 -0.08 0.05 -0.01 -0.02 -0.02 0.03 0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.08 0.05 -0.01 | 0.01 0.01 0.00 0.02 0.00 0.00 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.02 0.01 0.02 0.00 | 0.26 0.42 0.00 0.00 0.00 0.00 0.00 0.00 0.33 0.19 0.04 0.58 0.21 0.37 0.00 0.00 0.00 0.00 | 3137 3137 3137 3351 3351 3351 3351 3351 | 69 69 70 70 70 70 70 70 70 70 70 70 70 70 70 |
| Liberal Age White Woman College Religious Non-Metro California Pennsylvania New York Florida Illinois Liberal Age White Woman | -0.01 -0.08 0.05 -0.01 -0.02 -0.02 0.03 0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.08 0.05 -0.01 -0.01 -0.08 | 0.01 0.01 0.00 0.02 0.00 0.00 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.02 0.00 | 0.26 0.42 0.00 0.00 0.00 0.00 0.00 0.33 0.19 0.04 0.21 0.37 0.00 0.00 0.00 0.00 0.05 0.00 | 3137 3137 3137 3351 3351 3351 3351 3351 | 69 69 70 70 70 70 70 70 70 70 70 70 70 70 70 |
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| Liberal Age White Woman College Religious Non-Metro California Pennsylvania New York Florida Ilihoois Liberal Age White Woman College Religious | -0.01 -0.01 -0.08 0.05 -0.01 -0.02 -0.02 0.03 0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.05 -0.01 -0.02 -0.02 -0.02 -0.02 -0.03 | 0.01 0.01 0.00 0.02 0.00 0.00 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.02 0.00 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.00 0.00 0.00 0.00 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.00 | 0.26 0.42 0.00 0.00 0.00 0.00 0.00 0.00 0.03 0.19 0.04 0.58 0.21 0.37 0.00 0.00 0.00 0.00 0.00 0.00 0.00 | 3137 3137 3137 3351 3351 3351 3351 3351 | 69 69 70 70 70 70 70 70 70 70 70 70 70 70 70 |
| Liberal Age White Woman College Religious Non-Metro California Pennsylvania New York Florida Illinois Liberal Age White Woman College Religious Non-Metro | -0.01 -0.01 -0.08 0.05 -0.01 -0.02 -0.02 0.03 0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.02 -0.02 -0.02 -0.02 0.03 0.05 -0.01 -0.01 -0.02 -0.01 -0.02 -0.02 -0.01 -0.02 -0.02 -0.02 -0.01 -0.02 -0.01 -0.02 -0.00 | 0.01 0.01 0.00 0.02 0.00 0.00 0.00 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.02 0.01 0.02 0.01 0.02 0.00 | $\begin{array}{c} 0.26\\ 0.42\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.33\\ 0.19\\ 0.04\\ 0.58\\ 0.21\\ 0.37\\ 0.00\\$ | 3137 3137 3351 3351 3351 3351 3351 3351 | 69 69 70 70 70 70 70 70 70 70 70 70 70 70 70 |
| Liberal Age White Woman College Religious Non-Metro California Pennsylvania New York Florida Ilihoois Liberal Age White Woman College Religious Non-Metro College Religious | -0.01 -0.01 -0.08 0.05 -0.01 -0.02 -0.02 0.03 0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.05 -0.01 -0.02 -0.02 0.03 0.01 -0.02 -0.02 -0.02 0.03 -0.01 -0.01 -0.01 -0.02 -0.02 -0.01 -0.01 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.03 -0.01 -0.01 -0.01 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.01 -0.01 -0.01 -0.02 -0.02 -0.02 -0.01 -0.02 -0.02 -0.02 -0.02 -0.01 -0.01 -0.01 -0.02 -0.02 -0.02 -0.01 -0.01 -0.02 -0.02 -0.01 -0.02 -0.02 -0.01 -0.02 -0.01 -0.02 -0.02 -0.02 -0.02 -0.01 -0.02 -0.02 -0.02 -0.01 -0.02 -0.01 -0.02 -0.02 -0.02 -0.01 -0.02 | 0.01 0.01 0.02 0.00 0.00 0.00 0.00 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.02 0.00 | $\begin{array}{c} 0.26\\ 0.42\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.01\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.16\\ 0.20\\ \end{array}$ | 3137 3137 3351 3351 3351 3351 3351 3351 | 69 69 70 70 70 70 70 70 70 70 70 70 70 70 70 |
| Liberai Age White Woman College Religious Non-Metro California Pennsylvania New York New York Horida Illinois Liberal Liberal Minte Woman College White Woman College Religious Non-Metro California | -0.01 -0.01 -0.08 0.05 -0.01 -0.02 0.03 0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.08 0.05 -0.01 -0.02 -0.02 0.03 0.01 -0.02 -0.02 0.03 0.01 -0.02 -0.02 -0.03 0.01 -0.02 -0.03 -0.01 -0.01 -0.03 -0.01 -0.01 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.03 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.02 -0.02 -0.02 -0.03 -0.01 -0.02 -0.03 -0.01 -0.02 -0.03 -0.01 -0.02 -0.03 -0.01 -0.02 -0.03 -0.01 -0.02 -0.03 -0.01 -0.02 -0.03 -0.01 -0.02 -0.02 -0.02 -0.03 -0.01 -0.02 -0.02 -0.03 -0.01 -0.02 -0.02 -0.03 -0.01 -0.02 -0.03 -0.01 -0.02 -0.03 -0.01 -0.02 -0.03 -0.01 -0.02 -0.03 -0.01 -0.03 -0.01 -0.02 -0.03 -0.01 -0.03 -0.03 -0.01 -0.03 -0.03 -0.01 -0.03 -0.03 -0.01 -0.03 -0.0 | 0.01 0.01 0.02 0.00 0.00 0.00 0.00 0.01 0.01 0.01 0.01 0.01 0.01 0.02 0.00 | $\begin{array}{c} 0.26\\ 0.42\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.01\\ 0.00\\$ | 3137 3137 3137 3351 3351 3351 3351 3351 | 69 69 70 70 70 70 70 70 70 70 70 70 70 70 70 |
| Liberai Age White Woman College Religious Non-Metro California Pennsylvania New York Florida Ilinois Liberal Age White College Religious Non-Metro California Pennsylvania | -0.01 -0.01 -0.08 0.05 -0.01 -0.02 -0.02 0.03 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.02 -0 | 0.01 0.01 0.02 0.02 0.00 0.00 0.00 0.01 0.01 0.01 0.01 0.01 0.01 0.02 0.01 0.02 0.00 0.00 0.00 0.00 0.01 0.02 0.00 0.01 0.02 0.01 0.02 0.01 0.02 0.01 0.02 0.01 0.02 0.01 0.02 0.01 0.02 0.01 0.02 0.01 0.01 0.02 0.01 0.02 0.00 0.01 0.02 0.02 0.02 0.01 0.02 0.01 0.02 0.01 0.02 0.02 0.01 0.02 | $\begin{array}{c} 0.26\\ 0.42\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.33\\ 0.19\\ 0.33\\ 0.21\\ 0.37\\ 0.00\\ 0.04\\ 0.58\\ 0.21\\ 0.00\\$ | 3137 3137 3137 3351 3351 3351 3351 3351 | 69 69 70 70 70 70 70 70 70 70 70 70 70 70 70 |
| Liberai Age White Woman College Religious Nom-Metro California New York Florida Illinois Liberal Liberal Liberal Liberal College Religious Non-Metro California Religious Non-Metro California Religious Non-Metro California Pennsylvania New York Florida | -0.01 -0.01 -0.08 0.05 -0.01 -0.02 -0.02 -0.02 -0.02 -0.02 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.02 -0.02 -0.02 -0.02 -0.03 -0.01 -0.01 -0.01 -0.01 -0.01 -0.02 -0.01 -0.01 -0.01 -0.01 -0.02 - | 0.01 0.01 0.02 0.02 0.00 0.00 0.00 0.01 0.01 0.01 0.01 0.01 0.01 0.02 0.01 0.02 0.01 0.02 0.01 0.02 0.01 0.02 0.01 0.02 0.01 0.02 0.01 0.02 0.01 0.02 0.01 0.02 0.01 0.02 0.02 0.00 0.02 0.00 0.00 0.00 0.01 0.01 0.02 0.01 0.02 0.00 0.00 0.01 0.01 0.01 0.02 0.00 0.02 0.00 0.00 0.00 0.01 0.02 0.00 0.02 0.00 0.02 0.00 | $\begin{array}{c} 0.26\\ 0.42\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.33\\ 0.19\\ 0.33\\ 0.21\\ 0.37\\ 0.00\\ 0.04\\ 0.58\\ 0.21\\ 0.00\\$ | 3137 3137 3137 3351 3351 3351 3351 3351 | 69 69 70 70 70 70 70 70 70 70 70 70 70 70 70 |
| Liberai Age White Woman College Religious Non-Metro California Pennsylvania New York Florida Illinois Liberal Age White Woman College Religious Non-Metro California Pennsylvania New York Florida Illinois New York Florida | -0.01 -0.01 -0.08 0.05 -0.01 -0.02 -0.02 -0.02 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.02 -0.02 -0.02 -0.01 -0.01 -0.01 -0.02 -0.01 -0.01 -0.02 -0.02 -0.02 -0.02 -0.01 -0.01 -0.02 -0.02 -0.01 -0.01 -0.01 -0.02 -0.02 -0.02 -0.02 -0.01 -0.02 -0.02 -0.02 -0.02 -0.02 -0.01 -0.01 -0.01 -0.02 - | 0.01 0.01 0.02 0.00 0.00 0.00 0.00 0.01 0.01 0.01 0.01 0.01 0.01 0.02 0.00 0.00 0.01 0.01 0.02 0.00 0.00 0.01 0.01 0.02 0.01 0.01 0.02 0.00 0.01 0.01 0.02 0.01 0.02 0.01 0.02 0.01 0.01 0.01 0.02 0.01 0.02 0.01 0.02 0.01 0.02 0.01 0.02 0.01 0.02 0.01 0.02 0.01 0.01 0.02 0.01 0.02 0.01 0.02 0.01 0.02 0.01 0.01 0.02 0.01 0.02 0.01 0.01 0.01 0.02 0.00 0.01 0.01 0.01 0.01 0.02 0.00 0.01 0.01 0.01 0.01 0.02 0.00 0.00 0.00 0.01 0.00 0.01 | $\begin{array}{c} 0.26\\ 0.42\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.03\\ 0.19\\ 0.04\\ 0.58\\ 0.21\\ 0.37\\ 0.00\\$ | 3137 3137 3137 3351 3351 3351 3351 3351 | 69 69 70 70 70 70 70 70 70 70 70 70 70 70 70 |
| Liberai Age White Woman College Religious Non-Metro California New York Florida Illinois Liberal Liberal Age White Woman College Religious Non-Metro California College Religious Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro College Religious Non-Metro California Non-Metro California Non-Metro College Religious Non-Metro California Non-Metro California Non-Metro College Religious Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro California Non-Metro Non-Met | -0.01 -0.01 -0.08 0.05 -0.01 -0.02 0.03 0.01 -0.02 0.03 -0.01 -0.02 0.03 -0.01 -0.01 -0.02 -0.02 0.03 0.01 -0.02 0.03 0.01 -0.02 0.03 -0.01 -0.02 0.03 -0.01 -0.02 -0.02 -0.02 -0.01 -0.02 -0.02 -0.01 -0.02 -0.02 -0.01 -0.02 -0.02 -0.01 -0.02 -0.02 -0.02 -0.01 -0.02 -0.01 -0.02 -0.02 -0.02 -0.01 -0.02 -0.02 -0.02 -0.01 -0.02 -0.02 -0.01 -0.02 -0.01 -0.02 -0.02 -0.01 -0.02 -0.01 -0.02 -0.02 -0.01 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.01 -0.03 -0.01 -0.02 -0.02 -0.02 -0.01 -0.01 -0.02 -0.02 -0.01 -0.01 -0.02 -0.01 -0.02 -0.01 -0.01 -0.02 -0.02 -0.01 -0.01 -0.02 -0.01 -0.01 -0.02 -0.01 -0.02 -0.01 -0.01 -0.02 -0.01 -0.02 -0.01 -0.02 -0.01 -0.01 -0.02 -0.01 -0.01 -0.02 -0.01 -0.02 -0.01 -0.01 -0.02 -0.01 -0.02 -0.01 -0.01 -0.02 -0.01 -0.02 -0.01 -0.01 -0.02 -0.01 -0.01 -0.02 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.02 -0.01 | 0.01 0.01 0.02 0.00 0.00 0.00 0.00 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.02 0.00 0.00 0.00 0.00 0.00 0.01 0.01 0.01 0.01 0.01 0.01 0.02 0.00 | $\begin{array}{c} 0.26\\ 0.42\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.03\\ 0.19\\ 0.03\\ 0.19\\ 0.03\\ 0.21\\ 0.33\\ 0.21\\ 0.00\\$ | 3137 3137 3137 3351 3558 35988 3598 3598 3598 3598 35988 3598 3598 3598 35988 359 | 69 69 70 70 70 70 70 70 70 70 70 70 70 70 70 |
| Liberal Age White Woman College Religious Non-Metro California Pennsylvania New York Florida Ilihoral Age White Woman College Religious Non-Metro California Pennsylvania New York Florida Ilihoris Liberal Age White White White White | -0.01 -0.01 -0.08 0.05 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.02 -0.01 -0.02 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.02 - | 0.01 0.01 0.02 0.00 0.00 0.00 0.00 0.01 0.02 0.00 0.00 0.00 0.01 0.01 0.01 0.02 0.00 0.00 0.00 0.02 0.00 | $\begin{array}{c} 0.26\\ 0.42\\ 0.00\\$ | 3137 3137 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3598 3975 3975 | 69 69 70 70 70 70 70 70 70 70 70 70 70 70 70 |
| Liberal Age White Woman College Religious Non-Metro California New York Florida Illinois Liberal Age White Woman College Religious Non-Metro California Pennsylvania New York Florida Illinois Liberal Age White Woman | -0.01 -0.01 -0.08 0.05 -0.02 -0.02 -0.02 0.03 0.01 -0.01 -0.01 -0.01 -0.01 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.01 -0.01 -0.01 -0.01 -0.01 -0.02 -0.01 -0. | 0.01 0.01 0.00 0.00 0.00 0.00 0.00 0.00 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.02 0.00 | $\begin{array}{c} 0.26\\ 0.42\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.33\\ 0.19\\ 0.33\\ 0.19\\ 0.33\\ 0.19\\ 0.00\\$ | 3137 3137 3351 3598 35975 3975 3975 | 69 69 70 70 70 70 70 70 70 70 70 70 70 70 70 |
| Liberal Age White Woman College Religious Non-Metro California Pennsylvania New York Florida Ilihoria Liberal Age White College Religious Non-Metro California New York Florida Ilihoris Liberal Age Winte Woman College Usersylvania New York Florida | -0.01 -0.01 -0.08 0.05 -0.01 -0.02 -0.02 -0.02 -0.02 -0.02 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.02 -0.01 -0.02 -0.02 -0.02 -0.02 -0.02 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.01 - | 0.01 0.01 0.00 0.00 0.00 0.00 0.00 0.00 0.01 0.01 0.01 0.01 0.01 0.02 0.00 0.02 0.00 | $\begin{array}{c} 0.26\\ 0.42\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.33\\ 0.19\\ 0.04\\ 0.33\\ 0.19\\ 0.00\\ 0.00\\ 0.05\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.017\\ 0.17\\ 0.17\\ 0.44\\ 0.00$ | 3137 3137 3137 3351 3598 3598 3975 3975 3975 3975 | 69 69 70 70 70 70 70 70 70 70 70 70 70 70 70 |
| Liberai Age White Religious Religious Non-Metro California Non-Metro California Non-Metro California Kliberal Age White Religious Non-Metro California Religious Non-Metro California Non-Metro California Non-Metro California Non-Metro California Liberal Age White Woman College Religious College Religious Religious College Religious College Religious | -0.01 -0.02 -0.08 -0.01 -0.02 -0.02 -0.02 -0.02 -0.02 -0.03 -0.01 -0.01 -0.01 -0.01 -0.01 -0.02 -0.02 -0.02 -0.02 -0.02 -0.01 -0.01 -0.02 -0.02 -0.01 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.01 -0.02 | 0.01 0.01 0.01 0.00 0.02 0.00 0.00 0.00 0.00 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.02 0.00 | $\begin{array}{c} 0.266\\ 0.422\\ 0.000\\ 0.000\\ 0.000\\ 0.000\\ 0.000\\ 0.000\\ 0.000\\ 0.000\\ 0.001\\ 0.000\\ 0.$ | 3137 3137 3137 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3598 3975 3975 3975 3975 3975 | 699 699 700 700 700 700 700 700 700 700 700 7 |
| Luberai Age White College College College College College Pomsylvania New York Florida Illinois Liberal Liberal Liberal Age Woman College Religious Nor-Metro College Pennsylvania New York Florida Illinois Liberal Liberal Age Woman College Pennsylvania New York Florida Illinois Liberal Liberal Mensylvania New York Florida Illinois Liberal Mage White Woman College White College Religious Nor-Metro College Neiros New York Florida Illinois Liberal Mage New York Florida Illinois Liberal Mite College Religious Nor-Metro College Neiros New York Florida New York Florida New York Florida Illinois Liberal New York Florida New York Florida New York Neiros New York Neiros Neiros New York Neiros New York Neiros New York Neiros New York Neiros Neiros Neiros New York Neiros Neiro | -0.01 -0.01 -0.08 0.05 -0.01 -0.02 -0.02 0.03 0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.01 -0.02 -0.01 -0.02 -0. | 0.01 0.01 0.01 0.00 0.02 0.00 0.00 0.00 | $\begin{array}{c} 0.266\\ 0.422\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.01\\ 0.01\\ 0.00\\ 0.0$ | 3137 3137 3137 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 35975 3975 3975 3975 3975 3975 | 699 699 690 700 700 700 700 700 700 700 700 700 7 |
| Liberai Age White Religious Religious Religious Non-Metro California Non-Metro California Liberal Age White Religious Non-Metro California Religious Religious Non-Metro California Liberal Age White Religious Religious Religious California Collegue Religious California Collegue Religious California Collegue Religious California Collegue Religious California Collegue Religious California Collegue Religious Collionnia Collegue Collegue Religious Collegue Religious Collegue Religious Collegue Religious Collegue Religious Collegue Religious Collegue Religious Collegue Religious Collegue Religious Collegue Religious Collegue Religious Collegue Religious Collegue Religious Collegue Religious Collegue Religious Collegue Collegue Religious Collegue Religious Collegue Religious Collegue Collegue Religious Collegue Collegue Collegue Collegue Collegue Collegue Collegue Collegue Collegue Collegue Collegue Collegue Collegue C | -0.01 -0.08 0.05 -0.01 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.03 0.03 -0.01 -0.01 -0.01 -0.08 -0.01 -0.08 -0.01 -0.02 -0 | 0.011 0.011 0.010 0.022 0.000 0.000 0.001 0.011 0.011 0.011 0.011 0.022 0.001 0.011 0.010 0.020 0.000 0.000 0.000 0.000 0.001 0.011 0.011 0.010 0.022 0.000 0.001 0.010 0.022 0.001 0.010 0.010 0.010 0.010 0.0000 0.0000 0.0000 0.0000 0.0000 0.0000 0.0000 0.0000 0.0000 0.00000 0.00000 0.00000 0.000000 0.00000000 | $\begin{array}{c} 0.266\\ 0.422\\ 0.000\\ 0.$ | 3137 3137 3137 3351 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 35975 3975 3975 3975 3975 | 699 699 697 700 700 700 700 700 700 700 700 700 7 |
| Liberai Age Oblege Religious Nom Metro College Pennsylvania New York Florida Illinois Liberal Liberal Age Wönnan College Religious Nom-Metro College Religious Nom-Metro College Religious Nom-Metro Liberal Liberal Liberal Liberal Liberal New York Florida Illinois Liberal Liberal Mensylvania New York Florida Illinois Liberal Mensylvania New York Florida Unios Liberal New York Florida Unios Liberal New York Florida New York Florida | -0.01 -0.02 -0.08 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.03 -0.01 -0.01 -0.01 -0.01 -0.01 -0.02 -0.02 -0.02 -0.02 -0.02 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.02 | 0.01 0.01 0.00 0.02 0.00 0.00 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.01 0.02 0.00 0.00 0.00 0.01 0.02 0.00 | $\begin{array}{c} 0.266\\ 0.420\\ 0.420\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.33\\ 0.19\\ 0.04\\ 0.57\\ 0.00\\ 0.01\\ 0.00\\ 0.$ | 3137 3137 3137 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3598 3975 3975 3975 3975 | 69 99 69 70 70 70 70 70 70 70 70 70 70 70 70 70 |
| Laberal Mente Woman College Religious Non-Metro California Pennsylvania New York Florida Illinois Liberal Age White Woman College Religious Non-Metro California Liberal Age White Florida Liberal Age White Religious Non-Metro California College Religious College Religious College Religious College Religious College Religious College Religious College Religious College Religious College Religious College Religious College Religious College Religious College Religious College Religious College Religious Non-Metro College Religious Non-Metro Collifornia Pennsylvania Non-Metro | -0.01 -0.01 -0.08 0.05 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.01 -0.02 -0.01 -0.01 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.02 -0.01 -0.01 -0.02 - | 0.01 0.01 0.02 0.02 0.00 0.00 0.00 0.00 | $\begin{array}{c} 0.266\\ 0.422\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.00\\ 0.01\\ 0.01\\ 0.00\\ 0.0$ | 3137 3137 3137 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3351 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 3598 35975 3975 3975 3975 3975 3975 3975 3975 3975 3975 3975 | 699 697 700 700 700 700 700 700 700 700 700 7 |

C.7.9 Temporal Durability (Heterocentrism Outcome, Control Coefficients)

Table C43: Control Coefficients For Models Characterizing Temporal Durability of Post-Pulse Effect (Heterocentrism, With Controls, Part 1)



Table C44: Control Coefficients For Models Characterizing Temporal Durability of Post-Pulse Effect (Heterocentrism, With Controls, Part 2)



Table C45: Control Coefficients For Models Characterizing Temporal Durability of Post-Pulse Effect (Heterocentrism, With Controls, Part 3)



Table C46: Control Coefficients For Models Characterizing Temporal Durability of Post-Pulse Effect (Heterocentrism, With Controls, Part 4)



Table C47: Control Coefficients For Models Characterizing Temporal Durability of Post-Pulse Effect (Heterocentrism, With Controls, Part 5)



Table C48: Control Coefficients For Models Characterizing Temporal Durability of Post-
Pulse Effect (Heterocentrism, With Controls, Part 6)

| covar_name | est | se 0.01 | pv | nobs 2074 | days_cut |
|--|---------------------------------|------------------------------|------------------------------|----------------------|----------------|
| California | -0.01 | 0.01 | 0.38 | 2974 | 63 |
| Pennsylvania Now York | 0.02 | 0.01 | 0.09 | 2974 | 63 |
| Florida | -0.01 | 0.01 | 0.24 | 2974 2974 | 63 |
| Illinois | -0.00 | 0.01 | 0.85 | 2974 | 63 |
| Liberal Ame | -0.08 | 0.00 | 0.00 | 3092 | 64 |
| White | -0.01 | 0.00 | 0.20 | 3092 | 64 |
| Woman | -0.01 | 0.00 | 0.23 | 3092 | 64 |
| College Religions | -0.00 | 0.01 | 0.47 | 3092 | 64 64 |
| Non-Metro | 0.02 | 0.01 | 0.02 | 3092 | 64 |
| California | -0.01 | 0.01 | 0.46 | 3092 | 64 |
| Pennsylvania New York | 0.01 | 0.01 | 0.16 | 3092 | 64 |
| Florida | -0.00 | 0.01 | 0.58 | 3092 | 64 |
| Illinois | -0.00 | 0.01 | 0.72 | 3092 | 64 |
| Liberal | -0.08 | 0.00 | 0.00 | 3156 | 65 |
| White | -0.01 | 0.01 | 0.01 | 3156 | 65 |
| Woman | -0.00 | 0.00 | 0.32 | 3156 | 65 |
| College | -0.00 | 0.01 | 0.43 | 3156 | 65 |
| Non-Metro | 0.03 | 0.01 | 0.00 | 3156 | 65 |
| California | -0.00 | 0.01 | 0.76 | 3156 | 65 |
| Pennsylvania | 0.01 | 0.01 | 0.24 | 3156 | 65 |
| New York Florido | -0.01 | 0.01 | 0.27 | 3156 | 65 |
| Illinois | -0.00 | 0.01 | 0.84 | 3156 | 65 |
| Liberal | -0.08 | 0.00 | 0.00 | 3201 | 66 |
| Age | 0.03 | 0.01 | 0.02 | 3201 | 66 |
| White | -0.01 | 0.00 | 0.20 | 3201 | 66 66 |
| College | -0.00 | 0.00 | 0.37 | 3201 | 66 |
| Religious | 0.03 | 0.01 | 0.00 | 3201 | 66 |
| Non-Metro | 0.01 | 0.01 | 0.08 | 3201 | 66 |
| California Pennevlyania | -0.00 | 0.01 | 0.96 | 3201 | 66 66 |
| New York | -0.01 | 0.01 | 0.34 | 3201 | 66 |
| Florida | -0.00 | 0.01 | 0.65 | 3201 | 66 |
| Illinois Liborol | -0.01 | 0.01 | 0.59 | 3201 | 66 |
| Age | 0.03 | 0.00 | 0.00 | 3255 | 67 |
| White | -0.01 | 0.00 | 0.15 | 3255 | 67 |
| Woman | -0.01 | 0.00 | 0.21 | 3255 | 67 |
| College | -0.01 | 0.00 | 0.25 | 3255 | 67 |
| Non-Metro | 0.04 | 0.01 | 0.00 | 3255 | 67 |
| California | -0.00 | 0.01 | 0.67 | 3255 | 67 |
| Pennsylvania | 0.01 | 0.01 | 0.19 | 3255 | 67 |
| New York Florida | -0.01 | 0.01 | 0.30 | 3255 | 67 |
| Illinois | -0.01 | 0.01 | 0.50 | 3255 | 67 |
| Liberal | -0.08 | 0.00 | 0.00 | 3344 | 68 |
| Age White | 0.03 | 0.01 | 0.02 | 3344 | 68 |
| Woman | -0.01 | 0.00 | 0.21 | 3344 | 68 |
| College | -0.01 | 0.00 | 0.17 | 3344 | 68 |
| Religious | 0.04 | 0.00 | 0.00 | 3344 | 68 |
| Non-Metro California | 0.02 | 0.01 | 0.02 | 3344 | 68 |
| Pennsylvania | 0.01 | 0.01 | 0.44 | 3344 | 68 |
| New York | -0.01 | 0.01 | 0.43 | 3344 | 68 |
| Florida | -0.01 | 0.01 | 0.42 | 3344 | 68 |
| Liberal | -0.08 | 0.01 | 0.00 | 3408 | 69 |
| Age | 0.03 | 0.01 | 0.04 | 3408 | 69 |
| White | -0.01 | 0.00 | 0.17 | 3408 | 69 |
| College | -0.01 | 0.00 | 0.09 | 3408 | 69 |
| Religious | 0.01 | 0.00 | 0.00 | 3408 | 69 |
| Non-Metro | 0.02 | 0.01 | 0.03 | 3408 | 69 |
| California | -0.00 | 0.01 | 0.46 | 3408 | 69 |
| i ennsylvania New York | -0.01 | 0.01 | 0.36 | 3408 | 69 |
| Florida | -0.01 | 0.01 | 0.41 | 3408 | 69 |
| Illinois | -0.00 | 0.01 | 0.83 | 3408 | 69 |
| Liberal | -0.08 | 0.00 | 0.00 | 3587 | 70 |
| White | -0.01 | 0.01 | 0.07 | 3587 | 70 |
| Woman | -0.01 | 0.00 | 0.03 | 3587 | 70 |
| College | -0.01 | 0.00 | 0.14 | 3587 | 70 |
| neligious Non-Metro | 0.04 | 0.00 | 0.00 | 3587 3587 | 70 70 |
| California | -0.01 | 0.01 | 0.36 | 3587 | 70 |
| Pennsylvania | 0.01 | 0.01 | 0.47 | 3587 | 70 |
| New York | -0.01 | 0.01 | 0.21 | 3587 | 70 |
| Illinois | -0.01 | 0.01 | 0.21 | 3587 3587 | 70 70 |
| Liberal | -0.08 | 0.00 | 0.00 | 3776 | 71 |
| Age | 0.03 | 0.01 | 0.08 | 3776 | 71 |
| white Woman | -0.01 | 0.00 | 0.02 | 3776 | 71 |
| College | -0.01 | 0.00 | 0.14 | 3776 | 71 |
| Religious | 0.04 | 0.00 | 0.00 | 3776 | 71 |
| Non-Metro | 0.02 | 0.01 | 0.00 | 3776 | 71 |
| California Pennsylvania | -0.01 | 0.01 | 0.37 | 3776 | 71 |
| New York | -0.01 | 0.01 | 0.19 | 3776 | 71 |
| Florida | -0.01 | 0.01 | 0.17 | 3776 | 71 |
| Illinois | 0.01 | 0.01 | 0.66 | 3776 | 71 |
| Laberal | -0.08 | 0.00 | 0.00 | 4037 | 72 |
| White | -0.02 | 0.01 | 0.31 | 4037 | 72 |
| Woman | -0.01 | 0.00 | 0.02 | 4037 | 72 |
| College | -0.01 | 0.00 | 0.04 | 4037 | 72 |
| D.U.I. | 0.04 | 0.00 | 0.00 | 4037 4037 | 72 |
| Religious Non-Metro | 0.09 | | | 10 At \$ 1 | 14 |
| Religious Non-Metro California | 0.02 | 0.01 | 0.37 | 4037 | 72 |
| Religious Non-Metro California Pennsylvania | 0.02 -0.01 -0.01 | 0.01 0.01 0.01 | 0.37 0.52 | 4037 4037 | 72 72 |
| Religious Non-Metro California Pennsylvania New York | 0.02 -0.01 -0.01 -0.01 | 0.01 0.01 0.01 0.01 | 0.00 0.37 0.52 0.09 | 4037 4037 4037 | 72 72 72 |

C.8 Temporal Placebo Tests

Here, we show preexisting time trends are not driving our results. We estimate the influence of taking the PI S-IAT 15 and 20 days pre-Pulse relative to 16-30 and 21-40 days pre-Pulse on the *D*-score and heterocentrism. We also estimate the influence of taking the PI S-IAT after (2016-03-07 to 2016-06-11) relative to before (2016-01-01 to 2016-03-06) the median pre-treatment date. These placebo estimates are null, suggesting secular pro-gay time trends do not explain our findings (Figure C35).



Figure C35: Comparing True *post-Pulse* Coefficient to Placebo Coefficients To Rule Out Pre-Treatment Temporal Trends That Motivate Pro-Gay Attitudes. The x-axis is the type of estimate. True (15 days) is the true *post-Pulse* coefficient using a 15-day bandwidth. True (20 days) is the same with a 20-day bandwidth. Placebo 1 estimates the influence of taking the IAT in the 15 days prior to the Pulse massacre relative to the 16-30 days prior to the Pulse massacre. Placebo 2 estimates the influence of taking the IAT in the 20 days prior to the Pulse massacre relative to the 21-40 days prior to the Pulse massacre. Placebo 3 estimates the influence of taking the influence of taking the IAT after the median pre-treatment day (2016-03-07 to 2016-06-12) relative to the days before the median pre-treatment day (2016-01-01 to 2016-03-06). The y-axis is the coefficient. The left/right panel characterizes the influence of the true and placebo coefficients on the *D-score* and *heterocentrism*. 95% CIs displayed from HC2 robust SEs.
C.9 Prior and Post Year Temporal Placebo

Here, we attempt to rule out if systematic temporal trends near June motivate prosocial attitudes toward gay people other than the massacre. Thus, we assess the influence of placebo estimates comparing *D*-score and heterocentrism 15 and 20 days before and after June 12, the massacre calendar day, during the years 2010-2015 and 2017-2018. We find no consistent influence of these placebo estimates on the *D*-score and heterocentrism (Figure C36).



Figure C36: Temporal Placebo Tests Using IAT Data From Non-2016 Years. The x-axis is the IAT dataset at use (by year). The y-axis is the coefficient characterizing the influence of taking the IAT after June 12 (the calendar day of the Pulse nightclub shooting occurred). Panels A and B refer to estimates assessing the influence of the post-June 12th placebo on the *D-Score* and *Heterocentrism* outcomes. The top/bottom two panels are estimates using a 15/20 day bandwidth. 95% CIs displayed derived from HC2 robust standard errors.

C.10 Falsification Tests on Treatment-Irrelevant Group Attitudes

Here, we demonstrate our findings may not be due to a secular attitudinal trend in favor of marginalized groups through several falsification tests assessing if attitudes toward Black people, Asians, the differently-abled, Arabs, darker-skin people, and women shifts *post-Pulse* using the 15 and 20-day bandwidth samples.⁵⁷ Across 28 statistical tests, only 3 are significant (Section C.10), suggesting our findings are not driven by secular liberal attitudinal trends toward marginalized groups.

| Post-Pulse Coef. | \mathbf{SE} | \mathbf{p} | Ν | Outcome | Dataset | Bandwidth |
|------------------|---------------|--------------|-----------|---------------|--------------------------|-----------|
| -0.000 | 0.005 | 0.949 | 11310.000 | D-Score | Black/White IAT | 15 days |
| -0.003 | 0.003 | 0.377 | 10960.000 | White Bias | Black/White IAT | 15 days |
| -0.006 | 0.003 | 0.043 | 11039.000 | Ethnocentrism | Black/White IAT | 15 days |
| 0.012 | 0.015 | 0.434 | 1279.000 | D-Score | Asian/European IAT | 15 days |
| 0.011 | 0.011 | 0.320 | 1234.000 | White Bias | Asian/European IAT | 15 days |
| 0.006 | 0.014 | 0.670 | 1509.000 | D-Score | Disabled/Abled IAT | 15 days |
| -0.002 | 0.008 | 0.765 | 1484.000 | Abled Bias | Disabled/Abled IAT | 15 days |
| -0.009 | 0.009 | 0.319 | 1500.000 | Abledcentrism | Disabled/Abled IAT | 15 days |
| -0.013 | 0.013 | 0.327 | 1331.000 | D-Score | Arab/Non-Arab IAT | 15 days |
| -0.003 | 0.009 | 0.766 | 1267.000 | Non-Arab Bias | Arab/Non-Arab IAT | 15 days |
| -0.002 | 0.010 | 0.808 | 1310.000 | Ethnocentrism | Arab/Non-Arab IAT | 15 days |
| -0.014 | 0.009 | 0.145 | 3064.000 | D-Score | Dark Skin/Light Skin IAT | 15 days |
| -0.001 | 0.007 | 0.898 | 4550.000 | D-Score | Man/Woman (Career) IAT | 15 days |
| 0.004 | 0.010 | 0.702 | 2339.000 | D-Score | Man/Woman (Science) IAT | 15 days |
| -0.003 | 0.004 | 0.429 | 15506.000 | D-Score | Black/White IAT | 20 days |
| -0.006 | 0.003 | 0.013 | 15037.000 | White Bias | Black/White IAT | 20 days |
| -0.008 | 0.003 | 0.004 | 15151.000 | Ethnocentrism | Black/White IAT | 20 days |
| 0.008 | 0.013 | 0.518 | 1735.000 | D-Score | Asian/European IAT | 20 days |
| 0.011 | 0.009 | 0.218 | 1670.000 | White Bias | Asian/European IAT | 20 days |
| 0.010 | 0.012 | 0.399 | 1972.000 | D-Score | Disabled/Abled IAT | 20 days |
| 0.005 | 0.007 | 0.481 | 1938.000 | Abled Bias | Disabled/Abled IAT | 20 days |
| -0.003 | 0.008 | 0.736 | 1959.000 | Abledcentrism | Disabled/Abled IAT | 20 days |
| 0.005 | 0.012 | 0.638 | 1745.000 | D-Score | Arab/Non-Arab IAT | 20 days |
| 0.005 | 0.008 | 0.532 | 1663.000 | Non-Arab Bias | Arab/Non-Arab IAT | 20 days |
| 0.005 | 0.009 | 0.543 | 1717.000 | Ethnocentrism | Arab/Non-Arab IAT | 20 days |
| -0.009 | 0.008 | 0.249 | 4213.000 | D-Score | Dark Skin/Light Skin IAT | 20 days |
| -0.003 | 0.006 | 0.604 | 6624.000 | D-Score | Man/Woman (Career) IAT | 20 days |
| 0.007 | 0.008 | 0.416 | 3371.000 | D-Score | Man/Woman (Science) IAT | 20 days |

Table C49: Falsification Test on Treatment-Irrelevant Group Attitudes

This table characterizes falsification tests assessing the influence of taking an IAT *post-Pulse* on groups that are potentially unrelated to LGBTQ+. Not all datasets include the respective *D-score*, *bias*, and dominant group-centrism outcomes (hence their missingness in some IAT datasets). HC2 robust SEs displayed.

⁵⁷Falsification test data comes from separate Project Implicit surveys co-currently available to take in addition to the anti-gay attitude survey.

| | D S | core | Heteroo | entrism | D S | core | Heteroo | entrism | D S | core | Heteroo | entrism |
|------------------------|---------------------|---------------------|---------------------|---------------------|---------------------|---------------------|---------------------|---------------------|---------------------|---------------------|---------------------|---------------------|
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) | (11) | (12) |
| Post-Pulse | -0.01^{\dagger} | -0.01^{\dagger} | -0.02^{**} | -0.01^{**} | -0.02^{*} | -0.01 | -0.02^{**} | -0.02^{**} | -0.01^{*} | -0.00 | -0.01 | -0.01 |
| | (0.01) | (0.00) | (0.00) | (0.00) | (0.01) | (0.01) | (0.01) | (0.01) | (0.01) | (0.01) | (0.01) | (0.01) |
| Post-Pulse x Non-White | 0.01 | 0.00 | 0.01 | 0.00 | | | | | | | | |
| | (0.01) | (0.01) | (0.01) | (0.01) | | | | | | | | |
| Post-Pulse x Woman | | | | | 0.02 | 0.01 | 0.02^{\dagger} | 0.01 | | | | |
| | | | | | (0.01) | (0.01) | (0.01) | (0.01) | | | | |
| Post-Pulse x Liberal | | | | | | | | | 0.01 | -0.00 | -0.00 | -0.00 |
| | | | | | | | | | (0.01) | (0.01) | (0.01) | (0.01) |
| Non-White | 0.02^{*} | 0.01^{*} | 0.00 | 0.01 | | | | | | | | |
| | (0.01) | (0.01) | (0.01) | (0.01) | | | | | | | | |
| Woman | -0.02^{***} | -0.02^{***} | -0.01 | -0.01^{*} | -0.03^{***} | -0.02^{***} | -0.01^{*} | -0.01^{*} | -0.02^{***} | -0.02^{***} | -0.01 | -0.01^{*} |
| | (0.00) | (0.00) | (0.00) | (0.00) | (0.01) | (0.01) | (0.01) | (0.01) | (0.00) | (0.00) | (0.00) | (0.00) |
| Liberal | -0.07^{***} | -0.07^{***} | -0.08^{***} | -0.08^{***} | -0.07^{***} | -0.07^{***} | -0.08^{***} | -0.08^{***} | -0.07^{***} | -0.07^{***} | -0.08^{***} | -0.08^{***} |
| | (0.00) | (0.00) | (0.00) | (0.00) | (0.00) | (0.00) | (0.00) | (0.00) | (0.01) | (0.01) | (0.01) | (0.01) |
| Bandwidth | $15 \mathrm{ Days}$ | $20 \mathrm{ Days}$ | $15 \mathrm{ Days}$ | $20 \mathrm{~Days}$ | $15 \mathrm{~Days}$ | $20 \mathrm{~Days}$ | $15 \mathrm{ Days}$ | $20 \mathrm{~Days}$ | $15 \mathrm{ Days}$ | $20 \mathrm{~Days}$ | $15 \mathrm{~Days}$ | $20 \mathrm{ Days}$ |
| \mathbb{R}^2 | 0.12 | 0.12 | 0.16 | 0.16 | 0.12 | 0.12 | 0.17 | 0.16 | 0.12 | 0.12 | 0.16 | 0.16 |
| N | 3638 | 4907 | 3645 | 4920 | 3638 | 4907 | 3645 | 4920 | 3638 | 4907 | 3645 | 4920 |

Table C50: Assessing Heterogenous Influence of *Post-Pulse* (Study 2, Part 1)

Note: ***p < 0.001; **p < 0.01; *p < 0.05. All models adjust for age, white (when not assessing heterogeneity by non-white), woman, college education, religious, metropolitan residence, ideology, California, Pennsylvania, Florida, and Illinois state residence. All covariates rescaled between 0-1. HC2 robust SEs in parentheses.

Table C51: Assessing Heterogenous Influence of Post-Pulse (Study 2, Part 2)

| | DO | | TT / | | DO | | TT / | |
|---|----------|---------|-------------|----------|----------|-------------------|---------------------|--------------|
| | DS | core | Heteroo | centrism | DS | core | Heteroo | centrism |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Post-Pulse | 0.03 | 0.01 | 0.01 | -0.01 | -0.00 | -0.00 | -0.02^{*} | -0.02^{**} |
| | (0.02) | (0.01) | (0.03) | (0.03) | (0.01) | (0.01) | (0.01) | (0.01) |
| Post-Pulse x % LGBT (State) | -0.07 | -0.04 | -0.05 | -0.00 | | | | |
| | (0.03) | (0.02) | (0.07) | (0.07) | | | | |
| Post-Pulse x SS Couple Density (County) | | | | | -0.02 | -0.02 | 0.05 | 0.04 |
| | | | | | (0.03) | (0.03) | (0.03) | (0.02) |
| % LGBT (State) | -0.01 | -0.02 | -0.00 | -0.02 | | | | |
| | (0.03) | (0.02) | (0.05) | (0.04) | | | | |
| SS Couple Density (County) | | | | | -0.03 | -0.04^{\dagger} | -0.09^{**} | -0.08^{**} |
| | | | | | (0.03) | (0.02) | (0.02) | (0.02) |
| Bandwidth | 15 Days | 20 Days | 15 Days | 20 Days | 15 Days | 20 Days | $15 \mathrm{~Days}$ | 20 Days |
| R^2 | 0.12 | 0.12 | 0.17 | 0.16 | 0.12 | 0.12 | 0.17 | 0.16 |
| Ν | 3638 | 4907 | 3645 | 4920 | 3638 | 4907 | 3645 | 4920 |
| N Clusters | 52 | 52 | 52 | 52 | 739 | 848 | 738 | 848 |

Note: ***p < 0.001; **p < 0.01; *p < 0.05. All models adjust for age, white, woman, college education, religious, metropolitan residence, ideology, California, Pennsylvania, Florida, and Illinois state residence. Models 1-4 adjust for an interaction between *post-pulse* and an indicator for state residence missingness. Models 5-8 adjust for an interaction between *post-pulse* and an indicator for county residence missingness. All covariates rescaled between 0-1. HC2 robust SEs in parentheses but clustered at state and county-level for Models 1-4 and 5-8 respectively.

C.12 Balance Tests After Removing Days After Pulse Event

 Table C52: Covariate Balance Tests After Cutting Days Immediately After Pulse

 Massacre

| Days Cut | # Imbalanced Covariates | Imbalanced Covariates |
|----------|-------------------------|---|
| 1 | 0/12 | |
| 2 | 1/12 | Florida |
| 3 | 2/12 | White, Florida |
| 4 | 3/12 | Liberal, White, Florida |
| 5 | 2/12 | Liberal, White |
| 6 | 1/12 | White |
| <i>(</i> | 2/12 | White, Woman White, California |
| ° 9 | 2/12 5/12 | White California New York Florida Illinois |
| 10 | 3/12 | California, New York, Florida, Inniois |
| 10 | 3/12 | California, New York, Illinois |
| 12 | 3/12 | California, New York, Illinois |
| 13 | 4/12 | Liberal, California, New York, Illinois |
| 14 | 4/12 | Liberal, College, New York, Illinois |
| 15 | 7/12 | Liberal, Age, College, Religious, New York, Florida, Illinois |
| 16 | 6/12 | Liberal, Age, College, New York, Florida, Illinois |
| 17 | 5/12 | Liberal, Age, College, New York, Illinois |
| 18 | 4/12 | Liberal, Age, College, Illinois |
| 19 | 5/12 6/12 | Liberal Age, Woman College, Inniois |
| 20 | 5/12 | Liberal Age Woman College Florida |
| 22 | 5/12 | Liberal, Age, Woman, College, Florida |
| 23 | 6/12 | Liberal, Age, Woman, College, Religious, Florida |
| 24 | 4/12 | Liberal, Woman, College, Florida |
| 25 | 6/12 | Liberal, Age, Woman, College, Religious, Florida |
| 26 | 6/12 | Liberal, Age, Woman, College, Religious, Florida |
| 27 | 6/12 | Liberal, Woman, College, Religious, Non-metro, Florida |
| 28 | 5/12 | Liberal, Woman, College, Non-metro, Florida |
| 29 | 5/12 | Liberal, Woman, College, Non-metro, Florida |
| 30 | 5/12 6/12 | Liberal, Woman, College, Non-Inetro, Florida |
| 32 | 5/12 | Liberal, Woman, College, Religious, Florida |
| 33 | 5/12 | Liberal, Woman, College, Religious, Florida |
| 34 | 5/12 | Liberal, Woman, College, Religious, Florida |
| 35 | 5/12 | Liberal, Woman, College, Religious, Florida |
| 36 | 5/12 | Liberal, Woman, College, Religious, Florida |
| 37 | 5/12 | Liberal, Woman, College, Religious, Florida |
| 38 | 5/12 | Liberal, Woman, College, Religious, Florida |
| 39 | 6/12 | Liberal, Woman, College, Religious, Florida, Illinois Liberal, Woman, College, Beligious, Florida, Illinois |
| 40 | 6/12 | Liberal, Woman, College, Religious, Florida, Illinois |
| 41 42 | 6/12 | Liberal, Woman, College, Religious, Florida, Illinois |
| 43 | 6/12 | Liberal, Woman, College, Religious, Florida, Illinois |
| 44 | 6/12 | Liberal, Woman, College, Religious, Florida, Illinois |
| 45 | 6/12 | Liberal, Woman, College, Religious, Florida, Illinois |
| 46 | 5/12 | Liberal, Woman, College, Florida, Illinois |
| 47 | 5/12 | Liberal, Woman, College, Florida, Illinois |
| 48 | 5/12 | Liberal, Woman, College, Florida, Illinois |
| 49 | 5/12 | Liberal, Woman, College, Florida, Illinois |
| 50 | 5/12 6/12 | Liberal, Woman, College, Florida, Illinois Liberal, Woman, College, Popperlyania, Florida, Illinois |
| 52 | 6/12 | Liberal Woman College Peppsylvania Florida Illinois |
| 53 | 6/12 | Liberal, Woman, College, Pennsylvania, Florida, Illinois |
| 54 | 8/12 | Liberal, Age, Woman, College, Pennsylvania, New York, Florida, Illinois |
| 55 | 7/12 | Liberal, Woman, College, Pennsylvania, New York, Florida, Illinois |
| 56 | 7/12 | Liberal, Woman, College, Pennsylvania, New York, Florida, Illinois |
| 57 | 7/12 | Liberal, Woman, College, Pennsylvania, New York, Florida, Illinois |
| 58 | 7/12 | Liberal, Woman, College, Pennsylvania, New York, Florida, Illinois |
| 59 | 8/12 | Liberal, Woman, College, Religious, Pennsylvania, New York, Florida, Illinois |
| 60 61 | 7/12 | Liberal, Woman, College, Pennsylvania, New York, Florida, Illinois |
| 01 69 | 0/12 6/12 | Liberal, Woman, College, Pennsylvania, Florida, Illinois Liberal, Woman, College, Pennsylvania, New York, Illinois |
| 63 | 4/19 | Liberal Woman Colloge Illinois |
| 64 | 6/12 | Liberal, Woman, College, Non-metro, New York, Illinois |
| 65 | 5/12 | Liberal, Age, Woman. College. Illinois |
| 66 | 6/12 | Liberal, Age, Woman, College, New York, Florida |
| 67 | 5/12 | Liberal, Age, Woman, College, Florida |
| 68 | 4/12 | Age, Woman, College, Florida |
| 69 | 4/12 | Age, College, Pennsylvania, Florida |
| 70 | 5/12 | Age, White, College, New York, Florida |
| 71 | 5/12 | Age, White, New York, Florida, Illinois |
| 72 | 7/12 | Age, White, College, Keligious, New York, Florida, Illinois |

C.13 Sorting Test

Given respondents self-select into the S-IAT, we may be concerned systematic sorting induces bias (e.g. pro-gay people taking the survey *post-Pulse*). We do not believe sorting is a concern. If more pro-gay individuals were taking the survey *post-Pulse*, *post-Pulse* respondents would be younger, more liberal, less religious, and more college-educated, but they are not (Figure C34, Panels C-D). Second, if sorting were operative, we may expect more respondents taking the S-IAT *post-Pulse*. We conduct a difference-in-means comparing the number of daily respondents *post-Pulse* relative to pre-Pulse, and do not statistically find more respondents took the S-IAT *post-Pulse* (Table C53).

| Table C53: | Effect | of Pulse | On Nun | nber of | Project | Implicit | Sexuality | \mathbf{IAT} | Survey |
|-------------|--------------|----------|--------|---------|---------|----------|-----------|----------------|--------|
| Participant | \mathbf{s} | | | | | | | | |

| | # Of Pa | rticipants |
|----------------|-----------|------------|
| | (1) | (2) |
| Intercept | 111.47*** | 118.60*** |
| | (12.97) | (10.76) |
| Post-Pulse | 22.00 | 10.60 |
| | (17.21) | (14.09) |
| Bandwidth | 15-day | 20-day |
| \mathbb{R}^2 | 0.06 | 0.01 |
| Num. obs. | 30 | 40 |

***p < 0.001; ** p < 0.01; * p < 0.05

C.14 Heterogeneity By Conservatism

C.14.1 15, 20-Day Bandwidth Sample

| Table C54: | Heterogenous | Effect of Post- | Pulse on Ant | ti-Gay Attit | udes Condti | onal |
|------------|--------------|-----------------|--------------|--------------|-------------|------|
| On Conserv | vatism | | | | | |

| | D-S | core | Heteroo | entrism |
|---------------------------|---------------|---------------|---------------|---------------|
| | (1) | (2) | (3) | (4) |
| Post-Pulse x Conservative | 0.00 | 0.01 | 0.01 | 0.01 |
| | (0.01) | (0.01) | (0.01) | (0.01) |
| Post-Pulse | -0.01 | -0.01^{*} | -0.01^{***} | -0.01^{***} |
| | (0.01) | (0.00) | (0.00) | (0.00) |
| Conservative | 0.03^{**} | 0.02** | 0.05*** | 0.05*** |
| | (0.01) | (0.01) | (0.01) | (0.01) |
| Age | 0.06*** | 0.06*** | 0.04^{**} | 0.04** |
| | (0.02) | (0.01) | (0.01) | (0.01) |
| White | -0.02^{***} | -0.02^{***} | -0.01^{**} | -0.01^{***} |
| | (0.00) | (0.00) | (0.00) | (0.00) |
| Woman | -0.02^{***} | -0.02^{***} | -0.00 | -0.01 |
| | (0.00) | (0.00) | (0.00) | (0.00) |
| College-Educated | -0.02^{**} | -0.01^{***} | -0.01^{*} | -0.01 |
| | (0.01) | (0.00) | (0.00) | (0.00) |
| Religious | 0.03*** | 0.03*** | 0.04*** | 0.04*** |
| | (0.01) | (0.00) | (0.00) | (0.00) |
| Non-Metro | 0.01 | 0.00 | 0.01 | -0.00 |
| | (0.01) | (0.01) | (0.01) | (0.01) |
| Liberal | -0.06^{***} | -0.06^{***} | -0.06^{***} | -0.06^{***} |
| | (0.01) | (0.00) | (0.00) | (0.00) |
| California | -0.02^{*} | -0.01 | -0.00 | -0.01 |
| | (0.01) | (0.01) | (0.01) | (0.01) |
| Pennsylvania | 0.02 | 0.02 | 0.02 | 0.01 |
| | (0.01) | (0.01) | (0.01) | (0.01) |
| New York | -0.01 | -0.01 | 0.01 | -0.00 |
| | (0.01) | (0.01) | (0.01) | (0.01) |
| Florida | 0.02 | 0.01 | 0.00 | 0.00 |
| | (0.01) | (0.01) | (0.01) | (0.01) |
| Illinois | 0.00 | -0.00 | -0.01 | -0.01 |
| | (0.01) | (0.01) | (0.01) | (0.01) |
| Bandwidth (in Days) | 15 | 20 | 15 | 20 |
| \mathbb{R}^2 | 0.12 | 0.12 | 0.18 | 0.18 |
| Num. obs. | 3638 | 4907 | 3645 | 4920 |

***p < 0.001; **p < 0.01; *p < 0.05

C.14.2 15, 20-Day Bandwidth Sample, Comparing Pre-Pulse Respondents To Respondents Interviewed in 15 Days After 72 Days Post-Pulse

Table C55: Heterogenous Effect of *Post-Pulse* on Anti-Gay Attitudes Condtional On Conservatism (Comparing Respondents Interviewed 15 Days Before The Pulse Massacre To Those Interviewed 15 Days After 72 Days From the Pulse Massacre)

| | D-Score | Heterocentrism |
|---------------------------|---------------|----------------|
| | (1) | (2) |
| Post-Pulse x Conservative | 0.01 | -0.00 |
| | (0.01) | (0.01) |
| Post-Pulse | -0.01 | -0.00 |
| | (0.01) | (0.00) |
| Conservative | 0.02^{*} | 0.05^{***} |
| | (0.01) | (0.01) |
| Age | 0.05^{**} | 0.02 |
| | (0.02) | (0.01) |
| White | -0.01^{*} | -0.01^{*} |
| | (0.00) | (0.00) |
| Woman | -0.02^{***} | -0.01 |
| | (0.00) | (0.00) |
| College-Educated | -0.02^{***} | -0.01^{*} |
| | (0.01) | (0.00) |
| Religious | 0.03*** | 0.03*** |
| | (0.01) | (0.00) |
| Non-Metro | 0.01 | 0.02** |
| | (0.01) | (0.01) |
| Liberal | -0.07^{***} | -0.06^{***} |
| | (0.01) | (0.00) |
| California | -0.01 | -0.01 |
| | (0.01) | (0.01) |
| Pennsylvania | 0.03^{*} | 0.01 |
| | (0.02) | (0.01) |
| New York | -0.01 | -0.01 |
| | (0.01) | (0.01) |
| Florida | -0.01 | -0.01 |
| | (0.01) | (0.01) |
| Illinois | -0.01 | 0.00 |
| | (0.01) | (0.01) |
| Bandwidth (in Days) | 15 | 15 |
| \mathbb{R}^2 | 0.14 | 0.17 |
| Num. obs. | 3351 | 3587 |

***p < 0.001; **p < 0.01; *p < 0.05

C.15 Heterogeneity By Predicted Outcomes Assuming No Exposure to Pulse

Table C56: Assessing The Heterogenous Effect of Post-Pulse Conditional OnPredicted D-Score and Heterocentrism If Post-Pulse Indicator Is Equal To 0

| | D-Score | Heterocentrism | D-Score | Heterocentrism |
|---|--------------|----------------|--------------|----------------|
| | (1) | (2) | (3) | (4) |
| Post-Pulse x D-Score (Predicted) | 0.01 | | 0.07 | |
| | (0.09) | | (0.08) | |
| Post-Pulse x Heterocentrism (Predicted) | | 0.04 | | 0.04 |
| | | (0.08) | | (0.07) |
| Post-Pulse | -0.02 | -0.03 | -0.05 | -0.03 |
| | (0.05) | (0.04) | (0.04) | (0.03) |
| D-Score | 0.99^{***} | | 0.96^{***} | |
| | (0.07) | | (0.06) | |
| SSM Support (Predicted) | | 0.98^{***} | | 0.98^{***} |
| | | (0.06) | | (0.05) |
| Bandwidth (in Days) | 15 | 15 | 20 | 20 |
| Controls? | Ν | Ν | Ν | Ν |
| \mathbb{R}^2 | 0.12 | 0.16 | 0.12 | 0.16 |
| Num. obs. | 3638 | 3645 | 4907 | 4920 |

****p < 0.001; ***p < 0.01; *p < 0.05; †p < 0.1

D Study 3: Matthew Shepard

D.1 Media Data Details

We collect data on the number of gay-related newspaper articles in the New York Times and Washington Post. Data on the number of gay-related NYT newspaper articles per month are from the NYT article API. We use the **rtimes** package to query data from the NYT article API. Gay-related NYT newspaper articles include the terms "homosexual" or "gay" in their text (Figure 7, Panel A). Shepard-related articles are gay-related NYT newspaper articles with the terms "wyoming," "shepard," "student," "laramie," "beat," "beaten," "bias," and "hate (Figure 7, Panel C)."

Data on the number of gay-related Washington Post articles per month are acquired from the ProQuest Washington Post historical newspaper database (Figure 7, Panel B). Gayrelated articles are those that include the terms "homosexual," "gay," or "homosexuality" in their text.

D.2 Homosexuality = Morally Wrong Outcome

D.2.1 Study Details

The two studies we use to assess if the belief homosexuality is immoral decreased after Shepard's murder are the CNN/USA Today Jun 22-23 1998 poll and CNN/TIME Oct 14-15, 1998 poll. Both are nationally representative adult telephone surveys (N = 1016, N = 1036) and are population weighted to census demographic benchmarks. The CNN/USA Today poll was implemented by The Gallup Organization, and the CNN/TIME poll was implemented by Yankelovic Partners, Inc.

These two polls have largely similar sampling strategies (Voss et al., 1995). They are random digit dialing telephone polls. Their lists come from the same sample provider despite being implemented by different organizations (Survey Sampling Inc, SSI). Phone numbers are randomly selected based on a county's contribution to the total number of telephone households (e.g. if a county contains 20% of the national population, a telephone number will be randomly selected from that county with a probability of 20%). The key difference between the two organizations is weighting. Gallup weighs their data to Census statistics along sex, race, Census region, age and education. Yankelovic weighs their data to Census statistics along Census region, sex, race, education, and marital status. Therefore, Gallup weighs on age unlike Yankelovic, but Yankelovic weighs on marital status unlike Gallup. We do not believe differences in weighting generate a significant problem for inference. First, across 20 baseline covariates, only 2/20 are statistically imbalanced between the two samples, suggesting the Yankelovic and Gallup sampling strategies produce relatively similar samples despite weighting differences (Figure 8, Panel A). Second, after adjusting for imbalance between the two surveys, the *post-Shepard* coefficient is stable, suggesting sampling/compositional differences may not serve as a strong source of bias (Figure 8, Panel B).

The outcome item of interest from the CNN/USA Today Jun 1998 poll is "do you personally believe homosexual behavior is morally wrong or is not morally wrong" with response choices of 1) Yes, morally wrong and 2) No, not morally wrong. The outcome is binary, equal to 1 if the respondent indicates "Yes, morally wrong." The outcome item of interest from the CNN/TIME Oct 1998 poll is "do you personally think that homosexual relationships between consenting adults is morally wrong, or not a moral issue?" with response choices of 1) Yes, morally wrong and 2) Not a moral issue. The outcome is binary, equal to 1 if the respondent indicates "Yes, morally wrong." The weights, outcome, and baseline covariates are then stacked amongst each other across the two polls, with respondents from the CNN Oct. 1998 poll being defined as *post-Shepard* respondents (measured as a binary indicator equal to 1 if the respondent is from the October 1998 poll, 0 otherwise) and respondents from the CNN Jun. 1998 poll being defined as pre-Shepard respondents.

Although a benefit of these outcome items across the two surveys is they ask about the immorality of homosexuality very closely to the moment Matthew Shepard was murdered. they are worded slightly different from one another in that the post-Shepard survey references "homosexual relationships between consenting adults" while the pre-Shepard survey references "homosexual behavior." Therefore, it is plausible the decrease in support for the belief homosexuality is morally wrong may be a function of the specification that the homosexual behavior referenced in the post-Shepard survey relates specifically to behavior among consenting adults. Consequently, we re-estimate our findings with a different pre-Shepard survey from 1994 with a similar item wording. Consistent with the main findings, we find that respondents interviewed after Shepard's murder were less likely to believe homosexual relationships between consenting adults is morally wrong (see Figure 9). A shortcoming of the re-estimation is that our findings may be the result of secular time trends or intervening factors outside Shepard's murder. These alternative explanations are unlikely. First, the temporal placebo test comparing attitudes regarding "homosexual behavior" between April 1997-June 1998 on Figure 8, Panel B is statistically null. These findings suggest attitudes regarding the immorality of homosexuality were not trending in a liberal direction between 1994 to 1998 prior to Shepard's murder. Second, belief in the notion that "homosexual relationships between consenting adults" are "morally wrong" is *remarkably stable* between 1978-2004, with the exception of respondents interviewed in the few days after Matthew Shepard was murdered (Figure 9). These empirical findings suggest that item wording does not drive our main results and that Shepard's murder shifted anti-gay attitudes and not other temporal intervening factors.

D.2.2 Temporal Placebo Details

To conduct a temporal placebo test ruling out secular trends that may drive our finding that respondents interviewed after Shepard's murder were less likely to to believe homosexuality is morally wrong, we use a third survey, the Gallup Apr 11-13 1997 poll. The Gallup Apr 1997 poll is a nationally representative telephone survey (N = 1003) and is population weighted to census demographic benchmarks. The Gallup Apr 1997 poll includes an item asking respondents if they "personally believe homosexual behavior is morally wrong or is not morally wrong" with responses 1) Yes, morally wrong and 2) No, not morally wrong. We then compare the average level of support for belief homosexual behavior is morally wrong between the Gallup Apr. 1997 poll and CNN Jun. 1998 poll.

D.3 Regression Tables

D.3.1 Balance Test (Moral Wrong)

| Outcome | Post-Shepard Coef. | \mathbf{SE} | \mathbf{p} | \mathbf{N} |
|-------------------|--------------------|---------------|--------------|--------------|
| White | -0.08 | 0.02 | 0.00 | 2052 |
| Woman | -0.01 | 0.02 | 0.79 | 2052 |
| College | 0.02 | 0.02 | 0.26 | 2052 |
| Age (18-24) | 0.01 | 0.02 | 0.48 | 2052 |
| Age (25-29) | 0.01 | 0.02 | 0.63 | 2052 |
| Age (30-34) | -0.02 | 0.01 | 0.14 | 2052 |
| Age (35-39) | -0.02 | 0.01 | 0.30 | 2052 |
| Age (40-49) | 0.00 | 0.02 | 0.98 | 2052 |
| Age (50-64) | 0.01 | 0.02 | 0.73 | 2052 |
| Age $(65+)$ | -0.02 | 0.02 | 0.27 | 2052 |
| Income $(20-50k)$ | -0.02 | 0.02 | 0.52 | 2052 |
| Income (50-75k) | 0.00 | 0.02 | 0.84 | 2052 |
| Income (75k+) | -0.05 | 0.01 | 0.00 | 2052 |
| Democrat | 0.02 | 0.02 | 0.44 | 2052 |
| Registered | -0.02 | 0.02 | 0.45 | 2052 |
| Texas | 0.00 | 0.01 | 0.98 | 2052 |
| California | -0.01 | 0.02 | 0.37 | 2052 |
| New York | 0.01 | 0.01 | 0.60 | 2052 |
| Florida | 0.01 | 0.01 | 0.49 | 2052 |
| Pennsylvania | 0.01 | 0.01 | 0.51 | 2052 |

Table D57: Post-Shepard Balance Test (Moral Wrong, CNN Jun '98/Oct '98)

D.3.2 Temporal Placebo Test (Moral Wrong)

| | Moral Wrong (1) |
|--------------|--|
| Post-Placebo | -0.00 (0.02) |
| R^2 N | $\begin{array}{c} 0.00\\ 2019 \end{array}$ |
| Surveys | Gallup Apr '97/CNN Jun '98 |

| Table D | 58: Tem | poral P | lacebo | Tests |
|---------|---------|---------|--------|-------|
|---------|---------|---------|--------|-------|

Note: ****p < 0.001, ** p < 0.01, * p < 0.05, $^\dagger p < 0.1.$ HC2 robust standard errors in parentheses.

D.3.3 Influence of Shepard's Murder on Attitudes Concerning Homosexuality

| | Moral Wrong | | | | |
|-------------------|-----------------|-------------------|--|--|--|
| | (1) | (2) | | | |
| Post-Shepard | -0.11^{***} | -0.12^{***} | | | |
| | (0.02) | (0.02) | | | |
| White | ~ / | -0.10^{*} | | | |
| | | (0.04) | | | |
| Woman | | -0.08^{*} | | | |
| | | (0.03) | | | |
| College | | -0.17^{***} | | | |
| | | (0.04) | | | |
| Age (18-24) | | -0.24^{***} | | | |
| | | (0.07) | | | |
| Age $(25-29)$ | | -0.20^{**} | | | |
| | | (0.07) | | | |
| Age $(30-34)$ | | -0.13^{\dagger} | | | |
| | | (0.07) | | | |
| Age $(35-39)$ | | -0.04 | | | |
| | | (0.06) | | | |
| Age $(40-49)$ | | -0.18^{**} | | | |
| | | (0.06) | | | |
| Age $(50-64)$ | | -0.04 | | | |
| | | (0.06) | | | |
| Income $(20-50k)$ | | 0.01 | | | |
| | | (0.04) | | | |
| Income $(50-75k)$ | | -0.01 | | | |
| | | (0.06) | | | |
| Democrat | | 0.06^{+} | | | |
| | | (0.03) | | | |
| Registered | | -0.02 | | | |
| | | (0.05) | | | |
| \mathbb{R}^2 | 0.00 | 0.07 | | | |
| Ν | 2052 | 2052 | | | |
| State FE | Ν | Y | | | |
| Surveys | CNN Jun/Oct '98 | CNN Jun/Oct '98 | | | |

Table D59: Respondents Interviewed Post-Shepard Are Less Likely To Believe Homosexuality is Morally Wrong

Note: ***p < 0.001, **p < 0.01, *p < 0.05, †p < 0.1. HC2 robust standard errors in parentheses.

D.3.4 Falsification Tests

| Outcome | Post-Shepard Coef. | Post-Shepard Coef. SE p-value N | | $\mathbf{Survey}(\mathbf{s})$ | |
|--|--------------------|---------------------------------|------|-------------------------------|-------------------------------|
| Ban Abortion | 0.00 | 0.02 | 0.88 | 1757 | CNN Jan '98/CNN Oct '98 |
| Affirmative Action 1 | -0.04 | 0.03 | 0.14 | 1970 | CBS Dec '97/CBS Jul '00 |
| Affirmative Action 2 | 0.02 | 0.02 | 0.31 | 2741 | ANES '96-'98 |
| Death Penalty | 0.02 | 0.02 | 0.49 | 2557 | Kaiser Jul '98/Gallup Feb '99 |
| Black People Unintelligent | 0.01 | 0.00 | 0.05 | 4202 | GSS '98-'00 |
| Black People Lazy | 0.00 | 0.01 | 0.56 | 4202 | GSS '98-'00 |
| Spending 2 Aid Black People | 0.00 | 0.02 | 0.96 | 2790 | GSS '98-'00 |
| Black/White Inequality = Discrim. | 0.01 | 0.02 | 0.42 | 3748 | GSS '98-'00 |
| Black/White Inequality = In-Born Ability | 0.03 | 0.01 | 0.02 | 3748 | GSS '98-'00 |
| Black/White Inequality = No Education | 0.02 | 0.02 | 0.35 | 3748 | GSS '98-'00 |
| Black/White Inequality = No Motivation | 0.03 | 0.02 | 0.11 | 3748 | GSS '98-'00 |
| Oppose Living w/Black People | -0.01 | 0.01 | 0.31 | 4202 | GSS '98-'00 |
| Black Feeling Therm. | 0.04 | 0.01 | 0.00 | 2692 | ANES '96-'98 |
| Abortion Any Time | -0.01 | 0.02 | 0.56 | 3546 | GSS '98-'00 |
| Support Female Politicians | 0.01 | 0.02 | 0.67 | 3477 | GSS '98-'00 |
| Working Women Good | -0.06 | 0.02 | 0.00 | 3686 | GSS '98-'00 |
| Working Women Bad 1 | 0.04 | 0.02 | 0.01 | 3615 | GSS '98-'00 |
| Working Women Bad 2 | 0.04 | 0.02 | 0.07 | 2248 | GSS '98-'00 |

Table D60: Falsification Tests

Note: HC2 robust standard errors presented.

D.3.5 Event Study (Moral Wrong)

| | Moral Wrong | |
|-----------------------|-------------|---------------|
| | (1) | (2) |
| 1978 | 0.00 | 0.02 |
| | (0.03) | (0.02) |
| 1992 | 0.01 | -0.02 |
| | (0.02) | (0.02) |
| 1994 | | |
| | () | () |
| 1998 (Shepard Murder) | -0.05^{*} | -0.06^{*} |
| | (0.03) | (0.03) |
| 2001 | 0.02 | 0.01 |
| | (0.03) | (0.03) |
| 2004 | -0.02 | -0.03 |
| | (0.03) | (0.03) |
| Age $(25-34)$ | | -0.02 |
| | | (0.02) |
| Age $(35-49)$ | | 0.05^{**} |
| | | (0.02) |
| Age $(51+)$ | | 0.12*** |
| | | (0.02) |
| Woman | | -0.06*** |
| | | (0.01) |
| White | | -0.02 |
| | | (0.02) |
| College | | -0.14^{***} |
| D | | (0.02) |
| Democrat | | -0.07 |
| | | (0.01) |
| \mathbb{R}^2 | 0.00 | 0.03 |
| Ν | 6130 | 6129 |

Table D61: Event Study Characterizing Trends in Belief Homosexuality is Morally Wrong

Note: ***p < 0.001, **p < 0.01, *p < 0.05, †p < 0.1. Sample is a stacked dataset of surveys with similar items on the covariates displayed on this table. Surveys included in this sample are the TIME 1978, CNN 1992, CNN 1994, CNN 1998, CNN 2001, and CNN 2004 polls. The reference category for the "year" analysis is based on the level of *moral wrong* in the CNN 1994 poll. HC2 robust standard errors in parentheses.

D.4 Alternative Outcomes

We found two other items that could serve as potential candidates for assessing the influence of Shepard's murder on attitudes toward LGBTQ+ group members, however, we do not use them for various reasons. One item measured support for *legal recognition* of "marriages between homosexuals." But, there is a 3-year interval between the two surveys including this outcome item (Gallup Mar. 1996, N = 1008, Gallup Feb. 1999, N = 1054), and there are no surveys with comparable items concerning legal recognition of marriages between homosexuals prior to the baseline time period to conduct temporal placebo tests. Nevertheless, respondents surveyed *post-Shepard* are more likely to support legally recognizing same-sex marriage, consistent with H1 (see Section D.4.1 for details). Another item measures support for homosexuals serving in the armed forces using two surveys 7 months apart (hire *military*, Newsweek Jul. 1998, N = 602; Gallup Feb. 1999, N = 1054). Consistent with H1, we find respondents interviewed *post-Shepard* were more likely to support homosexuals serving in the military (Figure D38, Panel B). However, these effects may be a function of a secular attitudinal trend in support of incorporating homosexuals in the military, perhaps buttressed by Bill Clinton's efforts to implement Don't Ask Don't Tell in the 1990s. We demonstrate this is the case by showing that support for hiring gay people in the military is on an increasing trend from 1977-1996 (Figure D38, Panel C). Conversely, the moral wrong outcome is remarkably stable prior to Shepard's murder, making it an ideal candidate for assessing attitudinal shifts *post-Shepard* and their temporal sustainability (Figure 9).



D.4.1 Alternative Outcome: Legal Recognition

Figure D37: Influence of Shepard's Murder on Support for Legal Recognition of Same-Sex Marriages. All estimates include population weights. All covariates are scaled between 0-1. 95% CIs displayed derived from HC2 robust standard errors.

Data are from two polls stacked together. The first poll is the Gallup March 1996 Politics Polls (N = 1008). It was fielded from March 15-17, 1996 and is a telephone survey. The second poll is the Gallup February 1999 Service Poll (N = 1054). It was fielded from February 8-9, 1999. The main outcome of interest for this analysis is *legal recognition*. Legal recognition is from a common item in these two polls that asks respondents if they "think marriages between homosexuals should or should not be recognized by the law as valid, with the same rights as traditional marriages." Respondents can choose to reply "should be valid" or "should not be recognized." The outcome is measured equal to 1 if the respondent replies with "should be valid," and 0 otherwise. Figure D37, Panel A displays covariate composition balance between the pre- (Gallup 1996) and *post-Shepard* (Gallup 1999) surveys. Panel B displays the influence of being interviewed in the *post-Shepard* survey on respondents reporting that they believe marriages between homosexual should be recognized by the law as valid. Respondents interviewed *post-Shepard* report a 8 percentage point increase in support for the belief homosexuals should have their marriages legally recognized.



D.4.2 Alternative Outcome: Hire Military

Figure D38: Influence of Shepard's Murder on Support for Hiring Gay People To Serve In The Military. Panel A displays covariate balance between the Newsweek Jul '98 and Gallup Feb '99 polls used to assess the influence of being interviewed *post-Shepard* on attitudes toward hiring gay people to serve in the military. Panel B displays a temporal placebo test assessing if mass attitudes on hiring gay people in the military shift between Nov '96 and Jul '98 in addition to coefficients with and without covariate adjustment that assess the influence of being interviewed *post-Shepard* on support for hiring gay people in the military. Panel C displays an event study assessing trends in support for hiring gay people in the military relative to a survey in Jul 1998 (hence no CIs for that survey estimate). All estimates include population weights. All covariates are scaled between 0-1. 95% CIs displayed derived from HC2 robust standard errors.

The two studies we use to assess if the belief homosexuals should be hired for the military increases after Shepard's murder are a Newsweek Jul. 30-31 1998 poll and a Gallup Feb. 8-9 1999 poll. Both are nationally representative adult telephone surveys (N = 602, N = 1054) and are population weighted to census demographic benchmarks.

The outcome item of interest from the Newsweek poll is "Tell me if you think gays and lesbians should be hired as members of the armed forces" with response choices of 1) Should and 2) Should not. The outcome is binary, equal to 1 if the respondent indicates "Should." The outcome item of interest from the Gallup poll is "Do you think homosexuals should or should not be hired for the armed forces" with response choices of 1) Should and 2) Should not. The weights, outcome, and baseline covariates are then stacked amongst each other across the two polls, with respondents from the Gallup Feb. 1999 poll being defined as *post-Shepard* respondents (measured as a binary indicator equal to 1 if the respondent is from the October 1998 poll, 0 otherwise) and respondents from the Newsweek Jul. 1998 poll being defined as pre-Shepard respondents.

To conduct a temporal placebo test ruling out secular trends that may drive our finding that respondents interviewed after Shepard's murder were more likely to support hiring homosexuals in the military, we use a third survey, the Gallup Nov 21-24 1996 poll. The Gallup Nov 1996 poll is a nationally representative telephone survey (N = 1003) and is population weighted to census demographic benchmarks. The Gallup Nov 1996 poll includes an item asking respondents if they think "homosexuals should or should be hired for the armed forces" with responses 1) Should and 2) Should not. We then compare the average level of support for whether homosexuals should be hired for the armed forces in the Gallup 1996 poll with the Newsweek 1998 poll.

Figure D38, Panel A, demonstrates that respondents interviewed before and after Shepard's murder are similar on demographic, socio-economic, and political covariates. Figure D38, Panel B demonstrates that respondents interviewed *post-Shepard* are more likely to support gay people serving in the armed forces by 5 percentage points (p < 0.10), equivalent to 9% of the outcome standard deviation. However, Figure D38, Panel C demonstrates that support for hiring gay people in the military is on an upward trend between 1977-1996, suggesting these results may be a function of a progressive secular trend in support of incorporating gay people in the military, perhaps the result of Bill Clinton's push for Don't Ask Don't Tell policies.

D.5 Temporal Persistence Data Details

TIME 1978 poll (N = 1044): Nationally representative telephone poll sponsored by TIME magazine. Fielded March 14-30, 1978. Item we use asks respondents if "do you personally think that homosexual relationships between consenting adults is morally wrong or not a moral issue. How about? 1) Morally wrong, 2) Not a moral issue" Outcome is coded 1 if respondent indicates "morally wrong."

CNN 1992 poll (N = 1250): Nationally representative telephone poll sponsored by TIME magazine and CNN. Fielded May 13-14, 1992. Item we use asks respondents if "do you personally think that homosexual relationships between consenting adults is morally wrong or not a moral issue. How about? 1) Morally wrong, 2) Not a moral issue" Outcome is coded 1 if respondent indicates "morally wrong."

CNN 1994 poll (N = 800**):** Nationally representative telephone poll sponsored by TIME magazine and CNN. Fielded June 15-16, 1994. Item we use asks respondents if "do you personally think that homosexual relationships between consenting adults is morally wrong or not a moral issue. How about? 1) Morally wrong, 2) Not a moral issue." Outcome is coded 1 if respondent indicates "morally wrong."

CNN 1998 poll (N = 1036): Nationally representative telephone poll sponsored by TIME magazine and CNN. Fielded October 14-15, 1998. Item we use asks respondents if "do you personally think that homosexual relationships between consenting adults is morally wrong or not a moral issue. How about? 1) Morally wrong, 2) Not a moral issue." Outcome is coded 1 if respondent indicates "morally wrong."

CNN 2001 poll (N = 1000): Nationally representative telephone poll sponsored by TIME magazine and CNN. Fielded January 10-11, 2001. Item we use asks respondents if "do you personally think that homosexual relationships between consenting adults is morally wrong or not a moral issue. How about? 1) Morally wrong, 2) Not a moral issue." Outcome is coded 1 if respondent indicates "morally wrong."

CNN 2004 poll (N = 1000): Nationally representative telephone poll sponsored by TIME magazine and CNN. Fielded February 5-6, 2004. Item we use asks respondents if "do you personally think that homosexual relationships between consenting adults is morally wrong or not a moral issue. How about? 1) Morally wrong, 2) Not a moral issue." Outcome is coded 1 if respondent indicates "morally wrong."

D.6 Temporal Persistence CNN Poll (Jun. 1998) Exclusion Details

We do not use the CNN June 1998 poll on Figure 9 in our assessment of temporal persistence. This is because the *moral wrong* item in the CNN June 1998 poll references "homosexual behavior" as opposed to "homosexual relationships between consenting adults." Therefore, we focus on surveys with *moral wrong* outcome items using the "consenting adults" wording for the event study. However, two concerns may arise. First, one may be concerned using the "consenting adults" item in our main analysis on Figure 8 may inflate the post-Shepard coefficient since it may make respondents more comfortable with "homosexual relationships." Yet, we still observe a negative, significant, post-Shepard coefficient (-0.06) from the event study comparing a CNN 1994 poll to the CNN October 1998 poll. Second, one may be concerned the alternative analysis comparing polls between 1994-1998 may be biased by secular attitudinal time trends. But the absence of temporal trends in the "homosexual behavior" items (Figure 8, Panel B) between 1997-1998 suggests this is unlikely. Additionally, the stability of the "consenting adults" moral wrong outcome from 1978-1994 suggests secular attitudinal time trends are not influencing the *post-Shepard* coefficient. However, we caveat our findings by noting the distinct possibility there is an unobserved trend that exists only for the "homosexual behavior" outcome that does not exist for the "consenting adults" outcome. But this may be unlikely because we may expect prosocial secular trends toward people engaged in "homosexual relationships between consenting adults" than toward people engaged in "homosexual behavior" because of the clarity related to "consent."

D.7 Falsification Test Details

Here, we describe in greater detail the falsification tests characterized on Figure 8, Panel C. Only 4/18 outcomes are statistically significant and the *post-Shepard* coefficient is not consistently in support of non-LGBTQ+ marginalized groups, suggesting no systematic secular trend favoring marginalized groups. The Black feeling thermometer shifts in a favorable direction for Black people *post-Shepard*. However, the mass public increasingly attributes Black-White inequality to in-born ability *post-Shepard*, not discrimination, an unfavorable position toward Black people. The mass public is also less likely to believe women can establish a warm relationship with their children and more likely to believe their children will suffer if they work ("working women good", "working women bad 1") *post-Shepard*, both unfavorable attitudinal shifts toward women. Importantly, like the falsification tests in Study 1, the null effects of post-Shepard coefficients are not driven by independent shifts in social conservatism and/or religiosity despite their strong linkages with anti-LGBTQ+ beliefs, but rather, Shepard's murder.

D.8 Falsification Test Outcome Details

Outcome: Ban Abortion. **Surveys:** CNN Jan. '98, CNN Newsweek Oct. '98. **Pre-Shepard Outcome:** "Do you think abortions should be 1) legal under any circumstance, 2) legal under certain circumstances, or 3) illegal in all circumstances." Coded 1 if respondent indicates "legal under any circumstance" and 0 otherwise. **Post-Shepard Outcome:** Same as pre-Shepard

Outcome: Affirmative Action 1. **Surveys:** CBS Dec. '97, CBS Jul. '00. **Pre-Shepard Outcome:** "In order to make up for past discrimination, do you favor or oppose programs which make special efforts to help minorities get ahead?" 1) Favor, 2) Oppose. Coded 1 if respondent indicates favor, 0 otherwise. **Post-Shepard Outcome:** Same as pre-Shepard

Outcome: Affirmative Action 2. **Surveys:** ANES 96'-'98. **Pre-Shepard Outcome:** "Some people say that because of past discrimination, blacks should be given preference in hiring and promotion. Others say that such preference in hiring and promotion of blacks is wrong because it gives blacks advantages they haven't earned. What about your opinion – are you FOR or AGAINST preferential hiring and promotion of blacks? " 1) For preferential hiring and promotion of blacks, 2) Against preferential hiring and promotion of blacks. Coded 1 if respondent indicates for preferential hiring, 0 otherwise. **Post-Shepard Outcome:** Same as pre-Shepard

Outcome: Death Penalty. **Surveys:** Kaiser Jul. '98, Gallup Feb. '99 **Pre-Shepard Outcome:** "Do you favor or oppose the death penalty for persons convicted of murder?" 1) Favor, 2) Oppose. Coded 1 if favor, 0 otherwise. **Post-Shepard Outcome:** "Are you in favor of the death penalty for a person convicted of murder?" 1) Yes, in favor, 2) No, not in favor. Coded 1 if favor, 0 otherwise.

Outcome: Black People are Unintelligent. **Surveys:** GSS '98-'00. **Pre-Shepard Outcome:** "Do people in these groups tend to be unintelligent or tend to be intelligent? Where you you rate Blacks in general on this scale?" 1-7 scale from 1 = unintelligent to 7 = intelligent, reverse coded and rescaled between 0-1. **Post-Shepard Outcome:** Same as pre-Shepard

Outcome: Spending Too Little on Helping Black People **Surveys:** GSS '98-'00. **Pre-Shepard Outcome:** "We are faced with many problems in this country, none of which can be solved easily or inexpensively. I'm going to name some of these problems, and for each one I'd like you to tell me whether you think we're spending too much money on it, too little money, or about the right amount: improving the conditions of Blacks" Coded 1 if too little, 0 otherwise. **Post-Shepard Outcome:** Same as pre-Shepard

Outcome: Black-White Inequality is Because of Discrimination. **Surveys:** GSS '98-'00. **Pre-Shepard Outcome:** "On the average (Negroes/Blacks/African-Americans) have worse jobs, income, and housing than white people. Do you think these differences are: mainly due to discrimination" 1) Yes, 2) No. Coded 1 if yes, 0 otherwise. **Post-Shepard Outcome:** Same as pre-Shepard

Outcome: Black-White Inequality is Because of In-Born Ability. **Surveys:** GSS '98-'00. **Pre-Shepard Outcome:** "On the average (Negroes/Blacks/African-Americans) have worse jobs, income, and housing than white people. Do you think these differences are: Because most (Negroes/Blacks/African-Americans) have less in-born ability to learn?" 1) Yes, 2) No. Coded 1 if yes, 0 otherwise. **Post-Shepard Outcome:** Same as pre-Shepard

Outcome: Black-White Inequality is Because of No Chance for Education. **Surveys:** GSS '98-'00. **Pre-Shepard Outcome:** "On the average (Negroes/Blacks/African-Americans) have worse jobs, income, and housing than white people. Do you think these differences are: Because most (Negroes/Blacks/African-Americans) don't have the chance for education that it takes to rise out of poverty?" 1) Yes, 2) No. Coded 1 if yes, 0 otherwise. **Post-Shepard Outcome:** Same as pre-Shepard

Outcome: Black-White Inequality is Because of No Motivation. Surveys: GSS '98-'00. Pre-Shepard Outcome: "On the average (Negroes/Blacks/African-Americans) have worse jobs, income, and housing

than white people. Do you think these differences are: Because most (Negroes/Blacks/African-Americans) just don't have the motivation or will power to pull themselves up out of poverty?" 1) Yes, 2) No. Coded 1 if yes, 0 otherwise. **Post-Shepard Outcome:** Same as pre-Shepard

Outcome: Oppose Living with Black People. **Surveys:** GSS '98-'00. **Pre-Shepard Outcome:** "Now I'm going to ask you about different types of contact with various groups of people. In each situation would you please tell me whether you would be very much in favor of it happening, somewhat in favor, neither in favor nor opposed to it happening, somewhat opposed, or verymuch opposed to it happening? Living in a neighborhood where half of your neighbors were blacks?" 1-5 scale from 1 = Strongly Favor to 5 = Strongly Oppose. Coded 1 if oppose or strongly oppose, 0 otherwise. **Post-Shepard Outcome:** Same as pre-Shepard

Outcome: Black Feeling Thermometer. **Surveys:** ANES '96-'98. **Pre-Shepard Outcome:** "How would you rate Blacks?" 0-100 scale, rescaled between 0-1. **Post-Shepard Outcome:** Same as pre-Shepard

Outcome: Abortion Any Time. **Surveys:** GSS '98-'00. **Pre-Shepard Outcome:** 'Please tell me whether or not you think it should be possible for a pregnant woman to obtain a legal abortion if the woman wants it for any reason?" 1 if yes. **Post-Shepard Outcome:** Same as pre-Shepard

Outcome: Support Female Politicians. **Surveys:** GSS '98-'00. **Pre-Shepard Outcome:** "Tell me if you agree or disagree with this statement: Most men are better suited emotionally for politics than are most women" 1 if agree, 0 otherwise. **Post-Shepard Outcome:** Same as pre-Shepard

Outcome: Working Women Good. **Surveys:** GSS '98-'00. **Pre-Shepard Outcome:** "Now I'm going to read several more statements. As I read each one, please tell me whether you strongly agree, agree, disagree, or strongly disagree with it. For example, here is the statement: A working mother can establish just as warm and secure a relationship with her children as a mother who does not work." 1 if agree, 0 otherwise. **Post-Shepard Outcome:** Same as pre-Shepard

Outcome: Working Women Bad 1. **Surveys:** GSS '98-'00. **Pre-Shepard Outcome:** "Now I'm going to read several more statements. As I read each one, please tell me whether you strongly agree, agree, disagree, or strongly disagree with it. For example, here is the statement: A preschool child is likely to suffer if his or her mother works." 1 if agree, 0 otherwise. **Post-Shepard Outcome:** Same as pre-Shepard

Outcome: Working Women Bad 2. **Surveys:** GSS '98-'00. **Pre-Shepard Outcome:** "Now I'm going to read several more statements. As I read each one, please tell me whether you strongly agree, agree, disagree, or strongly disagree with it. For example, here is the statement: It is much better for everyone involved if the man is the achiever outside the home and the woman takes care of the home and family." 1 if agree, otherwise. **Post-Shepard Outcome:** Same as pre-Shepard

D.9 Assessing If Violence Against LGBTQ+ Community Segments Was Salient in 4 Months Between Surveys



Figure D39: There Was No New York Times Coverage of Hate Crimes Related to Gay People In Between June-October 1998. The x-axis is the month of 1998, the y-axis is the count of articles identified in the New York Times Historic Database (ProQuest) that are related to the following search term: ("hate crime" AND "gay") OR ("hate crime" AND "homosexual")

D.10 Other Intervening Events

Two other intervening events outside of other instances of violence against LGBTQ+ community segments during 1998 may explain our *post-Shepard* coefficient. First, Clinton signed Executive Order 13087 on May 1998, which prohibited discrimination over sexual orientation in the Federal workforce. If this explains our results, then we would expect the temporal placebo coefficient to be negative and statistically significant given the post-placebo survey is fielded on June 1998, after the executive order. The placebo coefficient is 0 and insignificant, suggesting Clinton's order does not explain our results (Figure 8, Panel B). Indeed, Clinton's order was not nearly as salient as Shepard's murder. There was no NYT coverage of his order on May or June 1998, the moment the executive order was signed (Figure D40). Second, Tammy Baldwin's 1998 House election run (the first open lesbian elected to Congress). This is unlikely because Baldwin's run was significantly less salient than Shepard's murder. There were only 2 NYT articles mentioning Baldwin during June-October 1998 but over 30 Shepard-related NYT articles on October 1998 (Figure D41).

D.10.1 Assessing If Clinton's Anti-Discrimination Executive Order Was Salient



Figure D40: There Were No New York Times Articles Related to Executive Order 13087 Near The Moment It Was Signed. The x-axis is the day, the y-axis is the count of articles identified in the New York Times Historic Database (rtimes package) that are related to the following search terms: "executive order 13087" OR "eeo executive order."

D.10.2 Assessing If Tammy Baldwin's Election Was Salient in 4 Months Between Surveys



Figure D41: There Were Only 2 New York Times Articles Related to Tammy Baldwin In Between June-October 1998. The x-axis is the month of 1998, the y-axis is the count of articles identified in the New York Times Historic Database (rtimes package) that are related to the following search term: "tammy baldwin." Annotations denote number of NYT articles for each specific month.

D.11 Evaluating Individual-Level Heterogeneity

| | Moral Wrong | | | | |
|--------------------------|--------------|--------------|--------------|--|--|
| | (1) | (2) | (3) | | |
| Post-Shepard | -0.07^{**} | -0.10^{**} | -0.02 | | |
| | (0.03) | (0.04) | (0.03) | | |
| Post-Shepard x Non-White | -0.15^{*} | | | | |
| | (0.06) | | | | |
| Post-Shepard x Woman | | -0.02 | | | |
| | | (0.05) | | | |
| Post-Shepard x Democrat | | | -0.20^{**} | | |
| | | | (0.05) | | |
| Non-White | 0.10^{*} | | | | |
| | (0.04) | | | | |
| Woman | -0.08^{*} | -0.08^{*} | -0.08^{*} | | |
| | (0.03) | (0.03) | (0.03) | | |
| Democrat | 0.06 | 0.06 | 0.06 | | |
| | (0.03) | (0.03) | (0.04) | | |
| \mathbb{R}^2 | 0.07 | 0.07 | 0.07 | | |
| Ν | 2052 | 2052 | 2052 | | |

Table D62: Heterogenous Influence of Post-Shepard (Study 3)

Note: ***p < 0.001; **p < 0.01; *p < 0.05. All models adjust for age, white (if not assessing heterogeneity by non-white), woman, college education, partisanship, voter registration, and Florida, Texas, California, New York, and Pennsylvania residence. HC2 robust SEs in parentheses.

E Study 4: Club Q

E.1 Representativeness Discussion

Table E63: Representativeness Assessment of 2022 PI S-IAT and T-IAT

| Survey | College-Educated | White | Woman | Age (18-29) | Age (30-44) | Age (45-59) | Age $(60+)$ |
|---------------|------------------|-------|-------|-------------|-------------|-------------|-------------|
| 2020 Census | 0.32 | 0.57 | 0.51 | 0.20 | 0.26 | 0.24 | 0.30 |
| 2022 PI S-IAT | 0.39 | 0.59 | 0.71 | 0.66 | 0.20 | 0.11 | 0.03 |
| 2022 PI T-IAT | 0.50 | 0.64 | 0.70 | 0.53 | 0.27 | 0.16 | 0.05 |

Table E63 displays marginals along college-education, race, gender, and age characterizing the composition of the adult U.S. population in the U.S. Census (2020) in addition to the 2022 Project Implicit Sexuality Implicit Association Test (PI S-IAT) and 2022 Project Implicit Transgender Implicit Association Test (PI T-IAT) surveys. The 2022 PI S-IAT and T-IAT surveys are more likely to be college-educated, women, and younger. Like Study 2, we are not particularly concerned about the lack of representativeness given prior research demonstrates non-representative samples respond similarly to external stimuli as representative samples (Coppock, 2019). Moreover, the primary purpose of Study 4 is to test **H4**, which posits that less salient events will not motivate prosocial attitudes toward LGBTQ+ community segments. Given the 2022 S-IAT and T-IAT samples are more college-educated, younger, and women, they may be more likely to perceive events that implicate marginalized social groups like LGBTQ+. Therefore, the 2022 S-IAT and T-IAT samples, despite being unrepresentative, possess an advantage in that they provide a hard test for **H4**.

E.2 Anti-Trans Attitudes Outcome Measurement Details

Anti-Trans D-Score: Measured by assessing the speed by which respondents associate negative/positive attributes (words) to images of trans/cis celebrities. Higher value suggest respondents associated negative attributes to trans people faster than they associated negative attributes to cis people. See Axt et al. (2021) for more details. This outcome is rescaled between 0-1 during the analysis.

Cis Bias: Scale from 1-7 from "I strongly prefer transgender people to cisgender people" to "I strongly prefer cisgender people to transgender people." This outcome is rescaled between 0-1 during the analysis, with 1 indicating maximum preferences for cisgender people.

Ciscentrism: Measured with two scales. One scale asks respondents to rate how warm they feel toward transgender people on a scale between 1-10. The other scale asks respondents to rate how warm they feel toward cisgender people on a scale between 1-10. We subtract the scale on warmth toward transgender people from the scale on warmth toward cisgender people. Therefore, higher values suggest more relative warmth toward cisgender people than transgender people. We rescale this measure between 0-1, with 1 indicating maximum warmth toward cisgender people relative to transgender people.

E.3 Salience of Club Q Relative to Pulse and Shepard

E.3.1 New York Times



Figure E42: Number of New York Times Articles Related to Matthew Shepard's Murder, the Pulse Massacre, and the Club Q Shooting In The Two Months After The Event(s). The x-axis is the respective event, the y-axis is the number of articles published in the New York Times in the two months after the incident. Data are from the ProQuest New York Times Historic Newspaper database. Search phrases for the respective incidents are: "matthew shepard AND (murder OR death OR killed)," "pulse AND shooting", and "club q AND shooting."





Figure E43: Count of News Articles Related to Violence Against LGBTQ+ People Six Months Before and After the Pulse Massacre and Club Q Massacre. Panels A-B, C-D, and E-F characterize the count of news articles (y-axis) over time (x-axis) containing the phrases "shooting," "LGBT," and "hate crime" respectively. Panels A, C, E and B, D, F characterize the count of articles over time 6 months before and after the Pulse and Club Q massacres respectively. Dashed vertical line denotes the moment the respective massacres occurred. The dark line characterizes a loess model fit on each side of the moment the respective massacres occurred. Data are from Mediacloud, an open-source platform for media analysis (see: https://www.mediacloud.org/). Annotations denote regression discontinuity-in-time estimates characterizing the effect of the respective massacres on the count of articles related to specific phrases (polynomial degree = 1, kernel = uniform, using CCT optimal bandwidth selection, see Calonico et al. (2015)).



Figure E44: Ratio of News Articles Related to Violence Against LGBTQ+ People vis-a-vis All News Articles Six Months Before and After the Pulse Massacre and Club Q Massacre. Panels A-B, C-D, and E-F characterize the ratio of news articles (y-axis) over time (x-axis) containing the phrases "shooting," "LGBT," and "hate crime" over all news articles respectively. Panels A, C, E and B, D, F characterize the count of articles over time 6 months before and after the Pulse and Club Q massacres respectively. Dashed vertical line denotes the moment the respective massacres occurred. The dark line characterizes a loess model fit on each side of the moment the respective massacres occurred. Data are from Mediacloud, an open-source platform for media analysis (see: https://www.mediacloud.org/). Annotations denote regression discontinuity-in-time estimates characterizing the effect of the respective massacres on the count of articles related to specific phrases (polynomial degree = 1, kernel = uniform, using CCT optimal bandwidth selection, see Calonico et al. (2015)).

Table E64: Assessing Coefficient Differences Between Post-Pulse and Post-Club Q on Media Salience

| Outcome | Topic | RDiT Coef. (Pulse) | RSE (Pulse) | RDiT Coef. (Club Q) | RSE (Club Q) | Coef. Difference | Difference t stat. | Difference p value |
|---------|------------|--------------------|-------------|---------------------|--------------|------------------|--------------------|---------------------------|
| Count | Shooting | 1240.232 | 147.036 | 155.632 | 103.825 | 1084.601 | 6.348 | 0.000 |
| Count | LGBT | 369.582 | 44.839 | 46.168 | 7.430 | 323.414 | 7.129 | 0.000 |
| Count | Hate Crime | 80.012 | 10.373 | 68.077 | 18.824 | 11.935 | 0.596 | 0.553 |
| Ratio | Shooting | 0.121 | 0.020 | 0.014 | 0.004 | 0.108 | 5.182 | 0.000 |
| Ratio | LGBT | 0.042 | 0.007 | 0.004 | 0.001 | 0.038 | 5.420 | 0.000 |
| Ratio | Hate Crime | 0.008 | 0.002 | 0.008 | 0.001 | 0.000 | 0.013 | 0.989 |

Note: All RDiT estimates use a uniform kernel and polynomial degree equal to 1 along with the optimal bandwidth selection mechanism by Calonico et al. (2015). Robust SEs displayed.

E.3.3 Google Trends



Figure E45: Google Search Intensity On Topics Related to LGBT, Hate Crimes, and Mass Shootings Over Time (2016-2022). The x-axis is month, the y-axis is the normalized search intensity for a particular search topic between 2016-2022. From left to right, dashed vertical lines denote the moment of the Pulse massacre and Club Q shooting. Panels A, B, and C characterize search intensity for the following search terms: "LGBT," "shooting," and "LGBT hate crime."

E.4 Balance Tests



E.4.1 Project Implicit Sexuality IAT Data (2022)

Figure E46: Covariate Balance Between Project Implicit Sexuality IAT Survey-Takers Before and After Club Q Massacre. Each coefficient is from a separate model regressing a balance covariate (y-axis) on a binary indicator for taking the Sexuality IAT after the Club Q massacre (*post-Club Q*). Each panel characterizes the sample bandwidth at use (1-40 days from the Club Q massacre) and sample size. Statistically significant coefficients are black, grey otherwise. 95% CIs displayed derived from HC2 robust standard errors.

E.4.2 Project Implicit Transgender IAT Data (2022)



Figure E47: Covariate Balance Between Project Implicit Transgender IAT Survey-Takers Before and After Club Q Massacre. Each coefficient is from a separate model regressing a balance covariate (y-axis) on a binary indicator for taking the Transgender IAT after the Club Q massacre (*post-Club Q*). Each panel characterizes the sample bandwidth at use (1-40 days from the Club Q massacre) and sample size. Statistically significant coefficients are black, grey otherwise. 95% CIs displayed derived from HC2 robust standard errors.
E.5 State-Level Anti-LGBTQ+ Bills Over Time By Partisan Control



Figure E48: Number of State-Level Anti-LGBTQ+ Bills Introduced Over Time By Partisan Control. X-axis is year, y-axis is the number of anti-LGBTQ+ bills introduced. Color denotes state government partisan control of governorship, upper, and lower house. Data on bill introductions are from the American Civil Liberties Union.

Data on partisan control of state government are from Grumbach (2022). Data on the introduction of anti-LGBTQ+ bills between 2018-2022 are from the American Civil Liberties Union (ACLU).⁵⁸ The data include the following types of bills:

- Anti-Transgender Bills
 - Single-sex facility restrictions
 - First amendment defense actions and religious exemptions
 - Restrictions on identification documents
 - Restrictions on health care/gender-affirming care
 - Restrictions on athletics
- Broader Anti-LGBTQ bills
 - Religious exemption bills
 - Religious freedom restoration acts
 - First amendment defense acts
 - Health care access restrictions
 - Adoption and foster care restrictions
 - Marriage-related exemptions

⁵⁸See https://www.aclu.org/past-legislation-affecting-lgbt-rights-across-country-2018, https://www.aclu.org/past-legislation-affecting-lgbt-rights-across-country-2020, https://www.aclu.org/legislation-affecting-lgbtq-rights-across-country-2021, and https://www.aclu.org/legislation-affecting-lgbtq-rights-across-country-2022 for source data.

- Restrictions on schools and student organizations
- Bills preempting local protections

E.6 Anti-LGBTQ+ Right Wing Protests Over Time



Figure E49: Number of Right-Wing Anti-LGBTQ+ Protests Over Time (2020-2022). X-axis is year, y-axis is the number of anti-LGBTQ+ protests. Data are from ACLED (see: https://acleddata.com/, protest keyword = "anti-LGBT")

E.7 Regression Tables

E.7.1 Post-Club Q Coefficients (S-IAT Dataset)

Table E65: Regression Table Characterizing Post-Club Q Coefficients (S-IAT Dataset)

| Club Q Coef. | SE | p-val | Ν | Dataset | Outcome | Bandwidth | Controls |
|--------------|------|-------|----------|---------------|--------------------|-----------|----------|
| 0.00 | 0.00 | 0.86 | 5645.00 | Sexuality IAT | D-Score (Anti-Gay) | 5.00 | No |
| 0.00 | 0.00 | 0.85 | 11068.00 | Sexuality IAT | D-Score (Anti-Gay) | 10.00 | No |
| -0.00 | 0.00 | 0.77 | 17246.00 | Sexuality IAT | D-Score (Anti-Gay) | 15.00 | No |
| -0.00 | 0.00 | 0.13 | 24118.00 | Sexuality IAT | D-Score (Anti-Gay) | 20.00 | No |
| -0.00 | 0.00 | 0.09 | 28949.00 | Sexuality IAT | D-Score (Anti-Gay) | 25.00 | No |
| -0.00 | 0.00 | 0.03 | 33221.00 | Sexuality IAT | D-Score (Anti-Gay) | 30.00 | No |
| -0.00 | 0.00 | 0.06 | 37519.00 | Sexuality IAT | D-Score (Anti-Gay) | 35.00 | No |
| -0.00 | 0.00 | 0.05 | 41263.00 | Sexuality IAT | D-Score (Anti-Gay) | 40.00 | No |
| 0.00 | 0.00 | 0.85 | 5645.00 | Sexuality IAT | D-Score (Anti-Gay) | 5.00 | Yes |
| -0.00 | 0.00 | 0.78 | 11068.00 | Sexuality IAT | D-Score (Anti-Gay) | 10.00 | Yes |
| -0.00 | 0.00 | 0.91 | 17246.00 | Sexuality IAT | D-Score (Anti-Gay) | 15.00 | Yes |
| -0.00 | 0.00 | 0.19 | 24118.00 | Sexuality IAT | D-Score (Anti-Gay) | 20.00 | Yes |
| -0.00 | 0.00 | 0.16 | 28949.00 | Sexuality IAT | D-Score (Anti-Gay) | 25.00 | Yes |
| -0.00 | 0.00 | 0.06 | 33221.00 | Sexuality IAT | D-Score (Anti-Gay) | 30.00 | Yes |
| -0.00 | 0.00 | 0.07 | 37519.00 | Sexuality IAT | D-Score (Anti-Gay) | 35.00 | Yes |
| -0.00 | 0.00 | 0.03 | 41263.00 | Sexuality IAT | D-Score (Anti-Gay) | 40.00 | Yes |
| 0.00 | 0.01 | 0.74 | 5743.00 | Sexuality IAT | Straight Bias | 5.00 | No |
| 0.01 | 0.00 | 0.10 | 11229.00 | Sexuality IAT | Straight Bias | 10.00 | No |
| 0.00 | 0.00 | 0.38 | 17538.00 | Sexuality IAT | Straight Bias | 15.00 | No |
| 0.00 | 0.00 | 0.88 | 24542.00 | Sexuality IAT | Straight Bias | 20.00 | No |
| 0.00 | 0.00 | 0.97 | 29456.00 | Sexuality IAT | Straight Bias | 25.00 | No |
| -0.00 | 0.00 | 0.68 | 33835.00 | Sexuality IAT | Straight Bias | 30.00 | No |
| -0.00 | 0.00 | 0.69 | 38195.00 | Sexuality IAT | Straight Bias | 35.00 | No |
| -0.00 | 0.00 | 0.68 | 41983.00 | Sexuality IAT | Straight Bias | 40.00 | No |
| 0.00 | 0.01 | 0.63 | 5743.00 | Sexuality IAT | Straight Bias | 5.00 | Yes |
| 0.01 | 0.00 | 0.11 | 11229.00 | Sexuality IAT | Straight Bias | 10.00 | Yes |
| 0.00 | 0.00 | 0.12 | 17538.00 | Sexuality IAT | Straight Bias | 15.00 | Yes |
| 0.00 | 0.00 | 0.41 | 24542.00 | Sexuality IAT | Straight Bias | 20.00 | Yes |
| 0.00 | 0.00 | 0.40 | 29456.00 | Sexuality IAT | Straight Bias | 25.00 | Yes |
| 0.00 | 0.00 | 0.73 | 33835.00 | Sexuality IAT | Straight Bias | 30.00 | Yes |
| 0.00 | 0.00 | 0.77 | 38195.00 | Sexuality IAT | Straight Bias | 35.00 | Yes |
| -0.00 | 0.00 | 0.91 | 41983.00 | Sexuality IAT | Straight Bias | 40.00 | Yes |
| 0.00 | 0.00 | 0.77 | 5782.00 | Sexuality IAT | Heterocentrism | 5.00 | No |
| 0.00 | 0.00 | 0.56 | 11299.00 | Sexuality IAT | Heterocentrism | 10.00 | No |
| 0.00 | 0.00 | 0.92 | 17631.00 | Sexuality IAT | Heterocentrism | 15.00 | No |
| -0.00 | 0.00 | 0.69 | 24691.00 | Sexuality IAT | Heterocentrism | 20.00 | No |
| -0.00 | 0.00 | 0.61 | 29632.00 | Sexuality IAT | Heterocentrism | 25.00 | No |
| -0.00 | 0.00 | 0.34 | 34037.00 | Sexuality IAT | Heterocentrism | 30.00 | No |
| -0.00 | 0.00 | 0.62 | 38414.00 | Sexuality IAT | Heterocentrism | 35.00 | No |
| -0.00 | 0.00 | 0.88 | 42225.00 | Sexuality IAT | Heterocentrism | 40.00 | No |
| 0.00 | 0.00 | 0.62 | 5782.00 | Sexuality IAT | Heterocentrism | 5.00 | Yes |
| 0.00 | 0.00 | 0.75 | 11299.00 | Sexuality IAT | Heterocentrism | 10.00 | Yes |
| 0.00 | 0.00 | 0.53 | 17631.00 | Sexuality IAT | Heterocentrism | 15.00 | Yes |
| 0.00 | 0.00 | 0.83 | 24691.00 | Sexuality IAT | Heterocentrism | 20.00 | Yes |
| 0.00 | 0.00 | 0.80 | 29632.00 | Sexuality IAT | Heterocentrism | 25.00 | Yes |
| -0.00 | 0.00 | 0.77 | 34037.00 | Sexuality IAT | Heterocentrism | 30.00 | Yes |
| 0.00 | 0.00 | 0.97 | 38414.00 | Sexuality IAT | Heterocentrism | 35.00 | Yes |
| -0.00 | 0.00 | 1.00 | 42225.00 | Sexuality IAT | Heterocentrism | 40.00 | Yes |
| | | | | ~ | | | |

HC2 robust SEs reported

Table E66: Regression Table Characterizing Post-Club Q Coefficients (T-IAT Dataset)

| Club Q Coef. | SE | p-val | N | Dataset | Outcome | Bandwidth | Controls |
|--------------|------|-------|----------|-----------------|----------------------|-----------|----------|
| 0.00 | 0.01 | 0.86 | 2010.00 | Transgender IAT | D-Score (Anti-Trans) | 5.00 | No |
| 0.00 | 0.00 | 0.80 | 4038.00 | Transgender IAT | D-Score (Anti-Trans) | 10.00 | No |
| -0.00 | 0.00 | 0.60 | 6185.00 | Transgender IAT | D-Score (Anti-Trans) | 15.00 | No |
| -0.00 | 0.00 | 0.44 | 8856.00 | Transgender IAT | D-Score (Anti-Trans) | 20.00 | No |
| -0.00 | 0.00 | 0.30 | 11013.00 | Transgender IAT | D-Score (Anti-Trans) | 25.00 | No |
| -0.00 | 0.00 | 0.10 | 12730.00 | Transgender IAT | D-Score (Anti-Trans) | 30.00 | No |
| -0.00 | 0.00 | 0.17 | 14453.00 | Transgender IAT | D-Score (Anti-Trans) | 35.00 | No |
| -0.00 | 0.00 | 0.12 | 16044.00 | Transgender IAT | D-Score (Anti-Trans) | 40.00 | No |
| -0.00 | 0.01 | 0.58 | 2010.00 | Transgender IAT | D-Score (Anti-Trans) | 5.00 | Yes |
| -0.00 | 0.00 | 0.62 | 4038.00 | Transgender IAT | D-Score (Anti-Trans) | 10.00 | Yes |
| -0.00 | 0.00 | 0.33 | 6185.00 | Transgender IAT | D-Score (Anti-Trans) | 15.00 | Yes |
| -0.00 | 0.00 | 0.22 | 8856.00 | Transgender IAT | D-Score (Anti-Trans) | 20.00 | Yes |
| -0.00 | 0.00 | 0.18 | 11013.00 | Transgender IAT | D-Score (Anti-Trans) | 25.00 | Yes |
| -0.00 | 0.00 | 0.06 | 12730.00 | Transgender IAT | D-Score (Anti-Trans) | 30.00 | Yes |
| -0.00 | 0.00 | 0.09 | 14453.00 | Transgender IAT | D-Score (Anti-Trans) | 35.00 | Yes |
| -0.00 | 0.00 | 0.04 | 16044.00 | Transgender IAT | D-Score (Anti-Trans) | 40.00 | Yes |
| 0.01 | 0.01 | 0.21 | 2114.00 | Transgender IAT | Cis Bias | 5.00 | No |
| 0.01 | 0.01 | 0.24 | 4261.00 | Transgender IAT | Cis Bias | 10.00 | No |
| -0.00 | 0.01 | 0.86 | 6516.00 | Transgender IAT | Cis Bias | 15.00 | No |
| -0.00 | 0.00 | 0.66 | 9327.00 | Transgender IAT | Cis Bias | 20.00 | No |
| -0.00 | 0.00 | 0.48 | 11586.00 | Transgender IAT | Cis Bias | 25.00 | No |
| -0.00 | 0.00 | 0.19 | 13390.00 | Transgender IAT | Cis Bias | 30.00 | No |
| -0.00 | 0.00 | 0.36 | 15189.00 | Transgender IAT | Cis Bias | 35.00 | No |
| -0.00 | 0.00 | 0.47 | 16851.00 | Transgender IAT | Cis Bias | 40.00 | No |
| 0.00 | 0.01 | 0.70 | 2114.00 | Transgender IAT | Cis Bias | 5.00 | Yes |
| 0.00 | 0.01 | 0.87 | 4261.00 | Transgender IAT | Cis Bias | 10.00 | Yes |
| -0.00 | 0.00 | 0.45 | 6516.00 | Transgender IAT | Cis Bias | 15.00 | Yes |
| -0.00 | 0.00 | 0.38 | 9327.00 | Transgender IAT | Cis Bias | 20.00 | Yes |
| -0.00 | 0.00 | 0.54 | 11586.00 | Transgender IAT | Cis Bias | 25.00 | Yes |
| -0.00 | 0.00 | 0.31 | 13390.00 | Transgender IAT | Cis Bias | 30.00 | Yes |
| -0.00 | 0.00 | 0.54 | 15189.00 | Transgender IAT | Cis Bias | 35.00 | Yes |
| -0.00 | 0.00 | 0.65 | 16851.00 | Transgender IAT | Cis Bias | 40.00 | Yes |
| 0.01 | 0.01 | 0.08 | 2148.00 | Transgender IAT | Ciscentrism | 5.00 | No |
| 0.01 | 0.00 | 0.06 | 4333.00 | Transgender IAT | Ciscentrism | 10.00 | No |
| 0.00 | 0.00 | 0.30 | 6627.00 | Transgender IAT | Ciscentrism | 15.00 | No |
| 0.00 | 0.00 | 0.86 | 9479.00 | Transgender IAT | Ciscentrism | 20.00 | No |
| -0.00 | 0.00 | 0.56 | 11764.00 | Transgender IAT | Ciscentrism | 25.00 | No |
| -0.00 | 0.00 | 0.07 | 13590.00 | Transgender IAT | Ciscentrism | 30.00 | No |
| -0.00 | 0.00 | 0.08 | 15412.00 | Transgender IAT | Ciscentrism | 35.00 | No |
| -0.00 | 0.00 | 0.08 | 17095.00 | Transgender IAT | Ciscentrism | 40.00 | No |
| 0.01 | 0.01 | 0.28 | 2148.00 | Transgender IAT | Ciscentrism | 5.00 | Yes |
| 0.00 | 0.00 | 0.30 | 4333.00 | Transgender IAT | Ciscentrism | 10.00 | Yes |
| 0.00 | 0.00 | 0.48 | 6627.00 | Transgender IAT | Ciscentrism | 15.00 | Yes |
| -0.00 | 0.00 | 0.84 | 9479.00 | Transgender IAT | Ciscentrism | 20.00 | Yes |
| -0.00 | 0.00 | 0.79 | 11764.00 | Transgender IAT | Ciscentrism | 25.00 | Yes |
| -0.00 | 0.00 | 0.18 | 13590.00 | Transgender IAT | Ciscentrism | 30.00 | Yes |
| -0.00 | 0.00 | 0.21 | 15412.00 | Transgender IAT | Ciscentrism | 35.00 | Yes |
| -0.00 | 0.00 | 0.22 | 17095.00 | Transgender IAT | Ciscentrism | 40.00 | Yes |

HC2 robust SEs reported

E.8 Evaluating Individual-Level Heterogeneity

E.8.1 Sexuality IAT

Table E67: Heterogeneous Influence of Club Q Massacre (S-IAT Dataset)

| Interaction | Coefficient | \mathbf{SE} | p-value | Dataset | Outcome | Bandwidth | Ν | R-Squared |
|--|-------------|---------------|---------|---------------|--------------------|-----------|-------|------------------|
| Post-Club Q x Non-White | 0.00 | 0.00 | 0.93 | Sexuality IAT | D-Score (Anti-Gay) | 20.00 | 24118 | 0.16 |
| Post-Club Q x Woman | -0.00 | 0.00 | 0.71 | Sexuality IAT | D-Score (Anti-Gay) | 20.00 | 24118 | 0.16 |
| Post-Club Q x Liberal | 0.00 | 0.00 | 0.87 | Sexuality IAT | D-Score (Anti-Gay) | 20.00 | 24118 | 0.16 |
| Post-Club Q x % LGBT (State) | 0.00 | 0.00 | 0.59 | Sexuality IAT | D-Score (Anti-Gay) | 20.00 | 24118 | 0.17 |
| Post-Club Q x SS Couple Density (County) | -0.00 | 0.00 | 0.68 | Sexuality IAT | D-Score (Anti-Gay) | 20.00 | 19057 | 0.17 |
| Post-Club Q x Non-White | -0.01 | 0.01 | 0.26 | Sexuality IAT | Straight Bias | 20.00 | 24542 | 0.23 |
| Post-Club Q x Woman | -0.01 | 0.01 | 0.08 | Sexuality IAT | Straight Bias | 20.00 | 24542 | 0.23 |
| Post-Club Q x Liberal | -0.00 | 0.01 | 0.69 | Sexuality IAT | Straight Bias | 20.00 | 24542 | 0.23 |
| Post-Club Q x % LGBT (State) | -0.00 | 0.00 | 0.78 | Sexuality IAT | Straight Bias | 20.00 | 24542 | 0.23 |
| Post-Club Q x SS Couple Density (County) | -0.00 | 0.00 | 0.21 | Sexuality IAT | Straight Bias | 20.00 | 19492 | 0.23 |
| Post-Club Q x Non-White | 0.00 | 0.00 | 0.95 | Sexuality IAT | Heterocentrism | 20.00 | 24691 | 0.25 |
| Post-Club Q x Woman | -0.01 | 0.00 | 0.15 | Sexuality IAT | Heterocentrism | 20.00 | 24691 | 0.25 |
| Post-Club Q x Liberal | -0.00 | 0.00 | 0.49 | Sexuality IAT | Heterocentrism | 20.00 | 24691 | 0.25 |
| Post-Club Q x % LGBT (State) | 0.00 | 0.00 | 0.95 | Sexuality IAT | Heterocentrism | 20.00 | 24691 | 0.25 |
| Post-Club Q x SS Couple Density (County) | -0.00 | 0.00 | 0.49 | Sexuality IAT | Heterocentrism | 20.00 | 19592 | 0.25 |

HC2 robust SEs reported. Each interaction coefficient is from a separate model.

E.8.2 Transgender IAT

Table E68: Heterogenous Influence of Club Q Massacre (T-IAT Dataset)

| Interaction | Coefficient | SE | p-value | Dataset | Outcome | Bandwidth | Ν | R-Squared |
|--|-------------|-------|---------|-----------------|----------------------|-----------|------|-----------|
| Post-Club Q x Non-White | 0.006 | 0.007 | 0.402 | Transgender IAT | D-Score (Anti-Trans) | 15.000 | 6185 | 0.116 |
| Post-Club Q x Woman | -0.001 | 0.007 | 0.861 | Transgender IAT | D-Score (Anti-Trans) | 15.000 | 6185 | 0.116 |
| Post-Club Q x Liberal | -0.008 | 0.007 | 0.197 | Transgender IAT | D-Score (Anti-Trans) | 15.000 | 6185 | 0.116 |
| Post-Club Q x % LGBT (State) | -0.006 | 0.005 | 0.219 | Transgender IAT | D-Score (Anti-Trans) | 15.000 | 6185 | 0.117 |
| Post-Club Q x SS Couple Density (County) | -0.002 | 0.001 | 0.072 | Transgender IAT | D-Score (Anti-Trans) | 15.000 | 4910 | 0.120 |
| Post-Club Q x Non-White | -0.006 | 0.010 | 0.580 | Transgender IAT | Cis Bias | 15.000 | 6516 | 0.188 |
| Post-Club Q x Woman | -0.024 | 0.011 | 0.035 | Transgender IAT | Cis Bias | 15.000 | 6516 | 0.189 |
| Post-Club Q x Liberal | -0.007 | 0.010 | 0.512 | Transgender IAT | Cis Bias | 15.000 | 6516 | 0.188 |
| Post-Club Q x % LGBT (State) | -0.007 | 0.006 | 0.216 | Transgender IAT | Cis Bias | 15.000 | 6516 | 0.190 |
| Post-Club Q x SS Couple Density (County) | -0.001 | 0.002 | 0.624 | Transgender IAT | Cis Bias | 15.000 | 5179 | 0.183 |
| Post-Club Q x Non-White | -0.009 | 0.007 | 0.163 | Transgender IAT | Ciscentrism | 15.000 | 6627 | 0.193 |
| Post-Club Q x Woman | -0.008 | 0.008 | 0.287 | Transgender IAT | Ciscentrism | 15.000 | 6627 | 0.193 |
| Post-Club Q x Liberal | -0.012 | 0.007 | 0.073 | Transgender IAT | Ciscentrism | 15.000 | 6627 | 0.193 |
| Post-Club Q x % LGBT (State) | -0.004 | 0.004 | 0.276 | Transgender IAT | Ciscentrism | 15.000 | 6627 | 0.195 |
| Post-Club Q x SS Couple Density (County) | 0.000 | 0.001 | 0.825 | Transgender IAT | Ciscentrism | 15.000 | 5252 | 0.190 |

 $\mathrm{HC2}$ robust SEs reported. Each interaction coefficient is from a separate model.

F Less Salient Violent Events

F.1 Salience: Search and Analysis Rules

Here, we assess the salience of several relatively prominent anti-LGBTQ+ violent events relative to the Pulse massacre, Matthew Shepard's murder, and the Club Q massacre between 2000-2022. The universe of events we assess is from this crowd-soruced list: https://en.wikipedia.org/wiki/History_of_violence_ against_LGBT_people_in_the_United_States. To assess salience, we assess the number of search hits related to each event from the New York Times.

The Google search term we use to assess salience is: site:nytimes.com "[name of victim]" AND LGBT OR LGBTQ OR gay OR lesbian OR bisexual OR queer OR transgender OR trans OR homophobic OR transphobic AND attack OR assault OR murder OR kill OR killed OR killing OR death"

In cases where a particular place is attacked (e.g. Pulse, or Club Q), we replace "name of victim" with the place the attack occurred (e.g. "Pulse," "Club Q").

F.2 Salience of Less Salient Violent Events (2000-2022)



Figure F50: Salience of Less Salient Violent Incidents Against LGBTQ+ Group Members Relative to the Pulse Massacre, Shepard's Murder, and the Club Q massacre. Panels A/B characterizes the salience (x-axis, number of NYT articles) of incidents (y-axis) from 2000-2009/2010-2022. Panel C characterizes the salience of Shepard's murder, the Pulse massacre, and the Club Q massacre. Annotations denote number of New York Times hits. See Section F.1 for information on measurement of violent incidents and salience.

F.3 Assessing Influence of Less Salient Violent Events on Prosocial Attitudes (2010-2022)



Figure F51: Influence of Less Salient Violent Incidents Against LGBTQ+ Group Members on Prosocial Attitudes Toward Gay People. Panels A/B characterize the influence of incidents on prosocial attitudes from 2010-2016/2017-2022. The x-axis is the post-incident coefficient, the y-axis is the name of victim and date of the respectively violent incident. Shape denotes outcome at use (*D-score, heterocentrism, straight bias*). Grey coefficients are statistically insignificant, black otherwise. Each panel contains two facets using data 15 days before and after the respective violent incident (left) and 20 days before and after the incident (right). 95% CIs displayed derived from HC2 robust SEs.

In this analysis, we examine the influence of less salient violent incidents against LGBTQ+ group members on prosocial attitudes toward gay people between 2010-2022 (see Figure F50, see also https://en.wikipedia.org/wiki/History_of_violence_against_LGBT_people_

in_the_United_States). Similar to Studies 2 and 4, we use Project Implicit Sexuality Implicit Association Test surveys on U.S. adults from 2010-2022 to conduct this analysis. In the analysis, we exclude less salient incidents where 1) there were days of missing data 15 and 20-days before and after the onset of a particular violent incident and 2) there were not 20 days of pre-treatment data for each respective yearly survey (e.g. if an incident occurred on January 7th in a particular year, where there is only 6 days of pre-treatment data for that particular year). Like Studies 2 and 4, We assess the effect of each incident on the *D-score*, straight bias, and heterocentrism.

G Validating Outcomes Across Studies

In this section, we show our outcomes capture the same concept despite differences in measurement and temporal domain across studies. If our outcomes are measuring the same concept across time, they should meet two criteria: 1) they should be highly correlated consistently with each other across several time periods, 2) they should have similar correlates over time. Yearly S-IAT surveys from 2010-2018 show SSM support (Study 1) is consistently strongly associated with the anti-gay *D*-score, straight bias, and heterocentrism outcomes (Study 2). The min-max association between SSM support and the Study 2 outcomes is 30%-100% of the Study 2 outcome scales after covariate adjustment (Figure G53). Although we can't correlate Study 2's outcomes with the moral wrong (Study 3) outcome due to data limitations, we can show SSM support is highly correlated with moral wrong for nearly three decades across several surveys between 1978-2004. Those who support SSM are 35-55 percentage points less likely to believe homosexuality is immoral (Figure G54). Given moral wrong is consistently highly correlated with SSM support over several decades and SSM support is consistently correlated with the Study 2 outcomes for a decade, we can safely assume the outcomes from Studies 1-3 are capturing a similar concept despite measurement and temporal differences. Moreover, socio-demographic and political correlates of the Study 2 and Study 3 outcomes are the same between 2010-2018 and 1978-2004 respectively (Figures G55-G56), suggesting safety in assuming the outcomes are measuring the same concept over time.⁵⁹ Additionally, we evaluate common correlates across Studies 1-4 and show, for the most part, all outcomes are correlated similarly with particular socio-demographic and political factors (Figure G57). These findings further suggest the outcomes from Studies 1-4 are capturing a similar concept despite measurement and temporal differences. These findings also validate our theoretic approach, which is to speak to prosocial attitudes in a broad, multi-dimensional manner.

 $^{^{59}{\}rm This}$ also validates our event study on Figure 9. Despite long-term differences in measuring moral wrong, the concept doesn't change much over time.

G.1 System 1 and System 2 =Related



Correlation Between Explicit Anti-Gay Attitudes and the D Score

Figure G52: The Explicit Anti-Gay Attitude Outcomes are Highly Correlated with the Implicit Anti-Gay Attitude Outcome (D Score). The x-axis is the coefficient for the respective explicit anti-gay attitude outcome (y-axis). Color denotes PI S-IAT dataset at use. The left and right panels characterize estimates without and with covariate adjustment (age, gender, white, college education, ideology, religious, non-metro resident, California resident, New York resident, Florida resident, and Illinois resident). All covariates rescaled between 0-1. 95% CIs displayed derived from HC2 robust SEs.

| Dataset | Covariate | Pearson's ρ With D Score | p-value |
|------------|----------------|-------------------------------|-----------|
| S-IAT 2010 | Heterocentrism | 0.39 | p < 0.001 |
| S-IAT 2010 | Straight Bias | 0.42 | p < 0.001 |
| S-IAT 2011 | Heterocentrism | 0.38 | p < 0.001 |
| S-IAT 2011 | Straight Bias | 0.41 | p < 0.001 |
| S-IAT 2012 | Heterocentrism | 0.39 | p < 0.001 |
| S-IAT 2012 | Straight Bias | 0.41 | p < 0.001 |
| S-IAT 2013 | Heterocentrism | 0.38 | p < 0.001 |
| S-IAT 2013 | Straight Bias | 0.40 | p < 0.001 |
| S-IAT 2014 | Heterocentrism | 0.39 | p < 0.001 |
| S-IAT 2014 | Straight Bias | 0.41 | p < 0.001 |
| S-IAT 2015 | Heterocentrism | 0.38 | p < 0.001 |
| S-IAT 2015 | Straight Bias | 0.40 | p < 0.001 |
| S-IAT 2016 | Heterocentrism | 0.41 | p < 0.001 |
| S-IAT 2016 | Straight Bias | 0.43 | p < 0.001 |
| S-IAT 2017 | Heterocentrism | 0.43 | p < 0.001 |
| S-IAT 2017 | Straight Bias | 0.45 | p < 0.001 |
| S-IAT 2018 | Heterocentrism | 0.43 | p < 0.001 |
| S-IAT 2018 | Straight Bias | 0.45 | p < 0.001 |

Table G69: Pearson's Rho Correlation Coefficients Characterizing AssociationBetween D Score and Explicit Measures of Anti-Gay Attitudes

G.2 Demonstrating Study 2 Outcomes = Associated With SSM Support



Figure G53: The *D* score, Straight Bias, and Heterocentrism Items are Highly Correlated With SSM Support Over an 8-year Period. The x-axis is the coefficient characterizing the relationship between the D score, straight bias, and heterocentrism (specified on y-axis) and support for same-sex marriage. Color denotes PI S-IAT dataset at use. The left and right panels characterize estimates without and with covariate adjustment (age, gender, white, college education, ideology, religious, non-metro resident, California resident, New York resident, Florida resident, and Illinois resident). All covariates rescaled between 0-1. 95% CIs displayed derived from HC2 robust SEs.

Across all PI S-IAT studies, *SSM support* is based on an item asking respondents "Do you think marriages between homosexuals should or should not be recognized by the law as valid, with the same rights as traditional marriages?" with response options: 1) should be valid, 2) should not be valid, 3) no opinion. We code SSM support as 1 if the respondent indicates "should be valid" and 0 otherwise.

The reason we do not use the *SSM support* measure as an outcome in Study 2 is because the item was not asked between January-July 2016, preventing us from using an unexpected-event-during-survey design with the outcome. Our estimates are from respondents who took the 2016 PI S-IAT survey after July 2016.

G.3 Demonstrating Moral Wrong (Study 3) Outcome = Associated With SSM Support



Figure G54: The Moral Wrong Item is Highly Correlated With SSM Support Over 3 Decades. The x-axis is the study at use. The y-axis is the moral wrong coefficient where support for gay marriage is the outcome. Color denotes the inclusion/exclusion of controls (age, gender, race, college-education, partian identification). All covariates are rescaled between 0-1. 95% CIs displayed derived from HC2 robust SEs.

Note: The TIME 1978 study does not have an explicit *SSM support* item. Instead, we use a proxy that characterizes whether respondents believe homosexual relationships are acceptable (see the measurement of SSM support across the studies characterized on Figure G53 below).

Homosexual Relationship Item (TIME 1978): Today there are many different kinds of lifestyles which people find acceptable, such as a husband staying home and caring for the children while the wife goes to work. How do you feel about this? Do you find it acceptable for other people but not for yourself, acceptable for other people and yourself, or not acceptable at all? Homosexual relationships. 1) Acceptable for others, 2) Acceptable for others and self, 3) Not acceptable. Coded 1 if response is "Acceptable for others" OR "Acceptable for others and self," 0 otherwise.

Gay Marriage Support Item (CNN 1994): Do you think marriages between homosexual men or homosexual women should be recognized as legal by the law? 1) Yes, 2) No. Coded 1 if response is "yes," 0 otherwise.

Gay Marriage Support Item (CNN 1998): Do you think marriages between homosexual men or between homosexual women should be recognized as legal by the law? 1) Yes, 2) No. Coded 1 if response is "yes," 0 otherwise.

Gay Marriage Support Item (CNN 2004): On another topic, do you think marriages between homosexual men or between homosexual women should be recognized as legal by the law, or not? 1) Yes, 2) No. Coded 1 if response is "yes," 0 otherwise.

G.4 Moral Wrong Outcome Correlates Between 1978-2004



Correlates of Moral Wrong Outcome Over Time

Figure G55: Correlates of *Moral Wrong* Outcome Over 3 Decades. The x-axis is coefficient for the respective covariate (y-axis, fully-specified model for each study). Color denotes the survey at use. All covariates are rescaled between 0-1. 95% CIs displayed derived from HC2 robust SEs.

G.5 Study 2 Outcome Correlates Between 2010-2018



Figure G56: Correlates of *D score*, *Heterocentrism*, and *Straight Bias* Outcomes Between 2010-2018. The x-axis is coefficient for the respective covariate (y-axis, fullyspecified model for each study). Color denotes the PI S-IAT dataset at use. All covariates are rescaled between 0-1. 95% CIs displayed derived from HC2 robust SEs.



G.6 Correlates Across All Studies

Figure G57: Consistent Correlates of Outcomes Across All Studies. The x-axis is coefficient for the respective covariate (y-axis, fully-specified model for each study). Color denotes the outcome and study dataset at use. All covariates are rescaled between 0-1. Unlike in the main text, the outcome for Study 1 is now reverse coded (*SSM No Support*). This is to maintain directional consistency with the outcomes from other studies. Moreover, there is no liberal ideology item in the Study 3 CNN surveys from 1998. The coefficient presented here for Study 3 is for Democratic partian identification, which is available in the 1998 CNN surveys and is known to be highly correlated with a liberal ideology. 95% CIs displayed derived from HC2 robust SEs.

H Assessing Heterogenous Decay

In this section, we assess if there is heterogeneous decay in the effects of violence against LGBTQ+ group members on prosocial attitudes toward LGBTQ+ group members for Studies 1-3. Given Studies 1-3 posit prosocial attitudinal shifts are temporally unsustainable, it may be the case that the sustainability of these shifts are conditional on the individual-level characteristics we theorize about on Section 3.1. We do not find heterogeneity in the temporal persistence of the effects we identify in Studies 1-3. We remove all respondents 1-10 days after the Pulse massacre and assess the differential effect of *Post-Pulse* by race, gender, liberalism, and LGBTQ+ geographic context and find no evidence of heterogeneous decay in Study 1 (Table H70). We remove all respondents 40 days after the Pulse massacre and assess the differential effect of *Post-Pulse* comparing respondents interviewed in the 15 days before Pulse relative to the 15 days after 40 days after Pulse by race, gender, liberalism, and LGBTQ+ geographic context and find largely null evidence of heterogeneous decay in Study 2 (Table H71). We assess the differential effect of being interviewed in 2001 or 2004 on moral wrong by race, gender, and Democratic partial partial and find largely null evidence of heterogeneous decay in Study 3 Table H72). Linear terms and control covariates are included in the models but omitted from the tables in these analyses. In summary, we find limited evidence of heterogeneous decay.

| | SSM Support | | | | | |
|--------------------------------|-------------|--------|--------|--------|--------|--|
| Post-Pulse x Non-White | 0.06 | | | | | |
| | (0.13) | | | | | |
| Post-Pulse x Female | | -0.12 | | | | |
| | | (0.12) | | | | |
| Post-Pulse x Liberal | | | -0.01 | | | |
| | | | (0.13) | | | |
| Post-Pulse x % LGBTQ (State) | | | | 0.02 | | |
| | | | | (0.65) | | |
| Post-Pulse x SS Couple Density | | | | | -0.05 | |
| | | | | | (0.53) | |
| \mathbb{R}^2 | 0.37 | 0.37 | 0.37 | 0.37 | 0.37 | |
| Num. obs. | 812 | 812 | 812 | 812 | 812 | |

 Table H70:
 There is no heterogeneous decay in the Post-Pulse effect

 $^{***}p < 0.001; \ ^{**}p < 0.01; \ ^{*}p < 0.05; \ ^{\dagger}p < 0.1$

Study 2 H.2

Table H71: Assessing Heterogenous Decay of Post-Pulse (Study 2, Part 1)

| | D-Score | Heterocentrism | D-Score | Heterocentrism | D-Score | Heterocentrism |
|------------------------|---------------|----------------|-------------------|----------------|---------------|----------------|
| Post-Pulse | -0.02^{**} | -0.02^{***} | -0.02^{\dagger} | -0.03^{***} | -0.01 | -0.02^{*} |
| | (0.01) | (0.00) | (0.01) | (0.01) | (0.01) | (0.01) |
| Post-Pulse x Non-White | -0.00 | 0.02^{*} | | | | |
| | (0.01) | (0.01) | | | | |
| Post-Pulse x Woman | | | -0.00 | 0.03^{**} | | |
| | | | (0.01) | (0.01) | | |
| Post-Pulse x Liberal | | | | | -0.01 | 0.01 |
| | | | | | (0.01) | (0.01) |
| Non-White | 0.02^{*} | 0.01 | | | | |
| | (0.01) | (0.01) | | | | |
| Woman | -0.03^{***} | 0.00 | -0.03^{***} | -0.01^{*} | -0.03^{***} | 0.00 |
| | (0.00) | (0.00) | (0.01) | (0.01) | (0.00) | (0.00) |
| Liberal | -0.08^{***} | -0.08^{***} | -0.08^{***} | -0.08^{***} | -0.08^{***} | -0.08^{***} |
| | (0.01) | (0.00) | (0.01) | (0.00) | (0.01) | (0.01) |
| \mathbb{R}^2 | 0.14 | 0.16 | 0.14 | 0.16 | 0.14 | 0.16 |
| Num. obs. | 3374 | 3360 | 3374 | 3360 | 3374 | 3360 |

 $^{***}p < 0.001; \ ^{**}p < 0.01; \ ^{*}p < 0.05; \ ^{\dagger}p < 0.1$

Table H72: Assessing Heterogenous Decay of Post-Pulse (Study 2, Part 2)

| | D-Score | Heterocentrism | $\operatorname{D-Score}$ | Heterocentrism |
|---|---------|----------------|--------------------------|------------------|
| Post-Pulse | 0.00 | -0.02 | -0.00 | -0.02^{**} |
| | (0.02) | (0.01) | (0.01) | (0.01) |
| Post-Pulse x % LGBT (State) | -0.04 | 0.02 | | |
| | (0.04) | (0.03) | | |
| Post-Pulse x SS Couple Density (County) | | | -0.07^{\dagger} | 0.05^{\dagger} |
| | | | (0.03) | (0.02) |
| % LGBT (State) | -0.00 | 0.01 | | |
| | (0.03) | (0.05) | | |
| SS Couple Density (County) | | | -0.03 | -0.09^{**} |
| | | | (0.03) | (0.02) |
| \mathbb{R}^2 | 0.14 | 0.16 | 0.14 | 0.17 |
| Num. obs. | 3374 | 3360 | 3374 | 3360 |
| N Clusters | 52 | 52 | 725 | 721 |
| | | | | |

 $^{***}p < 0.001; \ ^{**}p < 0.01; \ ^{*}p < 0.05; \ ^{\dagger}p < 0.1$

H.3 Study 3

| | Moral Wrong | | | | | |
|----------------------|-------------|--------|--------|--|--|--|
| Non-White x 2001 | 0.09 | | | | | |
| | (0.07) | | | | | |
| Non-White ≥ 2004 | 0.10 | | | | | |
| UU 0001 | (0.07) | 0.00 | | | | |
| Woman x 2001 | | -0.02 | | | | |
| Woman x 2004 | | (0.05) | | | | |
| Wollian x 2004 | | (0.05) | | | | |
| Democrat x 2001 | | (0.00) | -0.03 | | | |
| | | | (0.05) | | | |
| Democrat x 2004 | | | -0.03 | | | |
| | | | (0.06) | | | |
| \mathbb{R}^2 | 0.04 | 0.04 | 0.04 | | | |
| Num. obs. | 6129 | 6129 | 6129 | | | |

 Table H73: Heterogenous Decay of Post-Shepard (Study 3)

***p < 0.001; **p < 0.01; *p < 0.05

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